Data-Efficient Reinforcement Learning with Probabilistic Model Predictive Control

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1 Appendix

1.1 Lipschitz Continuity

Lemma 1. The moment matching mapping f_{MM} is Lipschitz continuous for controls defined over a compact set \mathcal{U} .

Proof: Lipschitz continuity requires that the gradient $\partial f_{MM}/\partial u_t$ is bounded. The gradient is

$$\frac{\partial f_{MM}}{\partial \boldsymbol{u}_t} = \frac{\partial \boldsymbol{z}_{t+1}}{\partial \boldsymbol{u}_t} = \left[\frac{\partial \boldsymbol{\mu}_{t+1}}{\partial \boldsymbol{u}_t}, \frac{\partial \boldsymbol{\Sigma}_{t+1}}{\partial \boldsymbol{u}_t} \right] . \tag{1}$$

The derivatives $\left[\frac{\partial \mu_{t+1}}{\partial u_t}, \frac{\partial \Sigma_{t+1}}{\partial u_t}\right]$ can be computed analytically [1].

We first show that the derivative $\partial \mu_{t+1}/\partial u_t$ is bounded. Defining $\boldsymbol{\beta}_d := (\boldsymbol{K}_d + \sigma_{f_d}^2 \boldsymbol{I})^{-1} \boldsymbol{y}_d$, from [1], we obtain for all state dimensions $d = 1, \ldots, D$

$$\mu_{t+1}^d = \sum_{i=1}^N \beta_{d_i} q_{d_i} \,, \tag{2}$$

$$q_{di} = \sigma_{f_d}^2 | \boldsymbol{I} + \boldsymbol{L}_d^{-1} \widetilde{\boldsymbol{\Sigma}}_t |^{-\frac{1}{2}} \times \tag{3}$$

$$\exp\left(-\frac{1}{2}(\widetilde{\boldsymbol{x}}_i - \widetilde{\boldsymbol{\mu}}_t)^T (\boldsymbol{L}_d + \widetilde{\boldsymbol{\Sigma}}_t)^{-1} (\widetilde{\boldsymbol{x}}_i - \widetilde{\boldsymbol{\mu}}_t)\right),$$
(4

where N is the size of the training set of the dynamics GP and \tilde{x}_i the *i*th training input. The corresponding gradient w.r.t. u_t is given by the last F elements of

$$\frac{\partial \mu_{t+1}^d}{\partial \widetilde{\boldsymbol{\mu}}_t} = \sum_{i=1}^N \beta_{d_i} \frac{\partial q_{d_i}}{\partial \widetilde{\boldsymbol{\mu}}_t}$$

$$= \sum_{i=1}^N \beta_{d_i} q_{d_i} (\widetilde{\boldsymbol{x}}_i - \widetilde{\boldsymbol{\mu}}_t)^T (\widetilde{\boldsymbol{\Sigma}}_t + \boldsymbol{L}_d)^{-1} \in \mathbb{R}^{1 \times (D+F)}$$
(6)

Let us examine the individual terms in the sum on the rhs in (6): For a given trained GP $\|\beta_d\| < \infty$ is constant. The definition of q_{d_i} in (4) contains an exponentiated negative quadratic term, which is bounded between [0, 1]. Since $I + L_d^{-1} \widetilde{\Sigma}_t$ is positive definite, the inverse determinant is defined and bounded. Finally,

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 $\sigma_{f_d}^2 < \infty$, which makes $q_{d_i} < \infty$. The remaining term in (6) is a vector-matrix product. The matrix is regular and its inverse exists and is bounded (and constant as a function of u_t . Since $u_t \in \mathcal{U}$ where \mathcal{U} is compact, we can also conclude that the vector difference in (6) is finite, which overall proves that f_{MM} is (locally) Lipschitz continuous and Lemma 1.

1.2 Sequential Quadratic Programming

We can use SQP for solving non-linear optimization problems (NLP) of the form,

$$\min_{u} f(\mathbf{x})$$
s.t. $b(\mathbf{x}) \ge 0$

$$c(\mathbf{x}) = 0$$

The Lagrangian \mathcal{L} associated with the NLP is

$$\mathcal{L}(\boldsymbol{x}, \boldsymbol{\lambda}, \boldsymbol{\sigma}) = f(\boldsymbol{x}) - \boldsymbol{\sigma}^T b(\boldsymbol{x}) - \boldsymbol{\lambda}^T c(\boldsymbol{x})$$
 (7)

where, λ and σ are Lagrange multipliers. Sequential Quadratic Programming (SQP) forms a quadratic (Taylor) approximation of the objective and linear approximation of constraints at each iteration k

$$\min_{\mathbf{d}} \quad f(\mathbf{x}_k) + \nabla f(\mathbf{x}_k)^T \mathbf{d} + \frac{1}{2} \mathbf{d}^T \nabla_{xx}^2 \mathcal{L}(\mathbf{x}, \boldsymbol{\lambda}, \boldsymbol{\sigma}) \mathbf{d}
\text{s.t.} \quad b(\mathbf{x}_k) + \nabla b(\mathbf{x}_k)^T \mathbf{d} \ge 0
 c(\mathbf{x}_k) + \nabla c(\mathbf{x}_k)^T \mathbf{d} = 0.$$
(8)

The Lagrange multipliers λ associated with the equality constraint are same as the ones defined in the control Hamiltonian \mathcal{H} ??. The Hessian matrix ∇^2_{xx} can be computed by exploiting the block diagonal structure introduced by the Hamiltonian [? ?].

Moment Matching Approximation [1]

Following the law of iterated expectations, for target dimensions $a = 1, \ldots, D$, we obtain the predictive mean

$$\mu_{t}^{a} = \mathbb{E}_{\tilde{\boldsymbol{x}}_{t-1}}[\mathbb{E}_{f_{a}}[f_{a}(\tilde{\boldsymbol{x}}_{t-1})|\tilde{\boldsymbol{x}}_{t-1}]] = \mathbb{E}_{\tilde{\boldsymbol{x}}_{t-1}}[m_{f_{a}}(\tilde{\boldsymbol{x}}_{t-1})]$$

$$= \int m_{f_{a}}(\tilde{\boldsymbol{x}}_{t-1})\mathcal{N}(\tilde{\boldsymbol{x}}_{t-1}|\tilde{\boldsymbol{\mu}}_{t-1},\tilde{\boldsymbol{\Sigma}}_{t-1})d\tilde{\boldsymbol{x}}_{t-1}$$

$$= \boldsymbol{\beta}_{a}^{T}\boldsymbol{q}_{a}, \qquad (9)$$

$$\boldsymbol{\beta}_a = (\boldsymbol{K}_a + \sigma_{w_a}^2)^{-1} \boldsymbol{y}_a \tag{10}$$

with $\mathbf{q}_a = [q_{a_1}, \dots, q_{a_n}]^T$. The entries of $\mathbf{q}_a \in \mathbb{R}^n$ are computed using standard results from multiplying and integrating over Gaussians and are given by

$$q_{a_{i}} = \int k_{a}(\tilde{\boldsymbol{x}}_{i}, \tilde{\boldsymbol{x}}_{t-1}) \mathcal{N}(\tilde{\boldsymbol{x}}_{t-1} | \tilde{\boldsymbol{\mu}}_{t-1}, \tilde{\boldsymbol{\Sigma}}_{t-1}) d\tilde{\boldsymbol{x}}_{t-1}$$

$$= \sigma_{f_{a}}^{2} |\tilde{\boldsymbol{\Sigma}}_{t-1} \boldsymbol{\Lambda}_{a}^{-1} + \boldsymbol{I}|^{-\frac{1}{2}} \exp\left(-\frac{1}{2}\boldsymbol{\nu}_{i}^{T}(\tilde{\boldsymbol{\Sigma}}_{t-1} + \boldsymbol{\Lambda}_{a})^{-1}\boldsymbol{\nu}_{i}\right) \text{with } \boldsymbol{\nu}_{i} \text{ taken from (12). Hence, the off-diagonal entries}$$

$$= \sigma_{f_{a}}^{2} |\tilde{\boldsymbol{\Sigma}}_{t-1} \boldsymbol{\Lambda}_{a}^{-1} + \boldsymbol{I}|^{-\frac{1}{2}} \exp\left(-\frac{1}{2}\boldsymbol{\nu}_{i}^{T}(\tilde{\boldsymbol{\Sigma}}_{t-1} + \boldsymbol{\Lambda}_{a})^{-1}\boldsymbol{\nu}_{i}\right) \text{of } \boldsymbol{\Sigma}_{t} \text{ are fully determined by (9)-(12)}$$

$$= \sigma_{f_{a}}^{2} |\tilde{\boldsymbol{\Sigma}}_{t-1} \boldsymbol{\Lambda}_{a}^{-1} + \boldsymbol{I}|^{-\frac{1}{2}} \exp\left(-\frac{1}{2}\boldsymbol{\nu}_{i}^{T}(\tilde{\boldsymbol{\Sigma}}_{t-1} + \boldsymbol{\Lambda}_{a})^{-1}\boldsymbol{\nu}_{i}\right) \text{of } \boldsymbol{\Sigma}_{t} \text{ are fully determined by (9)-(12)}$$

where we define

$$\boldsymbol{\nu}_i \coloneqq (\tilde{\boldsymbol{x}}_i - \tilde{\boldsymbol{\mu}}_{t-1}) \tag{12}$$

is the difference between the training input \tilde{x}_i and the mean of the test input distribution $p(\boldsymbol{x}_{t-1}, \boldsymbol{u}_{t-1})$.

Computing the predictive covariance matrix $\Sigma_t \in$ $\mathbb{R}^{D \times D}$ requires us to distinguish between diagonal elements and off-diagonal elements: Using the law of total (co-)variance, we obtain for target dimensions a, b = 1, ..., D

$$\sigma_{aa}^{2} = \mathbb{E}_{\tilde{\boldsymbol{x}}_{t-1}} \left[\operatorname{var}_{f} \left[x_{t}^{a} | \tilde{\boldsymbol{x}}_{t-1} \right] \right] + \mathbb{E}_{f, \tilde{\boldsymbol{x}}_{t-1}} \left[(\boldsymbol{x}_{t}^{a})^{2} \right] - (\mu_{t}^{a})^{2},$$
(13)

$$\sigma_{ab}^2 = \mathbb{E}_{f, \tilde{x}_{t-1}}[x_t^a x_t^b] - \mu_t^a \mu_t^b, \quad a \neq b,$$
 (14)

respectively, where μ_t^a is known from (9). The offdiagonal terms do not contain the additional term $\mathbb{E}_{\tilde{\boldsymbol{x}}_{t-1}}[\operatorname{cov}_{f}[x_{t}^{a}, x_{t}^{b}|\tilde{\boldsymbol{x}}_{t-1}]]$ because of the conditional independence assumption of the GP models. Different target dimensions do not covary for given \tilde{x}_{t-1} .

We start the computation of the covariance matrix with the terms that are common to both the diagonal and the off-diagonal entries: With $p(\tilde{x}_{t-1}) =$ $\mathcal{N}(\tilde{x}_{t-1} | \tilde{\mu}_{t-1}, \Sigma_{t-1})$ and the law of iterated expectations, we obtain

$$\mathbb{E}_{f,\tilde{\boldsymbol{x}}_{t-1}}[x_t^a, x_t^b] = \mathbb{E}_{\tilde{\boldsymbol{x}}_{t-1}} \left[\mathbb{E}_f[x_t^a | \tilde{\boldsymbol{x}}_{t-1}] \, \mathbb{E}_f[x_t^b | \tilde{\boldsymbol{x}}_{t-1}] \right]$$
$$= \int m_f^a(\tilde{\boldsymbol{x}}_{t-1}) m_f^b(\tilde{\boldsymbol{x}}_{t-1}) p(\tilde{\boldsymbol{x}}_{t-1}) d\tilde{\boldsymbol{x}}_{t-1} \quad (15)$$

because of the conditional independence of x_t^a and x_t^b given \tilde{x}_{t-1} . Using the definition of the mean function, we obtain

$$\mathbb{E}_{f,\tilde{\boldsymbol{x}}_{t-1}}[x_t^a x_t^b] = \boldsymbol{\beta}_a^T \boldsymbol{Q} \boldsymbol{\beta}_b, \qquad (16)$$

$$\boldsymbol{Q} \coloneqq \int k_a(\tilde{\boldsymbol{x}}_{t-1}, \boldsymbol{X})^T k_b(\tilde{\boldsymbol{x}}_{t-1}, \boldsymbol{X}) p(\tilde{\boldsymbol{x}}_{t-1}) d\tilde{\boldsymbol{x}}_{t-1}. \qquad (17)$$

Using standard results from Gaussian multiplications and integration, we obtain the entries Q_{ij} of $\mathbf{Q} \in \mathbb{R}^{n \times n}$

$$Q_{ij} = \frac{k_a(\tilde{\boldsymbol{x}}_i, \tilde{\boldsymbol{\mu}}_{t-1})k_b(\tilde{\boldsymbol{x}}_j, \tilde{\boldsymbol{\mu}}_{t-1})}{\sqrt{|\boldsymbol{R}|}} \exp\left(\frac{1}{2}\boldsymbol{z}_{ij}^T \boldsymbol{T}^{-1} \boldsymbol{z}_{ij}\right)$$
(18)

where we define

$$egin{aligned} m{R} \coloneqq & ilde{m{\Sigma}}_{t-1}(m{\Lambda}_a^{-1} + m{\Lambda}_b^{-1}) + m{I} \,, \quad m{T} \coloneqq m{\Lambda}_a^{-1} + m{\Lambda}_b^{-1} + ilde{m{\Sigma}}_{t-1}^{-1} \,, \ m{z}_{ij} \coloneqq m{\Lambda}_a^{-1} m{
u}_i + m{\Lambda}_b^{-1} m{
u}_j \,, \end{aligned}$$

of Σ_t are fully determined by (9)–(12), (14), (16), (17), and (18).

References

[1] M. P. Deisenroth, D. Fox, and C. E. Rasmussen. Gaussian Processes for Data-Efficient Learning in Robotics and Control. Transactions on Pattern Analysis and Machine Intelligence, 37(2):408–23, 2015.