

# Supplementary Material: Recommendation from Raw Data with Adaptive Compound Poisson Factorization

Olivier Gouvert, Thomas Oberlin, Cédric Févotte  
 IRIT, Université de Toulouse, CNRS, France  
 firstname.lastname@irit.fr

## 1 Stirling Numbers

The Stirling numbers of the three kinds are three different ways to partition  $y$  elements into  $n$  groups.

- The Stirling number of the first kind corresponds to the number of ways of partitioning  $y$  elements into  $n$  disjoint cycles.
- The Stirling number of the second kind corresponds to the number of ways of partitioning  $y$  elements into  $n$  non-empty subsets.
- The Stirling number of the third kind (also known as Lah number) corresponds to the number of ways of partitioning  $y$  elements into  $n$  non-empty ordered subsets.

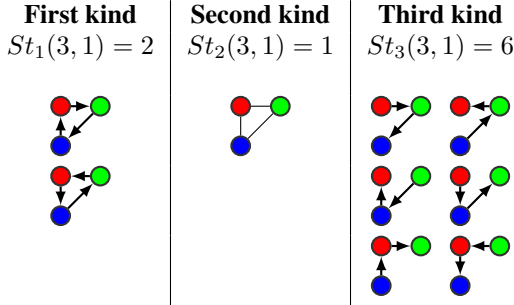


Figure 1: Illustration of the Stirling numbers of the three kinds for  $y = 3$  and  $n = 1$ .

## 2 Proof of limit cases

**Proposition 1.** *If there exists  $\theta^{raw}$  such that  $\lim_{\theta \rightarrow \theta^{raw}} \kappa^T \psi(\theta) = -\infty$ , then the posterior of dcPF tends to the posterior of PF as  $\theta$  goes to  $\theta^{raw}$ .*

**Proposition 2.** *If there exists  $\theta^{bin}$  such that  $\lim_{\theta \rightarrow \theta^{bin}} \kappa^T \psi(\theta) = +\infty$ , then the posterior of dcPF tends to the posterior of PF applied to binarized data as  $\theta$  goes to  $\theta^{bin}$ , i.e.:  $\lim_{\theta \rightarrow \theta^{bin}} p(\mathbf{W}, \mathbf{H} | \mathbf{Y}) = p(\mathbf{W}, \mathbf{H} | \mathbf{N} = \mathbf{Y}^b)$ .*

*Proof.* Let  $\lambda \in \mathbb{R}_+$ ,  $n \sim \text{Poisson}(\lambda)$  and  $y|n \sim ED(\theta, n\kappa)$  with support given by  $S = \{n, \dots, +\infty\}$ :

$$p(n|\lambda) = \frac{\lambda^n e^{-\lambda}}{n!}, \quad (1)$$

$$p(y|n) = \exp(y\theta - n\kappa^T \psi(\theta)) h(y, n\kappa), \quad y \in S, \quad (2)$$

where  $\kappa$  and  $\psi(\theta)$  can either be scalars or vectors of the same dimension. In both cases,  $\kappa^T \psi(\theta) \in \mathbb{R}$ . We denote by  $r = \lambda e^{-\kappa^T \psi(\theta)}$ .

We have the following posterior distribution for  $y > 0$ :

$$p(n|y) = \frac{r^n h(y, n\kappa) (n!)^{-1}}{\sum_{m=1}^y r^m h(y, m\kappa) (m!)^{-1}}, \quad n \in \{1, \dots, y\}. \quad (3)$$

Thus, for fixed  $\kappa$  and  $y > 0$ , we have that:

$$\sum_{m=1}^y r^m h(y, m\kappa) (m!)^{-1} \underset{r \rightarrow +\infty}{\sim} r^y h(y, y\kappa) (y!)^{-1} \quad (4)$$

$$\underset{r \rightarrow 0}{\sim} r h(y, \kappa). \quad (5)$$

It follows:

$$p(n|y) \xrightarrow{r \rightarrow +\infty} \delta_y(n) \quad (6)$$

$$p(n|y) \xrightarrow{r \rightarrow 0} \delta_1(n). \quad (7)$$

From these results we can deduce that, in dcPF, assuming:

- there exists  $\theta^{raw}$  such that  $\lim_{\theta \rightarrow \theta^{raw}} \kappa^T \psi(\theta) = -\infty$ ,
- there exists  $\theta^{bin}$  such that  $\lim_{\theta \rightarrow \theta^{bin}} \kappa^T \psi(\theta) = +\infty$ .

Then, we have the following limit cases:

$$\begin{aligned}
p(\mathbf{N}|\mathbf{Y}) &= \int_{\mathbf{W}, \mathbf{H}} p(\mathbf{N}|\mathbf{Y}, \mathbf{W}, \mathbf{H})p(\mathbf{W}, \mathbf{H}|\mathbf{Y})d\mathbf{W}d\mathbf{H} \\
&\xrightarrow{\theta \rightarrow \theta^{\text{raw}}} \int_{\mathbf{W}, \mathbf{H}} \delta_{\mathbf{Y}}(\mathbf{N}) p(\mathbf{W}, \mathbf{H}|\mathbf{Y})d\mathbf{W}d\mathbf{H} = \delta_{\mathbf{Y}}(\mathbf{N}) \\
&\xrightarrow{\theta \rightarrow \theta^{\text{bin}}} \int_{\mathbf{W}, \mathbf{H}} \delta_{\mathbf{Y}^b}(\mathbf{N}) p(\mathbf{W}, \mathbf{H}|\mathbf{Y})d\mathbf{W}d\mathbf{H} = \delta_{\mathbf{Y}^b}(\mathbf{N}).
\end{aligned} \tag{8}$$

And finally, for the posterior distribution:

$$p(\mathbf{W}, \mathbf{H}|\mathbf{Y}) = \int_{\mathbf{N}} p(\mathbf{W}, \mathbf{H}|\mathbf{N})p(\mathbf{N}|\mathbf{Y})d\mathbf{N} \tag{9}$$

$$\xrightarrow{\theta \rightarrow \theta^{\text{raw}}} p(\mathbf{W}, \mathbf{H}|\mathbf{N} = \mathbf{Y}) \tag{10}$$

$$\xrightarrow{\theta \rightarrow \theta^{\text{bin}}} p(\mathbf{W}, \mathbf{H}|\mathbf{N} = \mathbf{Y}^b), \tag{11}$$

where  $p(\mathbf{W}, \mathbf{H}|\mathbf{N})$  is the posterior of a PF model with raw or binarized observations respectively.  $\square$

### 3 Adaptivity of dcPF to over-dispersion

Table 1: Mean, variance and ratio var/mean of the non-zero values for each dataset. Learned parameters for each model and each dataset.

	<b>Taste Profile</b>	<b>NIPS</b>	<b>Last.fm</b>
mean of non-zeros	2.66	2.74	3.86
var of non-zeros	25.94	20.87	65.72
ratio var/mean	9.8	7.6	17.0
Log - $p$	0.80	0.74	0.90
ZTP - $p$	1.95	1.40	2.35
Geo - $p$	0.60	0.51	0.69
sh. NB - $p$	0.87	0.86	0.90
sh. NB - $\kappa_2$	0.21	0.17	0.27

Table 1 illustrates how the natural parameter  $\theta = \log(p)$  is strongly correlated to the variance-mean ratio of the non-zero values of the datasets. Hence, it illustrates the adaptivity of dcPF to over-dispersion.