

P R I N C I P A L W O R K S
O F
A . K . E R L A N G

1. THE THEORY OF PROBABILITIES AND
TELEPHONE CONVERSATIONS

First published in "Nyt Tidsskrift for Matematik" B, Vol. 20 (1909), p. 33.

Although several points within the field of Telephony give rise to problems, the solution of which belongs under the Theory of Probabilities, the latter has not been utilized much in this domain, so far as can be seen. In this respect the Telephone Company of Copenhagen constitutes an exception as its managing director, Mr. *F. Johannsen*, through several years has applied the methods of the theory of probabilities to the solution of various problems of practical importance; also, he has incited others to work on investigations of similar character. As it is my belief that some point or other from this work may be of interest, and as a special knowledge of telephonic problems is not at all necessary for the understanding thereof, I shall give an account of it below.

1. The probability of a certain number of calls being originated during a certain interval of time.

It is assumed that there is no greater probability of a call being attempted at one particular moment than at any other moment. Let a be the time interval given, n the average number of calls during the unit of time. We will find the probability S_0 of 0 calls being originated during the time a , and afterwards the probability S_x of exactly x calls being originated during the time a . As $\frac{na}{r}$ is the probability of calls during the time $\frac{a}{r}$ when r is infinitely great, $1 - \frac{na}{r}$ is, on the same assumption, the probability of 0 attempts being made during the time $\frac{a}{r}$. Hence we have

$$S_0 = \lim_{r=\infty} \left(1 - \frac{na}{r}\right)^r = e^{-na}. \quad (\text{I})$$

Now in order to find S_x , the time a can be divided into r equal elements where $r \geq x$, and x of these elements chosen, which can be done in $C_{r,x}$

ways. We are now seeking, firstly, the probability of 1 call being originated during each of the x elements; secondly, the probability of no attempts being made during the remaining time, which is $\frac{a(r-x)}{r}$. The former probability is, for $r = \infty$, $\left(\frac{na}{r}\right)^x$; the latter is according to (I)

$$e^{-\frac{na(r-x)}{r}}$$

Thus, we get $S_x = \lim_{r=\infty} C_{r,x} \left(\frac{na}{r}\right)^x e^{-\frac{na(r-x)}{r}}$, or, as $\lim_{r=\infty} \frac{C_{r,x}}{r^x} = \frac{1}{x!}$,

$$S_x = \frac{(na)^x}{x!} e^{-na}. \quad (\text{II})$$

This formula becomes less complicated if we let m denote na , the average number of calls arriving during the time a . Then we have

$$S_x = \frac{m^x}{x!} e^{-m}. \quad (\text{III})$$

2. The Law of Distribution.

When, in the formula thus found, x is allowed to assume the values of all whole numbers from 0 upwards, the formula will express a certain "law of error", or "law of distribution". It is at once obvious that the sum of all the probabilities is 1, as it should be; further, that the probabilities of an even number and of an odd number of calls are, respectively,

$$\frac{e^m + e^{-m}}{2e^m} \quad \text{and} \quad \frac{e^m - e^{-m}}{2e^m}.$$

The chief property of the law of distribution is that all "half-invariants" are equal to m (*T. N. Thiele: Theory of Observations, London, 1903*); here, I shall confine myself to showing that the mean square error is m . We get,

$$\begin{aligned} & \left(\frac{1^2 m}{1!} + \frac{2^2 m^2}{2!} + \frac{3^2 m^3}{3!} + \frac{4^2 m^4}{4!} + \dots \right) e^{-m} - m^2 \\ &= \left(\frac{m^2}{1!} + \frac{2 m^3}{2!} + \frac{3 m^4}{3!} + \dots + m + \frac{m^2}{1!} + \frac{m^3}{2!} + \frac{m^4}{3!} + \frac{m^5}{4!} + \dots \right) e^{-m} - m^2 \\ &= (m^2 e^m + m e^m) e^{-m} - m^2 = m. \end{aligned}$$

The simple suppositions leading to the simple formula (III) will not, of course, always be satisfied in practice. Let us suppose, *e. g.*, that a business firm has certain busy days every week corresponding to a mean value m_1 , and certain less busy days corresponding to a mean value m_2 . Let the busy portion of the week be p_1 , and the less busy portion p_2 , where $p_1 + p_2 = 1$. If it is desired here to express the variations in the number of calls from day to day in terms of one single law of distribution, we find that the mean value is

$$p_1 m_1 + p_2 m_2;$$

but the mean square error is

$$p_1 p_2 (m_1 - m_2)^2$$

greater than the mean value. As, however, the preceding simple theory proves, on the whole, to be corresponding fairly well with meter readings experienced, we shall stick to that in the following.

3. Delay in answering of telephone calls.

We will assume that each operator receives calls from a determinate group of subscribers only, the system being designed in such a way that she cannot get help from her neighbours even if she is occupied and they happen to be free at the moment. By choosing a suitable unit of time, we can fix the average at 1 call per unit of time. The establishing of a connexion lasts t units of time. If a call is originated while the operator is unoccupied, we shall here consider the delay in answering, or waiting time, as being non-existent (actually, a certain short space of time will pass before the signal is noticed). If, on the other hand, she is occupied with another call, then the calling subscriber will have to wait a certain time. The problem is now to determine the function $f(z)$, $f(z)$ representing the probability of the waiting time not exceeding z .

The probability that, at the moment a call arrives, the time having elapsed since the preceding call should be confined within the limits

$$y \text{ and } y + dy,$$

is $e^{-y} dy$. The probability that the waiting time of the preceding call has been less than $z + y - t$, is $f(z + y - t)$.

Hence, we obtain

$$f(z) = \int_{y=0}^{\infty} f(z + y - t) e^{-y} dy.$$

By differentiation with respect to z , this equation gives

$$f'(z) = \int_{y=0}^{\infty} f'(z + y - t) e^{-y} dy,$$

and by partial integration,

$$f(z) = f(z - t) + \int_{y=0}^{\infty} f'(z + y - t) e^{-y} dy.$$

Thus we have

$$f'(z) = f(z) - f(z - t). \tag{IV}$$

By integration, $f(z)$ can now be determined in a succession of intervals, on the assumption that $f(z) = 0$ for $z < 0$, jumps from 0 to $1 - t$ for $z = 0$, but varies continuously for all other values of z .

For

$0 < z < t,$
$t < z < 2t,$
$2t < z < 3t,$
$3t < z < 4t,$
$\dots \dots \dots$
$\dots \dots \dots$
$nt < z < (n + 1)t$

the results will then be, respectively,

$$f(z) = (1 - t) e^z$$

$$f(z) = (1 - t) (e^z + e^{z-t} (t - z))$$

$$\dots \dots \dots$$

$$f(z) = (1 - t) \left(e^z + \frac{e^{z-t} (t - z)}{1!} + \frac{e^{z-2t} (2t - z)^2}{2!} \dots \dots \frac{e^{z-nt} (nt - z)^n}{n!} \right)$$

This is easily proved by inserting in equation (IV) the value of $f(z)$ taken from the last of the above formulae, and the value of $f(z - t)$ taken from the last but one.

As to the numerical calculation, it will be advantageous to begin with making a table of the function

$$\frac{m^x}{x!} e^{-m}$$

for negative values of m (e. g. with intervals of 0.1), and for positive and integral values of x ; for this purpose, one of the existing tables of $\log x!$ will be a good help. The oldest of these — which is, also, still the best — is *C. F. Degen: Tabularum enneas* (Havniæ, 1824). The terms in the table of the function must then be summed along oblique lines, and the obtained sums multiplied by $1 - t$, to provide the values for the definitive table giving $f(z)$ as a function of z and t . A table of this kind is printed below (table 1).

Table 1. Values of $f(z)$.

$t \backslash z$	0.0	0.1	0.2	0.3	0.4	0.5	0.6	0.7	0.8	0.9	1.0
0.0	1.000	0.900	0.800	0.700	0.600	0.500	0.400	0.300	0.200	0.100	0.000
0.1	1.000	0.995	0.884	0.774	0.663	0.553	0.442	0.332	0.221	0.111	0.000
0.2	1.000	1.000	0.977	0.855	0.733	0.611	0.489	0.366	0.244	0.122	0.000
0.3	1.000	1.000	0.991	0.945	0.810	0.675	0.540	0.405	0.270	0.135	0.000
0.4	1.000	1.000	0.998	0.967	0.895	0.746	0.597	0.448	0.298	0.149	0.000
0.5	1.000	1.000	0.999	0.983	0.923	0.824	0.659	0.495	0.330	0.165	0.000
0.6	1.000	1.000	1.000	0.992	0.947	0.856	0.729	0.547	0.364	0.182	0.000
0.7	1.000	1.000	1.000	0.996	0.965	0.885	0.761	0.605	0.403	0.201	0.000
0.8	1.000	1.000	1.000	0.998	0.977	0.910	0.792	0.635	0.445	0.223	0.000
0.9	1.000	1.000	1.000	0.999	0.984	0.931	0.822	0.665	0.470	0.246	0.000
1.0	1.000	1.000	1.000	0.999	0.990	0.947	0.849	0.694	0.495	0.261	0.000
1.1	1.000	1.000	1.000	1.000	0.993	0.958	0.872	0.722	0.520	0.276	0.000
1.2	1.000	1.000	1.000	1.000	0.995	0.967	0.891	0.749	0.545	0.292	0.000
1.3	1.000	1.000	1.000	1.000	0.997	0.975	0.906	0.773	0.569	0.307	0.000
1.4	1.000	1.000	1.000	1.000	0.998	0.980	0.920	0.794	0.592	0.323	0.000
1.5	1.000	1.000	1.000	1.000	0.999	0.985	0.932	0.812	0.614	0.339	0.000
1.6	1.000	1.000	1.000	1.000	0.999	0.988	0.942	0.829	0.635	0.354	0.000
1.7	1.000	1.000	1.000	1.000	1.000	0.991	0.950	0.845	0.653	0.369	0.000
1.8	1.000	1.000	1.000	1.000	1.000	0.993	0.957	0.859	0.671	0.384	0.000
1.9	1.000	1.000	1.000	1.000	1.000	0.994	0.964	0.872	0.688	0.397	0.000
2.0	1.000	1.000	1.000	1.000	1.000	0.996	0.969	0.884	0.705	0.411	0.000

4. In order to facilitate the understanding of the preparation of the final table giving the values of $f(z)$, I shall give, in table 2, the values of the Poisson function

$$\frac{m^x}{x!} e^{-m}$$

for negative values of m . Incidentally, the values of $f(z)$ can be obtained in a different manner by means of a table giving the values of the said function for x being positive. In this case it is necessary, as previously mentioned, to add up all the terms placed along oblique lines, or diagonals; then, the number of terms to be considered is infinite, but most often convergence will be very rapid.

By multiplication by $1 - \alpha$ is thus obtained directly, not the probability $f(z)$ of an inferior delay in answering, but the probability $1 - f(z)$ of a superior delay for a fixed value of z . This is easily proved by means of a theorem by *J. L. W. V. Jensen**), according to which $\frac{1}{1 - \alpha}$ is equal to the sums of the terms situated along oblique lines in a complete Poisson table, i. e. one comprising positive as well as negative values of m . As to denotations, I have here used α , the symbol employed by Mr. *Johannsen* who was the first to discuss theoretically the important question of delays in answering**).

For great values of z , an approximated (asymptotic) formula may be employed which simplifies the calculation:

$$S = 1 - f(z) = \frac{1 - \alpha}{\alpha' - 1} e^{-z \frac{\alpha' - \alpha}{\alpha}},$$

where the figures α and α' ($\alpha < 1 < \alpha'$) are bound by the relation

$$\alpha e^{-\alpha} = \alpha' e^{-\alpha'}.$$

Finally, the average delay can be given the simple expression

$$M = \frac{\alpha}{2(1 - \alpha)}.$$

It should be remembered that it is an essential presupposition for the results stated above that the calls be of constant duration.

*) *Acta mathematica*, XXVI, 1902, p. 309.

***) *The Post Office Electrical Engineers' Journal*, October, 1910, p. 244, and January, 1911, p. 303.

