

Towards Tight Security Bounds for OMAC, XCBC and TMAC

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Abstract. OMAC — a single-keyed variant of CBC-MAC by Iwata and Kurosawa — is a widely used and standardized (NIST FIPS 800-38B, ISO/IEC 29167-10:2017) message authentication code (MAC) algorithm. The best security bound for OMAC is due to Nandi who proved that OMAC’s pseudorandom function (PRF) advantage is upper bounded by $O(q^2\ell/2^n)$, where n , q , and ℓ , denote the block size of the underlying block cipher, the number of queries, and the maximum permissible query length (in terms of n -bit blocks), respectively. In contrast, there is no attack with matching lower bound. Indeed, the best known attack on OMAC is the folklore birthday attack achieving a lower bound of $\Omega(q^2/2^n)$. In this work, we close this gap for a large range of message lengths. Specifically, we show that OMAC’s PRF security is upper bounded by $O(q^2/2^n + q\ell^2/2^n)$. In practical terms, this means that for a 128-bit block cipher, and message lengths up to 64 Gigabyte, OMAC can process up to 2^{64} messages before rekeying (same as the birthday bound). In comparison, the previous bound only allows 2^{48} messages. As a side-effect of our proof technique, we also derive similar tight security bounds for XCBC (by Black and Rogaway) and TMAC (by Kurosawa and Iwata). As a direct consequence of this work, we have established tight security bounds (in a wide range of ℓ) for all the CBC-MAC variants, except for the original CBC-MAC.

Keywords: OMAC, CMAC, XCBC, TMAC, CBC-MAC, PRF, tight security

1 Introduction

Message Authentication Code (or, MAC) algorithms are symmetric-key primitives which are used for data authenticity and integrity. The sender generates a short tag based on message and a secret key which can be recomputed by any authorized receiver. MACs are commonly designed either based on a hash function or a block cipher. CBC-MAC is a block cipher-based MAC (message authentication code) which is based on the CBC mode of operation invented by

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Ehrsam et al. [1]. Given an n -bit block cipher E instantiated with a key K , the CBC-MAC construction is defined recursively as follows: for any $x \in \{0, 1\}^n$, $\text{CBC}_{E_K}(x) := E_K(x)$. For all $m = (m[1], \dots, m[\ell]) \in (\{0, 1\}^n)^\ell$ where $\ell \geq 2$, we define

$$\text{CBC}_{E_K}(m) := E_K(\text{CBC}_{E_K}(m[1], \dots, m[\ell - 1]) \oplus m[\ell]) \quad (1)$$

It was an international standard, and has been proven secure for fixed-length messages or prefix-free message spaces (i.e., no message is a prefix to another message). Simple length extension attacks prohibit its usage for arbitrary length messages. However, appropriately chosen operations to process the last block can resist these attacks. One such idea was first applied in EMAC [2,3], where the CBC-MAC output was encrypted using an independently keyed block cipher. It worked for all messages with lengths that are divisible by the block size of the underlying block cipher. Black and Rogaway proposed [4] three-keyed constructions, ECBC, FCBC, and XCBC, which are proven to be secure against adversaries querying arbitrary length messages. Later, in back-to-back works, Iwata and Kurosawa proposed two improved constructions (in terms of the key size), namely, TMAC [5] that uses two keys, and OMAC¹ [7] that requires just a single key. Nandi proposed [8] GCBC1 and GCBC2, a slight improvement over OMAC in terms of the number of block cipher calls for multi-block messages.

1.1 Related Works and Motivation

It is well-established [9] that the security of any deterministic MAC can be quantified via the pseudorandom function (or PRF²) security. Consequently, most of the works on CBC-MAC variants analyze their PRF security. For constructions like ECBC, FCBC and EMAC, Pietrzak [10] showed a PRF bound of $O(q^2/2^n)$ for $\ell < 2^{n/8}$, where q and ℓ denote the number of messages and the maximum permissible length (no. of n -bit blocks) of the messages. Later, Jha and Nandi [11] discovered a flaw in the proof of the earlier bound and showed a bound of $O(q/2^{n/2})$ up to $\ell < 2^{n/4}$. However, in these constructions an extra (independent) block cipher is called at the end. Considering the number of block cipher calls, XCBC, TMAC and OMAC are better choices. XCBC uses two independent masking keys for the last block which are used depending on whether the last block is padded or not. In case of TMAC, the two masking keys are derived from a single n -bit key. OMAC optimized the key derivation further. Here, both the keys are derived using the underlying block cipher itself. Thus, it is much better in this respect. Classical bound for these constructions was $O(\sigma^2/2^n)$ [4,5], σ being the total number blocks among all the messages. Later, in a series of work [12,13,14,15], the improved bounds for XCBC, TMAC, and OMAC were shown to be in the form of $O(q^2\ell/2^n)$, $O(\sigma^2/2^n)$ and $O(\sigma q/2^n)$. Interestingly, it has also

¹ This is same as CMAC [6] — a NIST recommended AES based MAC — for appropriate choice of constants.

² A keyed construction is called a PRF if it is computationally infeasible to distinguish it from a random function.

been shown in [16] that if we use a PRF, instead of a block cipher in these constructions, there is an attack with roughly $\Omega(q^2\ell/2^n)$ advantage, which is tight. No such attack is known in the presence of a block cipher. This gives an implicit motivation to study the exact security of these constructions in the presence of block ciphers. In this paper, we aim to show birthday-bound security for these block cipher based MACs for a suitable range of message lengths.

In a different paradigm but with similar motivations, recently Chattopadhyay et al. [17] showed birthday-bound security for another standardized MAC called LightMAC [18]. However, similar result for original PMAC [19] is still an open problem (although a result is available for its variant in [20]). In addition to the improved bound for LightMAC, Chattopadhyay et al. proposed a new proof approach called the reset-sampling method. They also hinted (via a very brief discussion) that this method could be useful for proving better security for OMAC. However, the discussion in [17] is overly simplistic and contains no formal analysis of bad events. Indeed, the reset-sampling is more involved than anticipated in [17], giving rise to some crucial and tricky bad events (see section 4). To their credit, they do say that

A more formal and rigorous analysis of OMAC using reset-sampling will most probably require handling of several other bad events, and could be an interesting future research topic.

In this paper, we take up this topic and give a complete and rigorous analysis.

1.2 Our Contributions

In section 3, we show that the PRF advantages for OMAC, XCBC and TMAC are upper bounded by $O(q^2/2^n) + O(q\ell^2/2^n)$, which is almost tight in terms of the number of queries q while $\ell \ll 2^{n/4}$. This bound is not exactly the birthday bound $O(q^2/2^n)$, but for any fixed target advantage, in terms of the limit on q it behaves almost like the birthday bound for a fairly good range of ℓ (see the following discussion). The proof of our security bound is given in section 4 and follows the recently introduced reset-sampling approach [17]. These improved bounds, in combination with previous results [11,21] for EMAC, ECBC and FCBC, completely characterize (see Table 1.1) the security landscape of CBC-MAC variants for message lengths up to $2^{n/4}$ blocks.

A NOTE ON THE TIGHTNESS AND IMPROVEMENT IN BOUNDS: In Figure 1.3, we present a graph³ comparing the best known bound for OMAC [15], i.e., $B_1(\ell, q) = 10q^2\ell/2^n$, the ideal birthday bound, i.e., $B_{\text{id}} = q^2/2^n$, and the bound shown in this paper (see Theorem 3.1), i.e., $B_2(\ell, q) \approx \frac{16q^2}{2^n} + \frac{2q\ell^2}{2^n}$ (as the remaining terms are dominated by these two terms). In the graph, we show the trade-off curve for the parameters $X = \log \ell$ and $Y = \log q$, where \log denotes “log base 2”, for a fixed choice of advantage value, say $\epsilon = 2^{-a}$ for some $a \in \mathbb{N}$.

³ Using GeoGebra Classic available at <https://www.geogebra.org/classic>

Table 1.1: Summary of security (PRF advantage) bounds for the CBC-MAC family. Here n , q , ℓ , and σ denote the block size, number of queries, maximum permissible message length, and sum of message lengths of all q queries, respectively.

Scheme	State-of-the-art		This paper	
	Bound	Restriction	Bound	Restriction
CBC-MAC [1]	$O(\sigma q/2^n)$ [11,21]	$\ell = o(2^{n/3})$	-	-
EMAC [2,3]	$O(q^2/2^n + q\ell^2/2^n)$ [11,21]	-	-	-
ECBC,FCBC [4]	$O(q^2/2^n + q\ell^2/2^n)$ [11,21]	-	-	-
XCBC [4], TMAC [5]	$O(q^2\ell/2^n)$ [12] ¹	$\ell = o(2^{n/3})$	$O(q^2/2^n + q\ell^2/2^n)$	-
	$O(\sigma^2/2^n)$ [13] ¹	-		
OMAC [7]	$O(\sigma q/2^n)$ [15]	$\ell = o(2^{n/3})$	$O(q^2/2^n + q\ell^2/2^n)$	-

¹ σ^2 and $q^2\ell$ are incomparable, as they depend on the query length distribution.

Let $n_a := n - a$. Then, we have

$$B_{\text{id}} : Y = \frac{n_a}{2} \quad B_1 : X + 2Y = n_a - \log 10 \quad B_2 : \log(16 \cdot 2^{2Y} + 2 \cdot 2^{2X+Y}) = n_a.$$

Looking at the equation related to the bound B_2 we can see that it is actually a combination of two linear equations: $2Y = n_a - 4$ and $2X + Y = n_a - 1$, the choice depending on whether $16q^2/2^n$ or $2q\ell^2/2^n$ dominates. Precisely, the curve expressing the relation between $\log \ell$ and $\log q$ in B_2 is $\{(X, Y) : X \leq n/4, Y = \min\{(n_a - 4)/2, n_a - 1 - 2X\}\}$. From the above linear equations two important facts about the curve related to B_2 can be noticed:

- It remains very close to the straight line corresponding to B_{id} from $(0, \frac{n_a-4}{2})$ to $(\frac{n_a+2}{4}, \frac{n_a-4}{2})$ and then moves downward.
- At around $(\frac{n_a+1}{3}, \frac{n_a-5}{3})$ it starts to degrade below the curve related to B_1 .

For example, if we take $(n, a) = (128, 32)$, the bound proved in this paper is very close to the birthday bound for $\ell \leq 2^{25}$ and even after degrading, it remains better than the bound in [15] till $\ell \leq 2^{32}$. Moreover, if we take $(n, a) = (128, 64)$, q remains 2^{30} until $\ell \leq 2^{16}$ and degrades below the existing bound only after $\ell \geq 2^{22}$. Thus, if we consider the advantage in general terms, we can always take the minimum among the advantage proved in this paper and that proved in [15].

2 Preliminaries

For $n \in \mathbb{N}$, $[n]$ and (n) denote the sets $\{1, 2, \dots, n\}$ and $\{0\} \cup [n]$, respectively. The set of all bit strings (including the empty string \perp) is denoted $\{0, 1\}^*$. The length

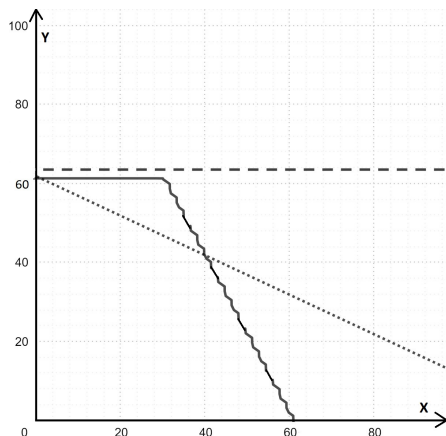
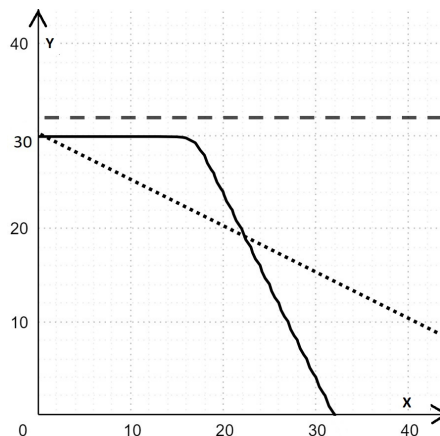
Fig. 1.1: For $\epsilon = 2^{-1}$ Fig. 1.2: For $\epsilon = 2^{-64}$

Fig. 1.3: $(\log \ell, \log q)$ -Trade-off Graph for the bounds of OMAC. For $n = 128$, and two different choices of the target advantage, $\epsilon = 2^{-1}$ (on the left), and $\epsilon = 2^{-64}$ (on the right), the above graphs show the relation between $X = \log \ell$ and $Y = \log q$. The *dashed*, *dotted* and *continuous* curves represent the equations B_{id} , B_1 , and B_2 , respectively.

of any bit string $x \in \{0, 1\}^*$, denoted $|x|$, is the number of bits in x . For $n \in \mathbb{N}$, $\{0, 1\}^n$ denotes the set of all bit strings of length n , and $\{0, 1\}^{\leq n} := \bigcup_{i=0}^n \{0, 1\}^i$. For $x, y \in \{0, 1\}^*$, $z = x||y$ denotes the concatenation of x and y . Additionally, x (resp. y) is called the *prefix* (resp. *suffix*) of z . For $x, y \in \{0, 1\}^*$, let $\text{Prefix}(x, y)$ denote the length of the largest possible common prefix of x and y . For $1 \leq k \leq n$, we define the falling factorial $(n)_k := n!/(n-k)! = n(n-1)\cdots(n-k+1)$. Any pair of q -tuples $\tilde{x} = (x_1, \dots, x_q)$ and $\tilde{y} = (y_1, \dots, y_q)$, are said to be *permutation compatible*, denoted $\tilde{x} \rightsquigarrow \tilde{y}$, if $(x_i = x_j) \iff (y_i = y_j)$, for all $i \neq j$. By an abuse of notation, we also use \tilde{x} to denote the set $\{x_i : i \in [q]\}$ for any \tilde{x} .

2.1 Security Definitions

DISTINGUISHERS: A (q, T) -distinguisher \mathcal{A} is an oracle Turing machine, that makes at most q oracle queries, runs in time at most T , and outputs a single bit. For any oracle \mathcal{O} , we write $\mathcal{A}^{\mathcal{O}}$ to denote the output of \mathcal{A} after its interaction with \mathcal{O} . By convention, $T = \infty$ denotes computationally unbounded (information-theoretic) and deterministic distinguishers. In this paper, we assume that the distinguisher is non-trivial, i.e., it never makes a duplicate query. Let $\mathbb{A}(q, T)$ be the class of all non-trivial distinguishers limited to q queries and T computations.

Primitives and Their Security: The set of all functions from \mathcal{X} to \mathcal{Y} is denoted $\mathcal{F}(\mathcal{X}, \mathcal{Y})$, and the set of all permutations of \mathcal{X} is denoted $\mathcal{P}(\mathcal{X})$. We

simply write $\mathcal{F}(a, b)$ and $\mathcal{P}(a)$, whenever $\mathcal{X} = \{0, 1\}^a$ and $\mathcal{Y} = \{0, 1\}^b$. For a finite set \mathcal{X} , $\mathbf{X} \leftarrow_{\$} \mathcal{X}$ denotes the uniform at random sampling of \mathbf{X} from \mathcal{X} .

PSEUDORANDOM FUNCTION: A $(\mathcal{K}, \mathcal{X}, \mathcal{Y})$ -keyed function F with key space \mathcal{K} , domain \mathcal{X} , and range \mathcal{Y} is a function $F : \mathcal{K} \times \mathcal{X} \rightarrow \mathcal{Y}$. We write $F_k(x)$ for $F(k, x)$.

The *pseudorandom function* or PRF advantage of any distinguisher \mathcal{A} against a $(\mathcal{K}, \mathcal{X}, \mathcal{Y})$ -keyed function F is defined as

$$\mathbf{Adv}_F^{\text{prf}}(\mathcal{A}) = \mathbf{Adv}_{F; \Gamma}(\mathcal{A}) := \left| \Pr_{\mathbf{K} \leftarrow_{\$} \mathcal{K}} (\mathcal{A}^{F_{\mathbf{K}}} = 1) - \Pr_{\Gamma \leftarrow_{\$} \mathcal{F}(\mathcal{X}, \mathcal{Y})} (\mathcal{A}^{\Gamma} = 1) \right|. \quad (2)$$

The *PRF insecurity* of F against $\mathbb{A}(q, T)$ is defined as

$$\mathbf{Adv}_F^{\text{prf}}(q, T) := \max_{\mathcal{A} \in \mathbb{A}(q, T)} \mathbf{Adv}_F^{\text{prf}}(\mathcal{A}).$$

PSEUDORANDOM PERMUTATION: For some $n \in \mathbb{N}$, a $(\mathcal{K}, \mathcal{B})$ -block cipher E with key space \mathcal{K} and block space $\mathcal{B} := \{0, 1\}^n$ is a $(\mathcal{K}, \mathcal{B}, \mathcal{B})$ -keyed function, such that $E(k, \cdot)$ is a permutation over \mathcal{B} for any key $k \in \mathcal{K}$. We write $E_k(x)$ for $E(k, x)$.

The *pseudorandom permutation* or PRP advantage of any distinguisher \mathcal{A} against a $(\mathcal{K}, \mathcal{B})$ -block cipher E is defined as

$$\mathbf{Adv}_E^{\text{prp}}(\mathcal{A}) = \mathbf{Adv}_{E; \Pi}(\mathcal{A}) := \left| \Pr_{\mathbf{K} \leftarrow_{\$} \mathcal{K}} (\mathcal{A}^{E_{\mathbf{K}}} = 1) - \Pr_{\Pi \leftarrow_{\$} \mathcal{P}(n)} (\mathcal{A}^{\Pi} = 1) \right|. \quad (3)$$

The *PRP insecurity* of E against $\mathbb{A}(q, T)$ is defined as

$$\mathbf{Adv}_E^{\text{prp}}(q, T) := \max_{\mathcal{A} \in \mathbb{A}(q, T)} \mathbf{Adv}_E^{\text{prp}}(\mathcal{A}).$$

2.2 H-coefficient Technique

Let \mathcal{A} be a computationally unbounded and deterministic distinguisher that's trying to distinguish the real oracle \mathcal{O}_1 from the ideal oracle \mathcal{O}_0 . The collection of all queries and responses that \mathcal{A} made and received to and from the oracle, is called the *transcript* of \mathcal{A} , denoted as ν . Let \mathbf{V}_1 and \mathbf{V}_0 denote the transcript random variable induced by \mathcal{A} 's interaction with \mathcal{O}_1 and \mathcal{O}_0 , respectively. Let \mathcal{V} be the set of all transcripts. A transcript $\nu \in \mathcal{V}$ is said to be *attainable* if $\Pr(\mathbf{V}_0 = \nu) > 0$, i.e., it can be realized by \mathcal{A} 's interaction with \mathcal{O}_0 .

Following these notations, we state the main result of the so-called H-coefficient technique [22, 23] in Theorem 2.1. A proof of this result is available in [23].

Theorem 2.1 (H-coefficient). *For $\epsilon_1, \epsilon_2 \geq 0$, suppose there is a set $\mathcal{V}_{\text{bad}} \subseteq \mathcal{V}$, referred as the set of all bad transcripts, such that the following conditions hold:*

- $\Pr(\mathbf{V}_0 \in \mathcal{V}_{\text{bad}}) \leq \epsilon_1$; and
- For any $\nu \in \mathcal{V} \setminus \mathcal{V}_{\text{bad}}$, ν is attainable and $\frac{\Pr(\mathbf{V}_1 = \nu)}{\Pr(\mathbf{V}_0 = \nu)} \geq 1 - \epsilon_2$.

Then, for any computationally unbounded and deterministic distinguisher \mathcal{A} , we have

$$\mathbf{Adv}_{\mathcal{O}_1; \mathcal{O}_0}(\mathcal{A}) \leq \epsilon_1 + \epsilon_2.$$

Reset-Sampling Method: In H-coefficient based proofs, often we release additional information to the adversary in order to make it easy to define the bad transcripts. In such scenarios, one has to define how this additional information is sampled, and naturally the sampling mechanism is construction specific. The reset-sampling method [17] is a sampling philosophy, within this highly mechanized setup of H-coefficient technique, where some of the variables are reset/resampled (hence the name) depending upon the consistency requirement for the overall transcript. We employ this sampling approach in our proof.

3 The CBC-MAC Family

Throughout, n denotes the *block size*, $\mathcal{B} := \{0, 1\}^n$, and any $x \in \mathcal{B}$ is referred as a *block*. For any non-empty $m \in \{0, 1\}^*$, $(m[1], \dots, m[\ell_m]) \stackrel{n}{\leftarrow} m$ denotes the *block parsing* of m , where $|m[i]| = n$ for all $1 \leq i \leq \ell_m - 1$ and $1 \leq |m[\ell_m]| \leq n$. In addition, we associate a boolean flag δ_m to each $m \in \{0, 1\}^*$, which is defined as

$$\delta_m := \begin{cases} -1 & \text{if } |m| = n\ell_m, \\ 0 & \text{otherwise.} \end{cases}$$

For any $m \in \{0, 1\}^{\leq n}$, we define

$$\bar{m} := \begin{cases} m \| 10^{n-|m|-1} & \text{if } |m| < n, \\ m & \text{otherwise.} \end{cases}$$

CBC FUNCTION: The CBC function, based on a permutation⁴ $\pi \in \mathcal{P}(n)$, takes as input a non-empty message $m \in \mathcal{B}^*$ and computes the output $\text{CBC}_\pi(m) := y_m^\pi[\ell_m]$ inductively as described below:

$y_m^\pi[0] = 0^n$ and for $1 \leq i \leq \ell_m$, we have

$$\begin{aligned} x_m^\pi[i] &:= y_m^\pi[i-1] \oplus m[i], \\ y_m^\pi[i] &:= \pi(x_m^\pi[i]), \end{aligned} \tag{4}$$

where $(m[1], \dots, m[\ell_m]) \stackrel{n}{\leftarrow} m$. For empty message, we define the CBC output as the constant 0^n . Figure 3.1 illustrates the evaluation of CBC function over a 4-block message m .

Given the definition of CBC_π , one can easily define all the variants of CBC-MAC. Here, we define XCBC, TMAC and OMAC—the three constructions that we study in this paper.

⁴ Instantiated with a block cipher in practical applications.

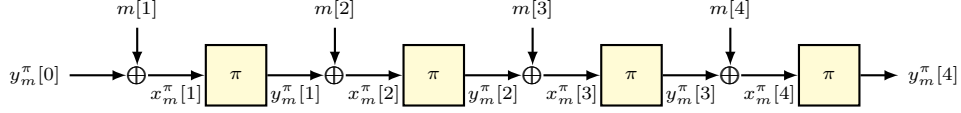


Fig. 3.1: Evaluation of CBC function over a 4-block message m .

XCBC: The XCBC algorithm is a three-key construction, based on a permutation $\pi \in \mathcal{P}(n)$ and keys $(L_{-1}, L_0) \in \mathcal{B}^2$, that takes as input a non-empty message $m \in \{0, 1\}^*$, and computes the output

$$\text{XCBC}_{\pi, L_{-1}, L_0}(m) := t = \pi \left(\text{CBC}_{\pi}(m^*) \oplus \overline{m[\ell_m]} \oplus L_{\delta_m} \right), \quad (5)$$

where $(m[1], \dots, m[\ell_m]) \stackrel{n}{\leftarrow} m$, and $m^* := m[1] \parallel \dots \parallel m[\ell_m - 1]$.

TMAC: The TMAC algorithm is a two-key construction, based on a permutation $\pi \in \mathcal{P}(n)$ and key $L \in \mathcal{B}$, that takes as input a non-empty message $m \in \{0, 1\}^*$, and computes the output

$$\text{TMAC}_{\pi, L}(m) := t = \pi \left(\text{CBC}_{\pi}(m^*) \oplus \overline{m[\ell_m]} \oplus \mu_{\delta_m} \odot L \right), \quad (6)$$

where $(m[1], \dots, m[\ell_m]) \stackrel{n}{\leftarrow} m$, $m^* := m[1] \parallel \dots \parallel m[\ell_m - 1]$, μ_{-1} and μ_0 are constants chosen from $\text{GF}(2^n)$ (viewing \mathcal{B} as $\text{GF}(2^n)$), such that $\mu_{-1}, \mu_0, 1 \oplus \mu_{-1}, 1 \oplus \mu_0$ are all distinct and not equal to either 0 or 1, and \odot denotes the field multiplication operation over $\text{GF}(2^n)$ with respect to a fixed primitive polynomial. For the sake of uniformity, we define $L_{\delta_m} := \mu_{\delta_m} \odot L$ in context of TMAC.

OMAC: The OMAC algorithm is a single-keyed construction, based on a permutation $\pi \in \mathcal{P}(n)$, that takes as input a non-empty message $m \in \{0, 1\}^*$, and computes the output

$$\text{OMAC}_{\pi}(m) := t = \pi \left(\text{CBC}_{\pi}(m^*) \oplus \overline{m[\ell_m]} \oplus \mu_{\delta_m} \odot \pi(0^n) \right), \quad (7)$$

where $(m[1], \dots, m[\ell_m]) \stackrel{n}{\leftarrow} m$, $m^* := m[1] \parallel \dots \parallel m[\ell_m - 1]$, μ_{-1} and μ_0 are constants chosen analogously as in the case of TMAC. For the sake of uniformity, we define $L_{\delta_m} := \mu_{\delta_m} \odot \pi(0^n)$ in context of OMAC.

Input and Output Tuples: In the context of CBC evaluation within OMAC, we refer to $x_m^{\pi} := (x_m^{\pi}[1], \dots, x_m^{\pi}[\ell_m - 1])$ and $y_m^{\pi} := (y_m^{\pi}[0], \dots, y_m^{\pi}[\ell_m - 1])$ as the *intermediate input* and *output* tuples, respectively, associated to π and m . We define the final input variable as $x_m^{\pi}[\ell_m] := y_m^{\pi}[\ell_m - 1] \oplus \overline{m[\ell_m]} \oplus \mu_{\delta_m} \odot \pi(0^n)$. Clearly, the input and output tuples (including the final input) are well defined for OMAC. Analogous definitions are possible (and useful in proof) for XCBC and TMAC as well. It is worth noting that the intermediate input tuple x_m^{π} is uniquely determined by the intermediate output tuple y_m^{π} and the message m , and it is independent of the permutation π . Going forward, we drop π from the notations, whenever it is clear from the context.

3.1 Tight Security Bounds for OMAC, XCBC and TMAC

The main technical result of this paper, given in Theorem 3.1, is a tight security bound for OMAC for a wide range of message lengths. The proof of this theorem is postponed to section 4. In addition, we also provide similar result for XCBC and TMAC in Theorem 3.2. We skip the proof since it is almost identical to the one for Theorem 3.1, and has slightly less relevance given that a more efficient and standardized algorithm OMAC already achieves similar security. In what follows we define

$$\begin{aligned} \epsilon'(q, \ell) := & \frac{16q^2 + q\ell^2}{2^n} + \frac{8q^2\ell^4 + 32q^3\ell^2 + 2q^2\ell^3}{2^{2n}} \\ & + \frac{3q^3\ell^5 + 143q^3\ell^6 + 11q^4\ell^3}{2^{3n}} + \frac{17q^4\ell^6 + 5462q^4\ell^8}{2^{4n}}. \end{aligned}$$

Theorem 3.1 (OMAC bound). *Let $q, \ell, \sigma, T > 0$. For $q + \sigma \leq 2^{n-1}$, the PRF insecurity of OMAC, based on block cipher E_K , against $\mathbb{A}(q, T)$ is given by*

$$\mathbf{Adv}_{\text{OMAC}_{E_K}}^{\text{prf}}(q, \ell, \sigma, T) \leq \mathbf{Adv}_E^{\text{prp}}(q + \sigma, T') + \frac{4\sigma}{2^n} + \epsilon'(q, \ell), \quad (8)$$

where q denotes the number of queries, ℓ denotes an upper bound on the number of blocks per query, σ denotes the total number of blocks present in all q queries, $T' = T + \sigma O(T_E)$ and T_E denotes the runtime of E .

Theorem 3.2 (XCBC-TMAC bound). *Let $q, \ell, \sigma, T > 0$. For $q + \sigma \leq 2^{n-1}$, the PRF insecurity of XCBC and TMAC, based on block cipher E_K and respective masking keys (L, L_{-1}, L_0) , against $\mathbb{A}(q, T)$ is given by*

$$\mathbf{Adv}_{\text{XCBC}_{E_K, L_{-1}, L_0}}^{\text{prf}}(q, \ell, \sigma, T) \leq \mathbf{Adv}_E^{\text{prp}}(q + \sigma, T') + \epsilon'(q, \ell) \quad (9)$$

$$\mathbf{Adv}_{\text{TMAC}_{E_K, L}}^{\text{prf}}(q, \ell, \sigma, T) \leq \mathbf{Adv}_E^{\text{prp}}(q + \sigma, T') + \epsilon'(q, \ell) \quad (10)$$

where q denotes the number of queries, ℓ denotes an upper bound on the number of blocks per query, σ denotes the total number of blocks present in all q queries, $T' = T + \sigma O(T_E)$ and T_E denotes the runtime of E .

Proof of this theorem is almost same as that of Theorem 3.1. The bad event on a collision on zero block input is redundant and hence dropped here. Rest of the proof remains the same and so we skip the details.

Remark 3.1. Note that the actual advantage cannot exceed 1. Let us denote $\frac{q^2}{2^n} = \alpha$ and $\frac{q\ell^2}{2^n} = \beta$. Looking at $\epsilon(q, \ell)$ (where $\epsilon(q, \ell) = \epsilon'(q, \ell) + \frac{4\sigma}{2^n}$ in case of OMAC and $\epsilon(q, \ell) = \epsilon'(q, \ell)$ in case of XCBC, TMAC), we see that any term in the expression is upper bounded by $c \cdot \alpha^s \beta^t$ for some constant c and $s, t \geq 0$ such that at least one of s and t is at least 1. As we can assume both α, β to be less than 1, each $\alpha^s \beta^t$ will be less than or equal to α or β . Thus, the above

PRF-advantage expressions for $\text{MAC} \in \{\text{OMAC}, \text{XCBC}, \text{TMAC}\}$ can be written as

$$\mathbf{Adv}_{\text{MAC}}^{\text{prf}}(q, \ell, \sigma) = O\left(\frac{q^2}{2^n}\right) + O\left(\frac{q\ell^2}{2^n}\right).$$

Indeed, under the assumption that $\ell \leq 2^{n/4-0.5}$ and $q \leq 2^{n/2-1}$, one can simplify the above bounds to $20q^2/2^n + 23q\ell^2/2^n$.

A NOTE ON THE PROOF APPROACH: In the analysis of OMAC, XCBC and TMAC, we have to handle the case that the final input collides with some intermediate input, the so-called *full collision* event. In earlier works the probability of this event is shown to be $q^2\ell/2^n$ (as there are less than $q\ell$ many intermediate inputs and q final inputs and any such collision happens with roughly $1/2^n$ probability). So, in a way they avoid handling this tricky event by disallowing it all together. In this work, we allow full collisions as long as the next intermediate input is not colliding with some other input (intermediate or final). Looking ahead momentarily, this is captured in **BadW3**. We can do this via the application of reset-sampling, resulting in a more amenable $(q^2/2^n + q\ell^2/2^n)$ bound.

4 Proof of Theorem 3.1

First, using the standard hybrid argument, we get

$$\mathbf{Adv}_{\text{OMAC}_{E_K}}^{\text{prf}}(q, \ell, \sigma, T) \leq \mathbf{Adv}_E^{\text{prf}}(q + \sigma, T') + \mathbf{Adv}_{\text{OMAC}_{\Pi}}^{\text{prf}}(q, \ell, \sigma, \infty). \quad (11)$$

Now, it is sufficient to bound $\mathbf{Adv}_{\text{OMAC}_{\Pi}}^{\text{prf}}(q, \ell, \sigma, \infty)$, where the corresponding distinguisher \mathcal{A} is computationally unbounded and deterministic. To bound this term, we employ the H-coefficient technique (see section 2.2), and the recently introduced *reset-sampling* method [17]. The remaining steps of the proof are given in the remainder of this section.

4.1 Oracle Description and Corresponding Transcripts

Real Oracle: The real oracle corresponds to OMAC_{Π} . It responds faithfully to all the queries made by \mathcal{A} . Once the query-response phase is over, it releases all the intermediate inputs and outputs, as well as the masking keys L_{-1} and L_0 to \mathcal{A} . We write $L = \Pi(0^n)$.

In addition, the real oracle releases three binary variables, namely, **FlagT**, **FlagW** and **FlagX**, all of which are degenerately set to 0. These flags are more of a technical requirement, and their utility will become apparent from the description of ideal oracle. For now, it is sufficient to note that these flags are degenerate in the real world.

Formally, we have $V_1 := (\tilde{M}, \tilde{T}, \tilde{X}, \tilde{X}^*, \tilde{Y}, L_{-1}, L_0, \text{FlagT}, \text{FlagW}, \text{FlagX})$, where

- $\tilde{M} = (M_1, \dots, M_q)$, the q -tuple of queries made by \mathcal{A} , where $M_i \in \{0, 1\}^*$ for all $i \in [q]$. In addition, for all $i \in [q]$, let $\ell_i := \left\lceil \frac{|M_i|}{n} \right\rceil$.

- $\tilde{\mathbf{T}} = (\mathsf{T}_1, \dots, \mathsf{T}_q)$, the q -tuple of final outputs received by \mathcal{A} , where $\mathsf{T}_i \in \mathcal{B}$.
- $\tilde{\mathbf{X}} = (\mathsf{X}_1, \dots, \mathsf{X}_q)$, where X_i denotes the intermediate input tuple for the i -th query.
- $\tilde{\mathbf{X}}^* = (\mathsf{X}_1[\ell_1], \dots, \mathsf{X}_q[\ell_q])$, where $\mathsf{X}_i[\ell_i]$ denotes the final input for the i -th query.
- $\tilde{\mathbf{Y}} = (\mathsf{Y}_1, \dots, \mathsf{Y}_q)$, where Y_i denotes the intermediate output tuple for the i -th query.
- L_{-1} and L_0 denote the two masking keys. Note that L_{-1} and L_0 are easily derivable from L . So we could have simply released L . The added redundancy is to aid the readers in establishing an analogous connection between this proof and the proof for XCBC and TMAC.
- $\mathsf{FlagT} = \mathsf{FlagW} = \mathsf{FlagX} = 0$.

From the definition of OMAC, we know that $\Pi(\mathsf{X}_i[a]) = \mathsf{Y}_i[a]$ for all $(i, a) \in [q] \times [\ell_i]$. So, *in the real world we always have* $(0^n, \tilde{\mathbf{X}}, \tilde{\mathbf{X}}^*) \rightsquigarrow (\mathsf{L}, \tilde{\mathbf{Y}}, \tilde{\mathbf{T}})$, *i.e.,* $(0^n, \tilde{\mathbf{X}}, \tilde{\mathbf{X}}^*)$ *is permutation compatible with* $(\mathsf{L}, \tilde{\mathbf{Y}}, \tilde{\mathbf{T}})$. We keep this observation in our mind when we simulate the ideal oracle.

Ideal Oracle: By reusing notations from the real world, we represent the ideal oracle transcript as $\mathsf{V}_0 := (\tilde{\mathbf{M}}, \tilde{\mathbf{T}}, \tilde{\mathbf{X}}, \tilde{\mathbf{X}}^*, \tilde{\mathbf{Y}}, \mathsf{L}_{-1}, \mathsf{L}_0, \mathsf{FlagT}, \mathsf{FlagW}, \mathsf{FlagX})$. This should not cause any confusion, as we never consider the random variables V_1 and V_0 jointly, whence the probability distributions of the constituent variables will always be clear from the context.

The ideal oracle transcript is described in three phases, each contingent on some predicates defined over the previous stages. Specifically, the ideal oracle first initializes $\mathsf{FlagT} = \mathsf{FlagW} = \mathsf{FlagX} = 0$, and then follows the sampling mechanism given below:

PHASE I (QUERY-RESPONSE PHASE): In the query-response phase, the ideal oracle faithfully simulates $\Gamma \leftarrow_{\mathcal{S}} \mathcal{F}(\{0, 1\}^*, \mathcal{B})$. Formally, for $i \in [q]$, at the i -th query $\mathsf{M}_i \in \{0, 1\}^*$, the ideal oracle outputs $\mathsf{T}_i \leftarrow_{\mathcal{S}} \mathcal{B}$. The partial transcript generated at the end of the query-response phase is given by $(\tilde{\mathbf{M}}, \tilde{\mathbf{T}})$, where

$$- \tilde{\mathbf{M}} = (\mathsf{M}_1, \dots, \mathsf{M}_q) \text{ and } \tilde{\mathbf{T}} = (\mathsf{T}_1, \dots, \mathsf{T}_q).$$

Now, we define a predicate on $\tilde{\mathbf{T}}$:

$$\mathsf{BadT} : \exists i \neq j \in [q], \text{ such that } \mathsf{T}_i = \mathsf{T}_j.$$

If BadT is true, then FlagT is set to 1, and $\tilde{\mathbf{X}}, \tilde{\mathbf{X}}^*$, and $\tilde{\mathbf{Y}}$ are defined degenerately: $\mathsf{X}_i[a] = \mathsf{Y}_i[b] = 0^n$ for all $i \in [q]$, $a \in [\ell_i]$, $b \in (\ell_i - 1)$. Otherwise, the ideal oracle proceeds to the next phase.

PHASE II (OFFLINE INITIAL SAMPLING PHASE): Onward, we must have $\mathsf{T}_i \neq \mathsf{T}_j$ whenever $i \neq j$, and $\mathsf{FlagT} = 0$, since this phase is only executed when BadT is false. In the offline phase, the ideal oracle's initial goal is to sample the input

and output tuples in such a way that the intermediate input and output tuples are permutation compatible. For now we use notations W and Z , respectively, instead of X and Y , to denote the input and output tuples. This is done to avoid any confusions in the next step where we may have to reset some of these variables. To make it explicit, W and Z respectively denote the input and output tuples before resetting, and X and Y denote the input and output tuples after resetting.

Let P be a key-value table representing a partial permutation of \mathcal{B} , which is initialized to empty, i.e., the corresponding permutation is undefined on all points. We write $P.\text{domain}$ and $P.\text{range}$ to denote the set of all keys and values utilized till this point, respectively. The ideal oracle uses this partial permutation P to maintain permutation compatibility between intermediate input and output tuples, in the following manner:

Initial sampling

```

 $L \leftarrow_{\$} \mathcal{B} \setminus \tilde{T}$ 
 $L_{-1} \leftarrow \mu_{-1} \odot L$ 
 $L_0 \leftarrow \mu_0 \odot L$ 
 $P(0^n) \leftarrow L$ 
for  $i = 1$  to  $q$  do
   $Z_i[0] \leftarrow 0^n$ 
  for  $a = 1$  to  $\ell_i - 1$  do
     $W_i[a] \leftarrow Z_i[a-1] \oplus M_i[a]$ 
    if  $W_i[a] \in P.\text{domain}$ 
       $Z_i[a] \leftarrow P(W_i[a])$ 
    else
       $Z_i[a] \leftarrow_{\$} \mathcal{B} \setminus (\tilde{T} \cup P.\text{range})$ 
       $P(W_i[a]) \leftarrow Z_i[a]$ 
   $W_i[\ell_i] \leftarrow Z_i[\ell_i - 1] \oplus \bar{M}_i[\ell_i] \oplus L_{\delta_{M_i}}$ 

```

At this stage we have $Z_i[a] = Z_j[b]$ if and only if $W_i[a] = W_j[b]$ for all $(i, a) \in [q] \times [\ell_i - 1]$ and $(j, b) \in [q] \times [\ell_j - 1]$. In other words, $(0^n, \tilde{W}) \rightsquigarrow (L, \tilde{Z})$. But it is obvious to see that the same might not hold between $(0^n, \tilde{W}, \tilde{W}^*)$ and $(L, \tilde{Z}, \tilde{T})$. In the next stage our goal will be to reset some of the Z variables in such a way that the resulting input tuple is compatible with the resulting output tuple. However, in order to reset, we have to identify and avoid certain contentious input-output tuples.

IDENTIFYING CONTENTIOUS INPUT-OUTPUT TUPLES: We define several predicates on (\tilde{W}, \tilde{W}^*) , each of which represents some undesirable property of the sampled input and output tuples.

First, observe that L is chosen outside the set \tilde{T} . This leads to the first predicate:

BadW1 : $\exists(i, a) \in [q] \times [\ell_i]$, such that $(W_i[a] = 0^n)$ and $(\ell_i > 1 \implies a > 1)$.

since, if **BadW1** is true, then $(0^n, \tilde{W}^*)$ is not compatible with (L, \tilde{T}) . In fact, \neg **BadW1** implies that none of the inputs, except the first input which is fully in adversary's control, can possibly be 0^n . This stronger condition will simplify the analysis greatly. The second predicate simply states that the final input tuple is not permutation compatible with the tag tuple, i.e., we have

BadW2 : $\exists i \neq j \in [q]$, such that $W_i[\ell_i] = W_j[\ell_j]$.

At this point, assuming $\neg(\text{BadW1} \vee \text{BadW2})$ holds true, the only way we can have permutation incompatibility is if $W_i[a] = W_j[\ell_j]$, for some $i, j \in [q]$ and $a \in [\ell_i - 1]$. A simple solution will be to reset $Z_i[a]$ to T_j , for all such (i, a, j) . In order to do this, we need that the following predicates must be false:

BadW3 : $\exists i, j, k \in [q], a \in [\ell_i - 1], b \in [\ell_k]$, such that

$$(W_i[a] = W_j[\ell_j]) \wedge (W_i[a+1] = W_k[b]) \wedge \text{Prefix}(M_i, M_k) < \max\{a+1, b\}.$$

BadW4 : $\exists i, j, k \in [q], a \neq b \in [\ell_i - 1]$, such that

$$(W_i[a] = W_j[\ell_j]) \wedge (W_i[b] = W_k[\ell_k]).$$

BadW5 : $\exists i, j, k \in [q], a \in [\ell_i - 1], b \in [\ell_j - 1]$, such that

$$(W_i[a] = W_j[\ell_j]) \wedge (W_j[b] = W_k[\ell_k]).$$

If **BadW3** is true, then once $Z_i[a]$ is reset, we lose the permutation compatibility since, the reset next input, i.e., $X_i[a+1] = W_i[a+1] \oplus Z_i[a] \oplus T_j = M_i[a+1] \oplus T_j \neq W_k[b]$ with high probability, whereas $Z_i[a+1] = Z_k[b]$ with certainty. **BadW4** simply represents the scenario where we may have to apply the initial resetting to two indices in a single message. Looking ahead momentarily, this may lead to contradictory *induced* resettings. Avoiding this predicate makes the resetting operation much more manageable. Similarly, avoiding **BadW5**, is just proactive prevention of contradictory resetting at $Z_i[a]$, since if **BadW5** occurs, then we may have a case where $X_j[\ell_j]$ is reset due to induced resetting, leading to the case, $X_i[a] \neq X_j[\ell_j]$ and $Y_i[a] = T_j$, where recall that $Y_i[a]$ is the resetting value of $Z_i[a]$. We write

$$\text{BadW} := \text{BadW1} \vee \text{BadW2} \vee \text{BadW3} \vee \text{BadW4} \vee \text{BadW5}.$$

If **BadW** is true, then **FlagW** is set to 1, and $(\tilde{X}, \tilde{X}^*, \tilde{Y})$ is again defined degenerately, as in the case of **BadT**. Otherwise, the ideal oracle proceeds to the next and the final phase, i.e., the resetting phase.

PHASE III.A INITIAL RESETTING PHASE: At this stage we must have $\neg(\text{BadT} \vee \text{BadW})$, i.e., $\text{FlagW} = \text{FlagT} = 0$. We describe the resetting phase in two sub-stages. First, we identify the indices affected by the initial resetting operation.

Definition 4.1 (full collision index). Any $(i, a, j) \in [q] \times [\ell_i - 1] \times [q]$ is called a full collision index (FCI) if $W_i[a] = W_j[\ell_j]$. Additionally, let

$$\begin{aligned} \text{FCI} &:= \{(i, a, j) : i, j \in [q], a \in [\ell_i - 1], \text{ such that } (i, a, j) \text{ is an FCI}\} \\ \widetilde{\text{FCI}} &:= \{(i, a) \in [q] \times [\ell_i - 1] : \exists j \in [q], \text{ such that } (i, a, j) \text{ is an FCI}\} \end{aligned}$$

The first sub-stage, executes a resetting for full collision indices in the following manner:

1. For all $(i, a, j) \in \text{FCI}$, define $Y_i[a] := T_j$;
2. For all $(i, a, j) \in \text{FCI}$, define

$$X_i[a + 1] := W_i[a + 1] \oplus Z_i[a] \oplus Y_i[a] = \overline{M}_i[a + 1] \oplus T_j \oplus 1_{a=\ell_i-1} \odot L_{\delta_{M_i}},$$

where $1_{a=\ell_i-1}$ is an indicator variable that evaluates to 1 when $a = \ell_i - 1$, and 0 otherwise.

Once the initial resetting is executed, it may result in new permutation incompatibilities. This necessitates further resettings, referred as *induced resettings*, which require that the following predicates are false:

BadX1 : $\exists(i, a, j) \in \text{FCI}, k \in [q], b \in [\ell_k] \setminus \{1\}$, such that

$$(X_i[a + 1] = W_k[b]) \vee (X_i[a + 1] = 0^n).$$

BadX2 : $\exists(i, a, j) \in \text{FCI}, k \in [q]$, such that

$$(X_i[a + 1] = M_k[1]) \wedge (M_i[a + 2, \dots, \ell_i] = M_k[2, \dots, \ell_k]).$$

BadX3 : $\exists(i, a, j), (k, b, l) \in \text{FCI}$, such that $(X_i[a + 1] = M_k[1])$.

BadX4 : $\exists(i, a, k), (j, b, l) \in \text{FCI}$, such that

$$(X_i[a + 1] = X_j[b + 1]) \wedge (\text{Prefix}(M_i, M_j) < \max\{a + 1, b + 1\}).$$

Here, the variable highlighted in red denotes the update after initial resetting. Let's review these predicates in slightly more details. First, **BadX1**, represents the situation where after resetting the next input (highlighted text) collides with some intermediate input or 0^n . This would necessitate induced resetting at $Z_i[a + 1]$. In other words, if **BadX1** is false then no induced resettings occur, unless the next input collides with some first block input. This case is handled in the next two predicates. **BadX2** represents the situation when the next input collides with a first block and the subsequent message blocks are all same. This would induce a chain of resetting going all the way to the final input. As **BadT** is false, this would immediately result in a permutation incompatibility since tags are distinct. If **BadX2** is false, then the chain of induced resetting must end at some point. **BadX3** is used to avoid circular or contradictory resettings. It is analogous

to BadW5 defined earlier. If it is false, then we know that the k -th message is free from resetting, so the induced resetting will be manageable. Finally, BadX4 represents the situation when two newly reset variables collide. We write

$$\text{BadX1234} := \text{BadX1} \vee \text{BadX2} \vee \text{BadX3} \vee \text{BadX4}$$

If BadX1234 is true, then FlagX is set to 1, and $(\tilde{X}, \tilde{X}^*, \tilde{Y})$ is again defined degenerately, as in the cases of BadT and BadW . Otherwise, the ideal oracle proceeds to the second and the final sub-stage of resetting.

PHASE III.B INDUCED RESETTING PHASE: Here, the goal is to execute the induced resettings necessitated by the initial resetting operation.

First, we define the *index of induced resetting* for each $(i, a) \in \widetilde{\text{FCI}}$, as the smallest index j such that $X_i[a+1] = M_j[1]$ and

$$\text{Prefix}(M_i[a+2, \dots, \ell_i], M_j[2, \dots, \ell_j]) = \max\{\text{Prefix}(M_i[a+2, \dots, \ell_i], M_{j'}[2, \dots, \ell_{j'}]) : j' \in [q]\},$$

i.e., $\text{Prefix}(M_i[a+2, \dots, \ell_i], M_j[2, \dots, \ell_j])$ maximizes.

Definition 4.2 (induced collision sequence). *A sequence of tuples $((i, a+1, j, 1), \dots, (i, a+p+1, j, p+1))$ is called an induced collision sequence (ICS), if $(i, a) \in \widetilde{\text{FCI}}$, and j is the index of induced resetting for (i, a) , where $p := \text{Prefix}(M_i[a+2, \dots, \ell_i], M_j[2, \dots, \ell_j])$. The individual elements of an ICS are referred as induced collision index (ICI). Additionally, we let*

$$\begin{aligned} \text{ICI} &:= \{(i, a, j, b) : i, j \in [q], a \in [\ell_i - 1], b \in [\ell_j - 1], \text{ and } (i, a, j, b) \text{ is an ICI.}\} \\ \widetilde{\text{ICI}} &:= \{(i, a) \in [q] \times [\ell_i - 1] : \exists (j, b) \in [q] \times [\ell_j - 1], \text{ and } (i, a, j, b) \text{ is an ICI.}\} \end{aligned}$$

Now, as anticipated, in the second sub-stage of resetting, we reset the induced collision indices in the following manner:

1. For all $(i, a, j, b) \in \text{ICI}$, define $Y_i[a] := Z_j[b]$;
2. For all $(i, a, j, b) \in \text{ICI}$, define

$$X_i[a+1] := W_i[a+1] \oplus Z_i[a] \oplus Y_i[a] = \overline{M}_i[a+1] \oplus Z_j[b] \oplus 1_{a=\ell_i-1} \odot L_{\delta_{M_i}},$$

where $1_{a=\ell_i-1}$ is an indicator variable that evaluates to 1 when $a = \ell_i - 1$, and 0 otherwise.

Given $\neg \text{BadX1234}$, we know that the induced resetting must stop at some point before the final input. Now, it might happen that once the first chain of induced resetting stops, the next input again collides which may result in nested resetting or permutation incompatibility. The predicates BadX5 , BadX6 , and BadX7 below represent these scenarios.

- $\text{BadX5} : \exists (i, a, k, b) \in \text{ICI}, l \in [q], b \in [\ell_l - 1]$, such that

$$(X_i[a+2+p] = W_i[b]) \vee (X_i[a+2+p] = 0^n),$$

where $p := \text{Prefix}(M_i[a+2, \dots, \ell_i], M_k[2, \dots, \ell_k])$.

- **BadX6** : $\exists(i, a) \in \widetilde{\text{FCI}}, (j, b, k, c) \in \text{ICI}$, such that $(\text{X}_i[a+1] = \text{X}_j[b+2+p])$, where $p := \text{Prefix}(\text{M}_j[b+2, \dots, \ell_j], \text{M}_k[2, \dots, \ell_k])$.
- **BadX7** : $\exists(i, a, k, c), (j, b, l, d) \in \text{ICI}$, such that

$$(\text{X}_i[a+2+p] = \text{X}_j[b+2+p']) \wedge (\text{Prefix}(\text{M}_i, \text{M}_j) < \max\{a+2+p, b+2+p'\}),$$
 where $p := \text{Prefix}(\text{M}_i[a+2, \dots, \ell_i], \text{M}_k[2, \dots, \ell_k])$, and $p' := \text{Prefix}(\text{M}_j[b+2, \dots, \ell_j], \text{M}_l[2, \dots, \ell_l])$.

Here, the variables highlighted in red and blue denote the update after initial resetting and induced resetting, respectively. These predicates are fairly self-explanatory. First **BadX5** represents the situation that the immediate input after induced resetting collides with some intermediate input or 0^n . This may cause permutation incompatibility and would lead to nested induced resetting at $Z_i[a+2+p]$. **BadX6** handles a similar collision with a full collision resetted variable, and **BadX7** handles the only remaining case where the immediate inputs after two different induced resetting collides. Note that, $\neg(\text{BadX5} \vee \text{BadX6} \vee \text{BadX7})$ would imply that for each message resetting stops at some point before the final input, and the next input is fresh.⁵ We write

$$\text{BadX} := \text{BadX1} \vee \text{BadX2} \vee \text{BadX3} \vee \text{BadX4} \vee \text{BadX5} \vee \text{BadX6} \vee \text{BadX7}.$$

If **BadX** is true, then **FlagX** is set to 1, and $(\widetilde{\text{X}}, \widetilde{\text{X}}^*, \widetilde{\text{Y}})$ is again defined degenerately, as in the case of **BadT** and **BadW**. Otherwise, for any remaining index $(i, a) \in [q] \times (\ell_i - 1) \setminus (\widetilde{\text{FCI}} \cup \widetilde{\text{ICI}})$, the ideal oracle resets as follows:

1. define $\text{Y}_i[a] := \text{Z}_i[a]$;
2. define $\text{X}_i[a+1] := \text{W}_i[a+1]$.

At this point, the ideal oracle transcript is completely defined. Intuitively, if the ideal oracle is not sampling $(\widetilde{\text{X}}, \widetilde{\text{X}}^*, \widetilde{\text{Y}})$ degenerately at any stage, then we must have $(0^n, \widetilde{\text{X}}, \widetilde{\text{X}}^*) \rightsquigarrow (\text{L}, \widetilde{\text{Y}}, \widetilde{\text{T}})$. The following proposition justifies this intuition.

Proposition 4.1. *For $\neg(\text{BadT} \vee \text{BadW} \vee \text{BadX})$, we must have $(0^n, \widetilde{\text{X}}, \widetilde{\text{X}}^*) \rightsquigarrow (\text{L}, \widetilde{\text{Y}}, \widetilde{\text{T}})$.*

Proof. Let $\neg(\text{BadT} \vee \text{BadW} \vee \text{BadX})$ hold. Recall that $(0^n, \widetilde{\text{W}}, \widetilde{\text{W}}^*)$ may not be permutation compatible with $(\text{L}, \widetilde{\text{Z}}, \widetilde{\text{T}})$. For any $(i, a) \in \widetilde{\text{FCI}}$, there exists $i' \in [q]$ such that $\text{W}_i[a] = \text{W}_{i'}[a]$ but $\text{Z}_i[a] \neq \text{T}_{i'}$. We apply the initial resetting to solve this issue. However, as a result of initial resetting, induced resetting takes place. Our goal is to show that the non-occurrence of the bad events assures that the compatibility is attained in the final reset tuples $(0^n, \widetilde{\text{X}}, \widetilde{\text{X}}^*)$ and $(\text{L}, \widetilde{\text{Y}}, \widetilde{\text{T}})$. We prove all possible cases as follows:

⁵ Does not collide with any other input.

- $X_i[a] = 0^n \iff Y_i[a] = L$: If $a = 1$ and $X_i[a] = 0$, then $(i, a) \notin \widetilde{\text{FCI}}$ due to $\neg\text{BadW1}$. Also, $(i, 1) \notin \widetilde{\text{ICI}}$. Thus, $Y_i[a] = Z_i[a] = L$ and the converse also holds. Otherwise, due to $\neg\text{BadX1}$, $X_i[a]$ can not be equal to 0. Also, due to $\neg\text{BadW1}$, $Y_i[a]$ can not be equal to L.
- $X_i[a] = X_{i'}[\ell_{i'}] \iff Y_i[a] = T_{i'}$: For $(i, a) \in \widetilde{\text{FCI}}$, this equivalence holds. Otherwise, $X_i[a] = X_{i'}[\ell_{i'}]$ can not hold due to $\neg(\text{BadX1} \vee \text{BadX5})$. Also $Y_i[a] = T_{i'}$ can not hold due to definition of \widetilde{T} and $\neg\text{BadX2}$.
- $X_i[a] = X_j[b] \iff Y_i[a] = Y_j[b]$: To prove this part we divide it in the following subcases:

- $(i, a), (j, b) \notin \widetilde{\text{FCI}} \cup \widetilde{\text{ICI}}$: Since in this case the variables are simply re-named due to definitions of resetting and $\neg\text{BadW3}$, the result follows from $\widetilde{W} \iff \widetilde{Z}$.
- $(i, a), (j, b) \in \widetilde{\text{FCI}}$: Since $(i, a), (j, b) \in \widetilde{\text{FCI}}$, there exists unique $i', j' \in [q]$, such that $W_i[a] = W_{i'}[\ell_{i'}]$ and $W_j[b] = W_{j'}[\ell_{j'}]$. Now, note that $X_i[a] = W_i[a]$ and $X_j[b] = W_j[b]$ since $\widetilde{\text{FCI}} \cap \widetilde{\text{ICI}} = \emptyset$ due to $\neg\text{BadW4}$; $W_{i'}[\ell_{i'}] = X_{i'}[\ell_{i'}]$ and $W_{j'}[\ell_{j'}] = X_{j'}[\ell_{j'}]$ due to $\neg\text{BadW5}$. Therefore, we must have $X_{j'}[\ell_{j'}] = W_{j'}[\ell_{j'}] = W_j[b] = X_j[b] = X_i[a] = W_i[a] = W_{i'}[\ell_{i'}] = X_{i'}[\ell_{i'}]$, which is possible if and only if $i' = j'$ (since $\neg\text{BadW2}$ holds).
- $(i, a), (j, b) \in \widetilde{\text{ICI}}$: Since $(i, a), (j, b) \in \widetilde{\text{ICI}}$, there exists $i', j' \in [q]$ and $a' \in [\ell_{i'} - 1], b' \in [\ell_{j'} - 1]$, such that $X_i[a] = W_{i'}[a']$ and $X_j[b] = W_{j'}[b']$. Further, $(i', a'), (j', b') \notin \widetilde{\text{FCI}} \cup \widetilde{\text{ICI}}$ (due to $\neg\text{BadX3}$). If $X_j[b] = X_i[a]$, then we have $W_{j'}[b'] = W_{i'}[a']$. This gives us $Y_j[b] = Z_{j'}[b'] = Z_{i'}[a'] = Y_i[a]$ (due to $\widetilde{W} \iff \widetilde{Z}$). Similarly, $X_i[a] \neq X_j[b]$ implies $Y_i[a] \neq Y_j[b]$.
- $(i, a) \in \widetilde{\text{FCI}}$ and $(j, b) \in \widetilde{\text{ICI}}$: Since $(i, a) \in \widetilde{\text{FCI}}$, there exists a unique $i' \in [q]$, such that $X_i[a] = W_i[a] = W_{i'}[\ell_{i'}] = X_{i'}[\ell_{i'}]$ (the first equality is due to $\neg\text{BadW4}$, the second equality is due to the definition of full collision, the third equality is due to $\neg\text{BadW5}$). Since $(j, b) \in \widetilde{\text{ICI}}$, we also have $X_j[b] = W_{j'}[b']$. If $X_i[a] = X_j[b]$, then $W_{j'}[b'] = W_{i'}[\ell_{i'}]$. Thus, $(j', b') = (i', \ell_{i'})$ due to $\neg\text{BadX3}$. Now, we have $Y_i[a] = T_{i'}$. Also, $Y_j[b] = Y_{j'}[b'] = Y_{i'}[\ell_{i'}] = T_{i'}$. Therefore, $Y_i[a] = Y_j[b]$. Moreover, $X_i[a] \neq X_j[b]$ implies that $Y_i[a] \neq Y_j[b]$ due to similar arguments as above and also $\neg\text{BadT}$.
- $(i, a) \in \widetilde{\text{ICI}}$ and $(j, b) \in \widetilde{\text{FCI}}$: Similar as the above case.
- $(i, a) \in \widetilde{\text{FCI}} \cup \widetilde{\text{ICI}}$ and $(j, b) \notin \widetilde{\text{FCI}} \cup \widetilde{\text{ICI}}$: Since $(j, b) \notin \widetilde{\text{FCI}} \cup \widetilde{\text{ICI}}$, we have $X_j[b] = W_j[b]$ and $Y_j[b] = Z_j[b]$. Suppose, $(i, a) \in \widetilde{\text{FCI}}$. Then $X_i[a] = X_j[b]$

is not possible since it would imply that $(j, b) \in \widetilde{\text{FCI}}$. Also, $Y_i[a] = Y_j[b]$ is not possible since it would contradict the definition of $\widetilde{\text{T}}$. Now, suppose, $(i, a) \in \widetilde{\text{ICI}}$. Therefore, $X_i[a] = W_{i'}[a']$ for some $i' \in [q]$ and $a' \in [\ell_{i'} - 1]$. If $X_i[a] = X_j[b]$, then $W_j[b] = X_j[b] = X_i[a] = W_{i'}[a']$. So, $Y_j[b] = Z_j[b] = Z_{i'}[a'] = Y_i[a]$. Similarly, $X_i[a] \neq X_j[b]$ implies $Y_i[a] \neq Y_j[b]$.

- $(i, a) \notin \widetilde{\text{FCI}} \cup \widetilde{\text{ICI}}$ and $(j, b) \in \widetilde{\text{FCI}} \cup \widetilde{\text{ICI}}$: Similar as the above case.

4.2 Transcript Analysis

SET OF TRANSCRIPTS: Given the description of transcript random variable corresponding to the ideal oracle, we can now define the set of transcripts \mathcal{V} as the set of all tuples $\nu = (\widetilde{m}, \widetilde{t}, \widetilde{x}, \widetilde{x}^*, \widetilde{y}, l_{-1}, l_0, \text{flagT}, \text{flagW}, \text{flagX})$, where

- $\widetilde{m} = (m_1, \dots, m_q)$, where $m_i \in \{0, 1\}^*$ for $i \in [q]$. Let $\ell_i = \left\lceil \frac{m_i}{n} \right\rceil$ for $i \in [q]$.
- $\widetilde{t} = (t_1, \dots, t_q)$, where $t_i \in \mathcal{B}$ for $i \in [q]$.
- $\widetilde{x} = (x_1, \dots, x_q)$, where $x_i = (x_i[1], \dots, x_i[\ell_i - 1])$ for $i \in [q]$.
- $\widetilde{x}^* = (x_1[\ell_1], \dots, x_q[\ell_q])$.
- $\widetilde{y} = (y_1, \dots, y_q)$, where $y_i = (y_i[0] = 0^n, y_i[1], \dots, y_i[\ell_i - 1])$ for $i \in [q]$.
- $l_{-1} = \mu_{-1} \odot l, l_0 = \mu_0 \odot l$ where $l \in \mathcal{B}$ and μ_{-1}, μ_0 are constants chosen from $\text{GF}(2^n)$ as defined before.
- $\text{flagT}, \text{flagW}, \text{flagX} \in \{0, 1\}$.

Furthermore, the following must always hold:

1. if $\text{flagI} = 1$ for some $I \in \{\text{T}, \text{W}\}$, then $x_i[a] = y_j[b] = 0^n$ for all $i, j \in [q]$, $a \in [\ell_i]$, and $b \in [\ell_j - 1]$.
2. if $\text{flagT} = 0$, then t_i 's are all distinct.
3. if $\text{flagI} = 0$ for all $I \in \{\text{T}, \text{W}, \text{X}\}$, then $x_i[a] = y_i[a - 1] \oplus \overline{m}_i[a]$ and $(0^n, \widetilde{x}, \widetilde{y}^\oplus) \leftrightarrow (L, \widetilde{y}, \widetilde{t})$.

The first two conditions are obvious from the ideal oracle sampling mechanism. The last condition follows from Proposition 4.1 and the observation that in ideal oracle sampling for any $I \in \{\text{T}, \text{Z}, \text{X}\}$, $\text{FlagI} = 1$ if and only if BadI is true. Note that, condition 3 is vacuously true for real oracle transcripts.

BAD TRANSCRIPT: A transcript $\nu \in \mathcal{V}$ is called *bad* if and only if the following predicate is true:

$$(\text{FlagT} = 1) \vee (\text{FlagW} = 1) \vee (\text{FlagX} = 1).$$

In other words, we term a transcript bad if the ideal oracle sets $(\widetilde{X}, \widetilde{X}^*, \widetilde{Y})$ degenerately. Let

$$\mathcal{V}_{\text{bad}} := \{\nu \in \mathcal{V} : \nu \text{ is bad.}\}.$$

All other transcript $\nu' = (\tilde{m}, \tilde{t}, \tilde{x}, \tilde{x}^*, \tilde{y}, l_{-1}, l_0, \text{flagT}, \text{flagW}, \text{flagX}) \in \mathcal{V} \setminus \mathcal{V}_{\text{bad}}$ are called *good*. From the preceding characterization of the set of transcripts, we conclude that for any good transcript ν' , we must have $(0^n, \tilde{x}, \tilde{x}^*) \rightsquigarrow (L, \tilde{y}, \tilde{t})$. Henceforth, we drop flagT , flagW , and flagX for any good transcript with an implicit understanding that $\text{flagT} = \text{flagW} = \text{flagX} = 0$.

Following the H-coefficient mechanism, we have to upper bound the probability $\Pr(\mathbf{V}_0 \in \mathcal{V}_{\text{bad}})$ and lower bound the ratio $\Pr(\mathbf{V}_1 = \nu) / \Pr(\mathbf{V}_0 = \nu)$ for any $\nu \in \mathcal{V} \setminus \mathcal{V}_{\text{bad}}$.

Lemma 4.1 (bad transcript analysis). *For $q + \sigma \leq 2^{n-1}$, we have*

$$\Pr(\mathbf{V}_0 \in \mathcal{V}_{\text{bad}}) \leq \frac{4\sigma}{2^n} + \frac{16q^2 + q\ell^2}{2^n} + \frac{8q^2\ell^4 + 32q^3\ell^2 + 2q^2\ell^3}{2^{2n}} + \frac{3q^3\ell^5 + 143q^3\ell^6 + 11q^4\ell^3}{2^{3n}} + \frac{17q^4\ell^6 + 5462q^4\ell^8}{2^{4n}}.$$

The proof of this lemma is postponed to section 5.

GOOD TRANSCRIPT: Now, fix a good transcript $\nu = (\tilde{m}, \tilde{t}, \tilde{x}, \tilde{x}^*, \tilde{y}, l_{-1}, l_0)$. Let σ be the total number of blocks (and one additional for 0^n) and $\sigma' := |\tilde{x} \cup \{0^n\}|$. Since, ν is good, we have $(0^n, \tilde{x}, \tilde{x}^*) \rightsquigarrow (L, \tilde{y}, \tilde{t})$. Then, we must have $|\tilde{x}^*| = q$. Further, let $|\tilde{x} \cap \tilde{x}^*| = r$. Thus, $|\{0^n\} \cup \tilde{x} \cup \tilde{x}^*| = q + \sigma' - r$.

Real world: In the real world, the random permutation Π is sampled on exactly $q + \sigma' - r$ distinct points. Thus, we have

$$\Pr(\mathbf{V}_1 = \nu) = \frac{1}{(2^n)_{q+\sigma'-r}}. \quad (12)$$

Ideal world: In the ideal world, we employed a two stage sampling. First of all, we have

$$\Pr(\tilde{\mathbf{T}} = \tilde{t}, \mathbf{P}(0^n) = L) \leq \frac{1}{2^{nq}}, \quad (13)$$

since each \mathbf{T}_i is sampled uniformly from the set \mathcal{B} independent of others. Now, observe that all the full collision and induced collision indices are fully determined from the transcript ν itself. In other words, we can enumerate the set $\tilde{\mathbf{C}}\mathbf{I} := \widetilde{\mathbf{F}}\mathbf{C}\mathbf{I} \cup \widetilde{\mathbf{I}}\mathbf{C}\mathbf{I}$. Now, since the transcript is good, we must have $|\tilde{\mathbf{C}}\mathbf{I}| = \sigma - \sigma' + |\tilde{x} \cap \tilde{x}^*| = \sigma - \sigma' + r$, and for all indices $(i, a) \notin \tilde{\mathbf{C}}\mathbf{I}$, we have $\mathbf{Y}_i[a] = \mathbf{Z}_i[a]$. Thus, we have

$$\begin{aligned} \Pr(\mathbf{Y}_i[a] = y_a^i \wedge (i, a) \notin \tilde{\mathbf{C}}\mathbf{I} \mid \tilde{\mathbf{T}} = \tilde{t}) &= \Pr(\mathbf{Z}_i[a] = y_a^i \wedge (i, a) \notin \tilde{\mathbf{C}}\mathbf{I} \mid \tilde{\mathbf{T}} = \tilde{t}) \\ &= \frac{1}{(2^n - q)_{\sigma'-r}}, \end{aligned} \quad (14)$$

where the second equality follows from the fact that truncation⁶ of a without replacement sample from a set of size $(2^n - q)$ is still a without replacement sample from the same set. We have

$$\begin{aligned} \Pr(V_0 = \omega) &= \Pr(\tilde{T} = \tilde{t}) \times \Pr(\tilde{Y} = \tilde{y} \mid \tilde{T} = \tilde{t}) \\ &\leq \frac{1}{2^{nq}} \times \Pr(Y_i[a] = y_i[a] \wedge (i, a) \notin \tilde{CI} \mid \tilde{T} = \tilde{t}) = \frac{1}{2^{nq}(2^n - q)^{\sigma' - r}}. \end{aligned} \quad (15)$$

The above discussion on good transcripts can be summarized in shape of the following lemma.

Lemma 4.2. *For any $\nu \in \mathcal{V} \setminus \mathcal{V}_{\text{bad}}$, we have $\frac{\Pr(V_1 = \nu)}{\Pr(V_0 = \nu)} \geq 1$.*

Proof. The proof follows from dividing (12) by (15).

Using Theorem 2.1, and Lemma 4.1 and 4.2, we get

$$\begin{aligned} \text{Adv}_{\text{OMAC}_\Pi}^{\text{prf}}(q, \ell, \sigma, \infty) &\leq \frac{4\sigma}{2^n} + \frac{16q^2 + q\ell^2}{2^n} + \frac{8q^2\ell^4 + 32q^3\ell^2 + 2q^2\ell^3}{2^{2n}} \\ &\quad + \frac{3q^3\ell^5 + 143q^3\ell^6 + 11q^4\ell^3}{2^{3n}} + \frac{17q^4\ell^6 + 5462q^4\ell^8}{2^{4n}}. \end{aligned} \quad (16)$$

Theorem 3.1 follows from (11) and (16).

5 Proof of Lemma 4.1

In supplementary material A, we recall the definition and properties of a combinatorial tool, called structure graphs [24,11], which will be highly useful in our proof. Our aim will be to bound the probability of bad events only when they occur in conjunction with some “manageable” structure graphs. In all other cases, we upper bound the probability by the probability of realizing an unmanageable structure graph. Formally, we say that the structure graph $\mathcal{G}_P(\tilde{M})$ is manageable if and only if:

1. for all $i \in [q]$, we have $\text{Acc}(\mathcal{G}_P(M_i)) = 0$, i.e., each M_i -walk is a path.
2. for all distinct $i, j \in [q]$, we have $\text{Acc}(\mathcal{G}_P(M_i, M_j)) \leq 1$.
3. for all distinct $i, j, k \in [q]$, we have $\text{Acc}(\mathcal{G}_P(M_i, M_j, M_k)) \leq 2$.
4. for all distinct $i, j, k, l \in [q]$, we have $\text{Acc}(\mathcal{G}_P(M_i, M_j, M_k, M_l)) \leq 3$.

Let unman denote the event that $\mathcal{G}_P(\tilde{M})$ is unmanageable. Then, using Corollary A.1, we have

$$\Pr(\text{unman}) \leq \Pr(\exists i \in [q] : \text{Acc}(\mathcal{G}_P(M_i)) \geq 1) + \Pr(\exists i < j \in [q] : \text{Acc}(\mathcal{G}_P(M_i, M_j)) \geq 2)$$

⁶ Removing some elements from the tuple.

$$\begin{aligned}
& + \Pr(\exists i < j < k \in [q] : \text{Acc}(\mathcal{G}_P(M_i, M_j, M_k)) \geq 3) \\
& + \Pr(\exists i < j < k < l \in [q] : \text{Acc}(\mathcal{G}_P(M_i, M_j, M_k, M_l)) \geq 4) \\
\leq & \sum_{i \in [q]} \frac{(\ell_i - 1)^2}{2^n} + \sum_{i < j \in [q]} \frac{(\ell_i + \ell_j - 2)^4}{2^{2n}} + \sum_{i < j < k \in [q]} \frac{(\ell_i + \ell_j + \ell_k - 3)^6}{2^{3n}} \\
& + \sum_{i < j < k < l \in [q]} \frac{(\ell_i + \ell_j + \ell_k + \ell_l - 4)^8}{2^{4n}} \\
\leq & \frac{q\ell^2}{2^n} + \frac{8q^2\ell^4}{2^{2n}} + \frac{121.5q^3\ell^6}{2^{3n}} + \frac{5461.34q^4\ell^8}{2^{4n}}. \tag{17}
\end{aligned}$$

From now on we only consider manageable graphs. Observe that apart from the fact that a manageable graph is just a union of M_i -paths, there is an added benefit that it has no zero collision. Let $\text{TU} := \neg(\text{BadT} \vee \text{unman})$ and $\text{TUW} := \neg(\text{BadT} \vee \text{unman} \vee \text{BadW})$. Now, we have

$$\begin{aligned}
\Pr(\mathbf{V}_0 \in \mathcal{V}_{\text{bad}}) &= \Pr((\text{FlagT} = 1) \vee (\text{FlagW} = 1) \vee (\text{FlagX} = 1)) \\
&\stackrel{1}{\leq} \Pr(\text{BadT} \vee \text{BadW} \vee \text{BadX}) \\
&\leq \Pr(\text{BadT}) + \Pr(\text{BadW} | \neg \text{BadT}) + \Pr(\text{BadX} | \neg(\text{BadT} \vee \text{BadW})) \\
&\stackrel{2}{\leq} \Pr(\exists i \neq j : T_i = T_j) + \Pr(\text{BadW} | \neg \text{BadT}) + \Pr(\text{BadX} | \neg(\text{BadT} \vee \text{BadW})) \\
&\stackrel{3}{\leq} \frac{q^2}{2^{n+1}} + \Pr(\text{unman}) + \Pr(\text{BadW} | \text{TU}) + \Pr(\text{BadX} | \text{TUW}) \\
&\stackrel{4}{\leq} \frac{0.5q^2 + q\ell^2}{2^n} + \frac{8q^2\ell^4}{2^{2n}} + \frac{122q^3\ell^6}{2^{3n}} + \frac{5462q^4\ell^8}{2^{4n}} \\
&\quad + \Pr(\text{BadW} | \text{TU}) + \Pr(\text{BadX} | \text{TUW}) \tag{18}
\end{aligned}$$

Here, inequalities 1 and 2 follow by definition; 3 follows from the fact that T_i is chosen uniformly at random from \mathcal{B} for each i ; and 4 follows from (17).

BOUNDING $\Pr(\text{BadW} | \neg(\text{BadT} \vee \text{unman}))$: Let $\text{E}_i = \neg(\text{TU} \vee \text{BadW}_1 \vee \dots \vee \text{BadW}_i)$. We have

$$\begin{aligned}
\Pr(\text{BadW} | \text{TU}) &\leq \Pr(\text{BadW}_1 | \text{TU}) + \Pr(\text{BadW}_2 | \text{E}_1) + \Pr(\text{BadW}_3 | \text{E}_2) \\
&\quad + \Pr(\text{BadW}_4 | \text{E}_3) + \Pr(\text{BadW}_5 | \text{E}_4) \tag{19}
\end{aligned}$$

We bound the individual terms on the right hand side as follows:

Bounding $\Pr(\text{BadW}_1 | \text{TU})$: Fix some $(i, a) \in [q] \times [\ell_i]$. The only way we can have $\overline{W}_i[a] = 0^n$, for $1 < a < \ell_i$, is if $Z_i[a-1] = M_i[a]$. This happens with probability at most $(2^n - q)^{-1}$. For $a = \ell_i$, the equation

$$\mu_{\delta_{M_i}} \odot \mathbf{L} \oplus Z_i[\ell_i - 1] \oplus \overline{M}_i[\ell_i] = 0^n$$

must hold non-trivially. The probability that this equation holds is bounded by at most $(2^n - q - 1)^{-1}$. Assuming $q + 1 \leq 2^{n-1}$, and using the fact that there can be at most σ choices for (i, a) , we have

$$\Pr(\text{BadW1}|\text{TU}) \leq \frac{2\sigma}{2^n}. \quad (20)$$

Bounding $\Pr(\text{BadW2}|\text{E1})$: Fix some $i \neq j \in [q]$. Since $\neg\text{unman}$ holds, we know that $\text{Acc}(\mathcal{G}_P(\mathbf{M}_i, \mathbf{M}_j)) \leq 1$. We handle the two resulting cases separately:

- (A) $\text{Acc}(\mathcal{G}_P(\mathbf{M}_i, \mathbf{M}_j)) = 1$: Suppose the collision source of the only accident are (i, a) and (j, b) . Then, we have the following system of two equations

$$\begin{aligned} Z_i[a] \oplus Z_j[b] &= M_i[a+1] \oplus M_j[b+1] \\ (\mu_{\delta_{M_i}} \oplus \mu_{\delta_{M_j}}) \odot L \oplus Z_i[\ell_i - 1] \oplus Z_j[\ell_j - 1] &= \bar{M}_i[\ell_i] \oplus \bar{M}_j[\ell_j] \end{aligned}$$

Suppose $\delta_{M_i} \neq \delta_{M_j}$, i.e. $\mu_{\delta_{M_i}} \oplus \mu_{\delta_{M_j}} \neq 0^n$. Using the fact that $\neg\text{BadW1}$ holds, we infer that $L \notin \{Z_i[a], Z_j[b], Z_i[\ell_i - 1], Z_j[\ell_j - 1]\}$. So, the two equations are linearly independent, whence the rank is 2 in this case. Again, using Lemma A.4, and the fact that there are at most $q^2/2$ choices for i and j , and ℓ^2 choices for a and b , we get

$$\Pr(\text{BadW2} \wedge \text{Case A} \wedge \delta_{M_i} \neq \delta_{M_j} | \text{E1}) \leq \frac{q^2 \ell^2}{2(2^n - q - \sigma + 2)^2}.$$

Now, suppose $\delta_{M_i} = \delta_{M_j}$, i.e. $\mu_{\delta_{M_i}} \oplus \mu_{\delta_{M_j}} = 0^n$. Then, we can rewrite the system as

$$\begin{aligned} Z_i[a] \oplus Z_j[b] &= M_i[a+1] \oplus M_j[b+1] \\ Z_i[\ell_i - 1] \oplus Z_j[\ell_j - 1] &= \bar{M}_i[\ell_i] \oplus \bar{M}_j[\ell_j] \end{aligned}$$

We can have two types of structure graphs relevant to this case, as illustrated in Figure 5.1. For type 1 all variables are distinct. So, the two equations are



Fig. 5.1: Accident-1 manageable graphs for two messages. The solid and dashed lines correspond to edges in \mathcal{W}_i and \mathcal{W}_j , respectively. * denotes optional parts in the walk.

linearly independent, whence the rank is 2 in this case. Again, using Lemma A.4, we get

$$\Pr(\text{BadW2} \wedge \text{Case A} \wedge \delta_{M_i} = \delta_{M_j} \wedge \text{Type 1} | \text{E1}) \leq \frac{q^2 \ell^2}{2(2^n - q - \sigma + 2)^2}.$$

For type 2, it is clear that $Z_j[\ell_j - 1] = Z_i[\ell_i - 1]$. So, we can assume that the second equation holds trivially, thereby deriving a system in $Z_i[a]$ and $Z_j[b]$, with rank 1. Further, a and b are uniquely determined as $\ell_i - p$ and $\ell_j - p$, where p is the longest common suffix of M_i and M_j . So we have

$$\Pr(\text{BadW2} \wedge \text{Case A} \wedge \delta_{M_i} = \delta_{M_j} \wedge \text{Type 2|E1}) \leq \frac{q^2}{2(2^n - q - \sigma + 1)}.$$

(B) $\text{Acc}(\mathcal{G}_P(M_i, M_j)) = 0$: In this case, we only have one equation of the form

$$(\mu_{\delta_{M_i}} \oplus \mu_{\delta_{M_j}}) \odot \mathbf{L} \oplus Z_i[\ell_i - 1] \oplus Z_j[\ell_j - 1] = \overline{M}_i[\ell_i] \oplus \overline{M}_j[\ell_j]$$

If $\delta_{M_i} \neq \delta_{M_j}$, we have an equation in three variables, namely \mathbf{L} , $Z_i[\ell_i - 1]$, and $Z_j[\ell_j - 1]$; and if $\delta_{M_i} = \delta_{M_j}$, we have an equation in two variables, namely $Z_i[\ell_i - 1]$, and $Z_j[\ell_j - 1]$. In both the cases, the equation can only hold non-trivially, i.e., rank is 1. Using Lemma A.4, we get

$$\Pr(\text{BadW2} \wedge \text{Case B|E1}) \leq \frac{q^2}{2(2^n - q - \sigma + 1)}.$$

On combining the three cases, we get

$$\Pr(\text{BadW2|E1}) \leq \frac{q^2}{2^n - q - \sigma + 1} + \frac{q^2 \ell^2}{(2^n - q - \sigma + 2)^2}. \quad (21)$$

Bounding $\Pr(\text{BadW3|E2})$: Fix some $i, j, k \in [q]$. Since $\neg \text{unman}$ holds, we must have $\text{Acc}(\mathcal{G}_P(M_i, M_j, M_k)) \leq 2$. Accordingly, we have the following three cases:

(A) $\text{Acc}(\mathcal{G}_P(M_i, M_j, M_k)) = 2$: Suppose (α_1, β_1) and (α_2, β_2) are collision source leading to one of the accident, and (α_3, β_3) and (α_4, β_4) are collision source leading to the other accident. Then, considering $W_i[a] = W_j[\ell_j]$, we have the following system of equations

$$\begin{aligned} Z_{\alpha_1}[\beta_1] \oplus Z_{\alpha_2}[\beta_2] &= M_{\alpha_1}[\beta_1 + 1] \oplus M_{\alpha_2}[\beta_2 + 1] \\ Z_{\alpha_3}[\beta_3] \oplus Z_{\alpha_4}[\beta_4] &= M_{\alpha_3}[\beta_3 + 1] \oplus M_{\alpha_4}[\beta_4 + 1] \\ Z_j[a - 1] \oplus \mu_{\delta_{M_j}} \odot \mathbf{L} \oplus Z_j[\ell_j - 1] &= \overline{M}_j[\ell_j] \oplus M_i[a] \end{aligned}$$

The first two equations are independent by definition. Further, using $\neg \text{BadW1}$, we can infer that the last equation is also independent of the first two equations. Thus the system has rank 3. There are at most $q^3/6$ choices for (i, j, k) , and for each such choice we have 3 choices for $(\alpha_1, \alpha_2, \alpha_3, \alpha_4)$ and at most ℓ^5 choices for $(\beta_1, \beta_2, \beta_3, \beta_4, a)$. Using Lemma A.4, we have

$$\Pr(\text{BadW3} \wedge \text{Case A|E2}) \leq \frac{q^3 \ell^5}{2(2^n - q - \sigma + 3)^3}.$$

- (B) $\text{Acc}(\mathcal{G}_P(\mathbf{M}_i, \mathbf{M}_j, \mathbf{M}_k)) = 1$: Suppose (α_1, β_1) and (α_2, β_2) are collision source leading to the accident. First consider the case $a < \ell_i - 1$ and $b < \ell_k$. In this case, we have the following system of equations

$$\begin{aligned} Z_{\alpha_1}[\beta_1] \oplus Z_{\alpha_2}[\beta_2] &= \mathbf{M}_{\alpha_1}[\beta_1 + 1] \oplus \mathbf{M}_{\alpha_2}[\beta_2 + 1] \\ Z_i[a - 1] \oplus \mu_{\delta_{\mathbf{M}_j}} \odot \mathbf{L} \oplus Z_j[\ell_j - 1] &= \overline{\mathbf{M}}_j[\ell_j] \oplus \mathbf{M}_i[a] \\ Z_i[a] \oplus Z_k[b - 1] &= \mathbf{M}_i[a + 1] \oplus \mathbf{M}_k[b] \end{aligned}$$

The first two equations are clearly independent. Further, since $\mathbf{M}_i \neq \mathbf{M}_k$, the last equation must correspond to a true collision as a consequence of the accident. So, the rank of the above system is 2. Once we fix (i, j, k) and (a, b) , we have at most 3 choices for (α_1, α_2) , and β_1 and β_2 are uniquely determined as $a + 1 - p$ and $b - p$, where p is the largest common suffix of $\mathbf{M}_i[1, \dots, a + 1]$ and $\mathbf{M}_k[1, \dots, b]$. So, we have

$$\Pr(\text{BadW3} \wedge \text{Case B} \wedge a < \ell_i - 1 \wedge b < \ell_k | \mathbf{E2}) \leq \frac{q^3 \ell^2}{2(2^n - q - \sigma + 2)^2}.$$

Now, suppose $a = \ell_i - 1$. Then we can simply consider the first two equations

$$\begin{aligned} Z_{\alpha_1}[\beta_1] \oplus Z_{\alpha_2}[\beta_2] &= \mathbf{M}_{\alpha_1}[\beta_1 + 1] \oplus \mathbf{M}_{\alpha_2}[\beta_2 + 1] \\ Z_j[\ell_i - 2] \oplus \mu_{\delta_{\mathbf{M}_j}} \odot \mathbf{L} \oplus Z_j[\ell_j - 1] &= \overline{\mathbf{M}}_j[\ell_j] \oplus \mathbf{M}_i[\ell_i - 1] \end{aligned}$$

Clearly, the two equations are independent. We have at most q^3 choices for (i, j, k) , 3 choices for (α_1, α_2) , and ℓ^2 choices for (β_1, β_2) . So we have

$$\Pr(\text{BadW3} \wedge \text{Case B} \wedge a = \ell_i - 1 | \mathbf{E2}) \leq \frac{q^3 \ell^2}{2(2^n - q - \sigma + 2)^2}.$$

The case where $a < \ell_i - 1$ and $b = \ell_k$ can be handled similarly by considering the first and the third equations.

- (C) $\text{Acc}(\mathcal{G}_P(\mathbf{M}_i, \mathbf{M}_j, \mathbf{M}_k)) = 0$: In this case, we know that the three paths, \mathcal{W}_i , \mathcal{W}_j , and \mathcal{W}_k do not collide. This implies that we must have $a = \ell_i - 1$, or $b = \ell_k$ or both, in order for $\mathbf{W}_i[a + 1] = \mathbf{W}_k[b]$ to hold. First, suppose both $a = \ell_i - 1$ and $b = \ell_k$. Then, we have the following system of equations:

$$\begin{aligned} Z_j[\ell_i - 2] \oplus \mu_{\delta_{\mathbf{M}_j}} \odot \mathbf{L} \oplus Z_j[\ell_j - 1] &= \overline{\mathbf{M}}_j[\ell_j] \oplus \mathbf{M}_i[\ell_i - 2] \\ (\mu_{\delta_{\mathbf{M}_i}} \oplus \mu_{\delta_{\mathbf{M}_k}}) \odot \mathbf{L} \oplus Z_i[\ell_i - 1] \oplus Z_k[\ell_k - 1] &= \overline{\mathbf{M}}_i[\ell_i] \oplus \overline{\mathbf{M}}_k[\ell_k] \end{aligned}$$

Using the properties of μ_{-1} and μ_0 , and $\neg\text{BadW1}$, we can conclude that the above system has rank 2. There are at most $q^3/6$ choices for (i, j, k) , and at most ℓ^2 choices for (a, b) . So, we have

$$\Pr(\text{BadW3} \wedge \text{Case C} \wedge a = \ell_i - 1 \wedge b = \ell_k | \mathbf{E2}) \leq \frac{q^3 \ell^2}{6(2^n - q - \sigma + 2)^2}.$$

The remaining two cases are similar. We handle the case $a = \ell_i - 1$ and $b < \ell_k$, and the other case can be handled similarly. We have the following system of equations

$$\begin{aligned} Z_j[\ell_i - 2] \oplus \mu_{\delta_{M_j}} \odot L \oplus Z_j[\ell_j - 1] &= \overline{M}_j[\ell_j] \oplus M_i[\ell_i - 2] \\ \mu_{\delta_{M_i}} \odot L \oplus Z_i[\ell_i - 1] \oplus Z_k[b - 1] &= \overline{M}_i[\ell_i] \oplus M_k[b] \end{aligned}$$

If $\delta_{M_i} \neq \delta_{M_j}$, then using the same argument as above, we can conclude that the system has rank 2, and we get

$$\Pr(\text{BadW3} \wedge \text{Case C} \wedge a = \ell_i - 1 \wedge b < \ell_k \wedge \delta_{M_i} \neq \delta_{M_j} | \mathbf{E2}) \leq \frac{q^3 \ell^2}{6(2^n - q - \sigma + 2)^2}.$$

So, suppose $\delta_{M_i} = \delta_{M_j}$. Now, in order for the second equation to be a consequence of the first equation, we must have $Z_i[\ell_i - 2] = Z_j[\ell_j - 1]$ and $Z_i[\ell_i - 1] = Z_k[b]$. The only way this happens trivially is if $M_i[1, \dots, \ell_i - 1] = M_j[1, \dots, \ell_j - 1]$ and $M_i[1, \dots, \ell_i - 1] = M_k[1, \dots, b]$. But, then we have $b = \ell_i - 1$, and once we fix (i, k) there's a unique choice for j , since $M_j[1, \dots, \ell_j - 1] = M_i[1, \dots, \ell_i - 1]$ and $\overline{M}_j[\ell_j] = \overline{M}_i[\ell_i] \oplus M_i[\ell_i - 2] \oplus M_k[b]$. So, we get

$$\Pr(\text{BadW3} \wedge \text{Case C} \wedge a = \ell_i - 1 \wedge b < \ell_k \wedge \delta_{M_i} = \delta_{M_j} | \mathbf{E2}) \leq \frac{q^2}{2(2^n - q - \sigma + 1)}.$$

By combining all three cases, we have

$$\Pr(\text{BadW3} | \mathbf{E2}) \leq \frac{q^3 \ell^5}{2(2^n - q - \sigma + 3)^3} + \frac{2q^3 \ell^2}{(2^n - q - \sigma + 2)^2} + \frac{q^2}{2(2^n - q - \sigma + 1)}. \quad (22)$$

Bounding $\Pr(\text{BadW4} | \mathbf{E3})$: Fix some $i, j, k \in [q]$. The analysis in this case is very similar to the one in case of $\text{BadW3} | \mathbf{E2}$. So we will skip detailed argumentation whenever possible. Since $\neg \text{unman}$ holds, we must have $\text{Acc}(\mathcal{G}_P(M_i, M_j, M_k)) \leq 2$. Accordingly, we have the following three cases:

- (A) $\text{Acc}(\mathcal{G}_P(M_i, M_j, M_k)) = 2$: This can be bounded by using exactly the same argument as used in Case A for $\text{BadW3} | \mathbf{E2}$. So, we have

$$\Pr(\text{BadW4} \wedge \text{Case A} | \mathbf{E3}) \leq \frac{q^3 \ell^5}{2(2^n - q - \sigma + 3)^3}.$$

- (B) $\text{Acc}(\mathcal{G}_P(M_i, M_j, M_k)) = 1$: Suppose (α_1, β_1) and (α_2, β_2) are collision source leading to the accident. Without loss of generality we assume $a < b$. Specifically, $b \leq \ell_i - 1$ and $a \leq b - 2$ due to $\neg(\text{BadW2} \wedge \text{BadW3})$. First consider the case $b = \ell_i - 1$. In this case, considering $W_i[b] = W_k[\ell_k]$, we have the following system of equations

$$Z_{\alpha_1}[\beta_1] \oplus Z_{\alpha_2}[\beta_2] = M_{\alpha_1}[\beta_1 + 1] \oplus M_{\alpha_2}[\beta_2 + 1]$$

$$Z_i[b-1] \oplus \mu_{\delta_{M_k}} \odot L \oplus Z_k[\ell_k-1] = \overline{M}_k[\ell_k] \oplus M_i[b]$$

Using a similar argument as used in previous such cases, we establish that the two equations are independent. Now, once we fix (i, j, k) , we have exactly one choice for b , at most 3 choices for (α_1, α_2) , and ℓ^2 choices for (β_1, β_2) . So, we have

$$\Pr(\text{BadW4} \wedge \text{Case B} \wedge b = \ell_i - 1 | \mathbf{E3}) \leq \frac{q^3 \ell^2}{2(2^n - q - \sigma + 2)^2}.$$

Now, suppose $b < \ell_i - 1$. Here we can have two cases:

- (B.1) \mathcal{W}_i is involved in the accident: Without loss of generality assume that $\alpha_1 = i$ and $\beta_1 \in [\ell_i - 1]$. Then, we have the following system of equations:

$$\begin{aligned} Z_i[\beta_1] \oplus Z_{\alpha_2}[\beta_2] &= M_i[\beta_1 + 1] \oplus M_{\alpha_2}[\beta_2 + 1] \\ Z_i[a-1] \oplus \mu_{\delta_{M_j}} \odot L \oplus Z_j[\ell_j-1] &= \overline{M}_j[\ell_j] \oplus M_i[a] \\ Z_i[b-1] \oplus \mu_{\delta_{M_k}} \odot L \oplus Z_k[\ell_k-1] &= \overline{M}_k[\ell_k] \oplus M_i[b] \end{aligned}$$

Suppose $Z_i[\beta_1] = Z_i[a-1]$. Then, we must have $\beta_1 = a-1$ as the graph is manageable. In this case, we consider the first two equations. It is easy to see that the two equations are independent, and once we fix i, j, k , there are at most 2 choices for α_2 and ℓ^2 choices for (β_1, β_2) , which gives a unique choice for a . So, we have

$$\Pr(\text{BadW4} \wedge \text{Case B.1} \wedge \beta_1 = a-1 | \mathbf{E3}) \leq \frac{q^3 \ell^2}{2(2^n - q - \sigma + 2)^2}.$$

We get identical bound for the case when $Z_i[\beta_1] = Z_i[b-1]$. Suppose $Z_i[\beta_1] \notin \{Z_i[a-1], Z_i[b-1]\}$. Then, using the fact that there is only one accident in the graph and that accident is due to (i, β_1) and (α_2, β_2) , we infer that $Z_{\alpha_2}[\beta_2] \notin \{Z_i[a-1], Z_i[b-1]\}$. Now, the only way rank of the above system reduces to 2, is if $Z_i[a-1] = Z_k[\ell_k-1]$ and $Z_i[b-1] = Z_j[\ell_j-1]$ trivially. However, if this happens then a and b are uniquely determined by our choice of $(i, j, k, \beta_1, \alpha_2, \beta_2)$. See Figure 5.2 for the two possible structure graphs depending upon the value of α_2 . Basically, based on the choice of α_2 , $a \in \{\ell_k, \ell_k - \beta_2 + \beta_1\}$. Similarly, $b \in \{\ell_j, \ell_j - \beta_2 + \beta_1\}$. So, using Lemma A.4, we get

$$\Pr(\text{BadW4} \wedge \text{Case B.1} \wedge \beta_1 \notin \{a-1, b-1\} | \mathbf{E3}) \leq \frac{2q^3 \ell^2}{3(2^n - q - \sigma + 2)^2}.$$

- (B.2) \mathcal{W}_i is not involved in the accident: Without loss of generality assume $\alpha_1 = j$ and $\alpha_2 = k$. Then, we have the following system of equations:

$$\begin{aligned} Z_j[\beta_1] \oplus Z_k[\beta_2] &= M_j[\beta_1 + 1] \oplus M_k[\beta_2 + 1] \\ Z_i[a-1] \oplus \mu_{\delta_{M_j}} \odot L \oplus Z_j[\ell_j-1] &= \overline{M}_j[\ell_j] \oplus M_i[a] \end{aligned}$$

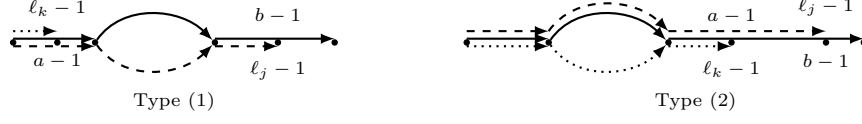


Fig. 5.2: Manageable graphs for case B.1. The solid, dashed and dotted lines correspond to edges in \mathcal{W}_i , \mathcal{W}_j , and \mathcal{W}_k , respectively.

$$Z_i[b-1] \oplus \mu_{\delta_{M_k}} \odot L \oplus Z_k[l_k-1] = \overline{M}_k[l_k] \oplus M_i[b]$$

Since the graph is manageable, $\{Z_i[a-1], Z_i[b-1]\} \cap \{Z_j[l_j-1], Z_k[l_k-1]\} \neq \emptyset$. Suppose $\{Z_i[a-1], Z_i[b-1]\} = \{Z_j[l_j-1], Z_k[l_k-1]\}$. Without loss of generality, assume $Z_i[a-1] = Z_k[l_k-1]$ and $Z_i[b-1] = Z_j[l_j-1]$. This can only happen if the resulting graph is of Type 2 form in Figure 5.2, which clearly shows that we have unique choices for a and b when we fix the other indices. Now, suppose $|\{Z_i[a-1], Z_i[b-1]\} \cap \{Z_j[l_j-1], Z_k[l_k-1]\}| = 1$. Then, we must have $Z_i[a-1] \in \{Z_j[\beta_1], Z_k[\beta_2]\}$ since $a < b$. Without loss of generality we assume that $Z_i[a-1] = Z_k[\beta_2]$ and $Z_i[b-1] = Z_j[l_j-1]$. Using similar argument as before, we conclude that a and b are fixed once we fix all other indices. So using Lemma A.4, we get

$$\Pr(\text{BadW4} \wedge \text{Case B.2} | \text{E3}) \leq \frac{2q^3 \ell^2}{3(2^n - q - \sigma + 2)^2}.$$

(C) $\text{Acc}(\mathcal{G}_P(M_i, M_j, M_k)) = 0$: In this case, we know that the three paths, \mathcal{W}_i , \mathcal{W}_j , and \mathcal{W}_k do not collide. We have the following system of equations:

$$Z_i[a-1] \oplus \mu_{\delta_{M_j}} \odot L \oplus Z_j[l_j-1] = \overline{M}_j[l_j] \oplus M_i[a]$$

$$Z_i[b-1] \oplus \mu_{\delta_{M_k}} \odot L \oplus Z_k[l_k-1] = \overline{M}_k[l_k] \oplus M_i[b]$$

Using a similar analysis as in case C of BadW3|E2, we get

$$\Pr(\text{BadW4} \wedge \text{Case C} | \text{E3}) \leq \frac{q^3 \ell^2}{6(2^n - q - \sigma + 2)^2} + \frac{q^2}{2(2^n - q - \sigma + 1)}.$$

By combining all three cases, we have

$$\Pr(\text{BadW4} | \text{E3}) \leq \frac{q^3 \ell^5}{2(2^n - q - \sigma + 3)^3} + \frac{3q^3 \ell^2}{(2^n - q - \sigma + 2)^2} + \frac{q^2}{2(2^n - q - \sigma + 1)}. \quad (23)$$

Bounding $\Pr(\text{BadW5} | \text{E4})$: Fix some $i, j, k \in [q]$. The analysis in this case is again similar to the analysis of BadW3|E2 and BadW4|E3. We have the following three cases:

(A) $\text{Acc}(\mathcal{G}_P(M_i, M_j, M_k)) = 2$: This can be bounded by using exactly the same argument as used in Case A for BadW3|E2. So, we have

$$\Pr(\text{BadW5} \wedge \text{Case A} | \text{E4}) \leq \frac{q^3 \ell^5}{2(2^n - q - \sigma + 3)^3}.$$

- (B) $\text{Acc}(\mathcal{G}_P(\mathbf{M}_i, \mathbf{M}_j, \mathbf{M}_k)) = 1$: Suppose (α_1, β_1) and (α_2, β_2) are collision source leading to the accident. In this case, we have the following system of equations

$$\begin{aligned} Z_{\alpha_1}[\beta_1] \oplus Z_{\alpha_2}[\beta_2] &= M_{\alpha_1}[\beta_1 + 1] \oplus M_{\alpha_2}[\beta_2 + 1] \\ Z_i[a - 1] \oplus \mu_{\delta_{M_j}} \odot L \oplus Z_j[\ell_j - 1] &= \bar{M}_j[\ell_j] \oplus M_i[a] \\ Z_j[b - 1] \oplus \mu_{\delta_{M_k}} \odot L \oplus Z_k[\ell_k - 1] &= \bar{M}_k[\ell_k] \oplus M_j[b] \end{aligned}$$

We can have two sub-cases:

- (B.1) Suppose the third equation is simply a consequence of the second equation. Then, we must have $\delta_{M_i} = \delta_{M_j}$ and $Z_i[a - 1] = Z_j[b - 1]$ and $Z_j[\ell_j - 1] = Z_k[\ell_k - 1]$ must hold trivially, since the graph is manageable. We claim that $a = b = \text{Prefix}(M_i[1], M_j[1]) + 1$. If not, then $M_i[\ell_i] = M_j[\ell_j]$ which in conjunction with $Z_j[\ell_j - 1] = Z_k[\ell_k - 1]$ implies that $W_i[\ell_i] = W_j[\ell_j]$ which contradicts **BadW2**. So, using Lemma [A.4](#), we get

$$\Pr(\text{BadW5} \wedge \text{Case B.1} | \mathbf{E4}) \leq \frac{q^3 \ell^2}{2(2^n - q - \sigma + 2)^2}.$$

- (B.2) The second and third equation are independent. Considering the subsystem consisting of these two equations, and using Lemma [A.4](#), we get

$$\Pr(\text{BadW5} \wedge \text{Case B.2} | \mathbf{E4}) \leq \frac{q^3 \ell^2}{6(2^n - q - \sigma + 2)^2}.$$

- (C) $\text{Acc}(\mathcal{G}_P(\mathbf{M}_i, \mathbf{M}_j, \mathbf{M}_k)) = 0$: We have the following system of equations:

$$\begin{aligned} Z_i[a - 1] \oplus \mu_{\delta_{M_j}} \odot L \oplus Z_j[\ell_j - 1] &= \bar{M}_j[\ell_j] \oplus M_i[a] \\ Z_i[b - 1] \oplus \mu_{\delta_{M_k}} \odot L \oplus Z_k[\ell_k - 1] &= \bar{M}_i[\ell_k] \oplus M_i[b] \end{aligned}$$

Let r denote the rank of the above system. Using a similar analysis as in case B.1 above, we conclude that $a = b = \text{Prefix}(M_i[1], M_j[1]) + 1$ if $r = 1$. Using Lemma [A.4](#), we get

$$\begin{aligned} \Pr(\text{BadW5} \wedge \text{Case C} \wedge r = 1 | \mathbf{E4}) &\leq \frac{q^2}{2(2^n - q - \sigma + 1)}. \\ \Pr(\text{BadW5} \wedge \text{Case C} \wedge r = 2 | \mathbf{E4}) &\leq \frac{q^3 \ell^2}{6(2^n - q - \sigma + 2)^2}. \end{aligned}$$

By combining all three cases, we have

$$\Pr(\text{BadW5} | \mathbf{E4}) \leq \frac{q^3 \ell^5}{2(2^n - q - \sigma + 3)^3} + \frac{5q^3 \ell^2}{6(2^n - q - \sigma + 2)^2} + \frac{q^2}{2(2^n - q - \sigma + 1)}. \quad (24)$$

Further, from Eq. (19)-(24), we have

$$\Pr(\text{BadW} | \text{TU}) \leq \frac{2\sigma}{2^n} + \frac{5q^2}{2(2^n - q - \sigma + 1)} + \frac{7q^3 \ell^2}{(2^n - q - \sigma + 2)^2} + \frac{3q^3 \ell^5}{2(2^n - q - \sigma + 3)^3}. \quad (25)$$

BOUNDING $\Pr(\text{BadX}|\text{TUW})$: For space consideration, in supplementary material [B](#), we show that

$$\begin{aligned} \Pr(\text{BadX}|\text{TUW}) \leq & \frac{2\sigma}{2^n} + \frac{10q^2}{2^n - q - \sigma + 1} + \frac{15q^3\ell^2 + q^2\ell^3}{(2^n - q - \sigma + 2)^2} \\ & + \frac{12q^3\ell^6 + 6q^4\ell^3}{(2^n - q - \sigma + 3)^3} + \frac{8q^4\ell^6}{(2^n - q - \sigma + 4)^4} \end{aligned} \quad (26)$$

Combining Eq. (18), (25), and (26), we have

$$\begin{aligned} \Pr(V_0 \in \mathcal{V}_{\text{bad}}) \leq & \frac{4\sigma}{2^n} + \frac{16q^2 + q\ell^2}{2^n} + \frac{8q^2\ell^4 + 32q^3\ell^2 + 2q^2\ell^3}{2^{2n}} \\ & + \frac{3q^3\ell^5 + 143q^3\ell^6 + 11q^4\ell^3}{2^{3n}} + \frac{17q^4\ell^6 + 5462q^4\ell^8}{2^{4n}}. \end{aligned} \quad (27)$$

6 Conclusion

In this paper we proved that OMAC, XCBC and TMAC are secure up to $q \leq 2^{n/2}$ queries, while the message length $\ell \leq 2^{n/4}$. As a consequence, we have proved that OMAC – a single-keyed CBC-MAC variant – achieves the same security level as some of the more elaborate CBC-MAC variants like EMAC and ECBC. This, in combination with the existing results [11,21], shows that the security is tight up to $\ell \leq 2^{n/4}$ for all CBC-MAC variants except for the original CBC-MAC. It could be an interesting future problem to extend our analysis and derive similar bounds for CBC-MAC over prefix-free message space. In order to prove our claims, we employed reset-sampling method by Chattopadhyay et al. [17], which seems to be a promising tool in reducing the length-dependency in single-keyed iterated constructions. Indeed, we believe that this tool might even be useful in obtaining better security bounds for single-keyed variants of many beyond-the-birthday-bound constructions.

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Supplementary Materials

A Structure Graphs

In this section, we recall the definition and properties of a combinatorial tool, called structure graphs [24, 11]. Fix a tuple of q distinct messages $\tilde{m} = (m_1, \dots, m_q)$, where $m_i \in \mathcal{B}^{\ell_i}$. Let $\sigma_i = \sum_{j \in [i]} \ell_j$, and $\sigma_q = \sigma$. Note that, we have assumed that the last block of each message is full n -bit long (if not it can be suitably padded). Let $\mathcal{Q} := \{(i, a) \in [q] \times [\ell_i - 1]\}$, and \leq be a natural ordering over \mathcal{Q} , defined as follows:

$$(i, a) \leq (i', a') \text{ if and only if } (i < i') \text{ or } (i = i' \text{ and } a \leq a').$$

In context of the poset $(\mathcal{Q}, \leq) = (\alpha_1 \leq \dots \leq \alpha_\sigma)$, we can naturally define $\alpha_i + j$ as α_{i+j} for any $i \in [\sigma]$ and $j \in [\sigma - i]$. One can define subtraction analogously. Sometimes we also use the subset $\mathcal{Q}^+ := \mathcal{Q} \setminus \{(i, 0) : i \in [q]\}$. Going forward, we sometimes write $v_i[a]$ succinctly as v_α for any $\alpha = (i, a) \in \mathcal{Q}$ and any appropriately defined notation v .

For the message tuple \tilde{m} and a permutation $\pi \in \mathcal{P}(n)$, let \tilde{z} denote the intermediate output tuple generated in OMAC function evaluation over each of the q messages in \tilde{m} , i.e., $z_i[0] = 0^n$, $z_i[a] = \pi(z_i[a-1] \oplus m_i[a])$, for all $(i, a) \in \mathcal{Q}$. Let $\text{in}(i, a) := \min\{(j, b) \leq (i, a) : z_i[a] = z_j[b]\}$.

STRUCTURE GRAPHS: Given the message tuple \tilde{m} and permutation π , the *structure graph* $\mathcal{G}_\pi(\tilde{m}) := (\mathcal{V}, \mathcal{E})$, is an edge-labeled directed graph, where the set of vertices $\mathcal{V} = \{\text{in}(\alpha) : \alpha \in \mathcal{Q}\}$, the set of edges $\mathcal{E} = \{e_\alpha := (\text{in}(\alpha - 1), \text{in}(\alpha)) : \alpha \in \mathcal{Q}^+\}$, and edge e_α is labeled m_α for all $\alpha \in \mathcal{Q}^+$. Note that, it is possible that $e_\alpha = e_\beta$ for some $\alpha, \beta \in \mathcal{Q}^+$, i.e., they represent the same edge with obviously the same label. When we consider a single message m_r , the resulting subgraph is simply a walk, that we call an m_r -walk and denote as \mathcal{W}_r , starting at node $(1, 0)$ and following the labels from $(m_r[1], \dots, m_r[\ell_r - 1])$. So, a structure graph can also be viewed as an union of m_i -walks for all $i \in [q]$.

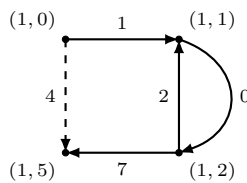


Fig. A.1: Structure graph corresponding to the messages $m_1 = (1, 0, 2, 0, 7, 1)$ and $m_2 = (4, 1)$, and permutation π , with $\pi(1) = 2$, $\pi(2) = 3$ and $\pi(4) = 5$. The solid lines correspond to edges in \mathcal{W}_1 , and dashed lines correspond to edges in \mathcal{W}_2 .

Example A.1. Let $m_1 = (1, 0, 2, 0, 7, 1)$ and $m_2 = (4, 1)$ be two messages and $\pi(1) = 2$; $\pi(2) = 3$; $\pi(4) = 5$ for some $\pi \in \mathcal{P}$. Then, we have $z_1 = (0, 2, 3, 2, 3, 5)$ and $z_2 = (0, 5)$. The corresponding structure graph $\mathcal{G}_\pi(m_1, m_2)$, illustrated in Figure A.1, has vertex set $\mathcal{V} = \{(1, 0), (1, 1), (1, 2), (1, 5)\}$ and edges set

$$\mathcal{E} = \{((1, 0), (1, 1)), ((1, 1), (1, 2)), ((1, 2), (1, 1)), ((1, 2), (1, 5)), ((1, 0), (1, 5))\}.$$

COLLISIONS AND ACCIDENTS: Suppose that $\mathcal{G}_\pi(\tilde{m})$ is revealed edge by edge in an orderly fashion following (\mathcal{Q}^+, \leq) . We say that an edge e_α leads to a *collision* if $\text{in}(\alpha)$ is already present in the partially revealed graph. A collision formed by edges e_α and e_β is generally denoted as $(\text{in}(\alpha - 1), \text{in}(\beta - 1); \gamma)$, where $\gamma = \text{in}(\alpha) = \text{in}(\beta)$. The only exception occurs when $\gamma = (1, 0)$ and there is no prior edge to $(1, 0)$, in which case the collision is denoted as $(\text{in}(\alpha - 1); \gamma)$, since prior to e_α there's no edge pointing to $(1, 0)$. This exceptional case is referred as a *zero collision*, and all other collisions are referred as *true collisions*. We refer to $\text{in}(\alpha - 1)$ (and $\text{in}(\beta - 1)$, if applicable) as *collision source*.

Note that it is not possible to recover the intermediate output tuple, by just looking at a given structure graph. Indeed, multiple intermediate output tuples may give the same structure graphs. However, a structure graph does preserve the collision relation between intermediate outputs. More precisely, let $Z_{\text{in}(\alpha)}$ denote the variable for the intermediate output corresponding to the vertex $\text{in}(\alpha)$, and $\tilde{Z} = (Z_{\text{in}(\alpha)} : \alpha \in \mathcal{Q})$. Obviously, we must have $Z_1[0] = 0^n$, otherwise the resulting intermediate output tuple is invalid. Now, any true collision $(\text{in}(\alpha - 1), \text{in}(\beta - 1); \gamma)$ implies a linear equation

$$Z_{\text{in}(\alpha-1)} \oplus Z_{\text{in}(\beta-1)} = m_\alpha \oplus m_\beta,$$

since both $Z_{\text{in}(\alpha-1)} \oplus m_\alpha$ and $Z_{\text{in}(\beta-1)} \oplus m_\beta$ must equal $\pi^{-1}(Z_\gamma)$. Any new true collision can either give a linear equation that is linearly dependent on the linear equations due to previously discovered true collisions, or it may give an independent linear equation. True collisions of the latter type are referred as *accidents*. At a high level, accidents denote the “surprising” collisions in CBC function computation. Obviously, the number of true collisions is at least the number of accidents. The following definition due to Jha and Nandi [11] gives a formula for the number of accidents.

Definition A.1 ([11]). *Consider the structure graph $\mathcal{G}_\pi(\tilde{m})$ associated with the message tuple \tilde{m} and permutation π . Let $\mathcal{S}(\mathcal{G}_\pi(\tilde{m}))$ be the system of linear equations formed by the true collisions of $\mathcal{G}_\pi(\tilde{m})$, and let r denote the rank of $\mathcal{S}(\mathcal{G}_\pi(\tilde{m}))$. Let $\text{Acc}(\mathcal{G}_\pi(\tilde{m}))$ be the set of accidents of $\mathcal{G}_\pi(\tilde{m})$. Then, the number of accidents is defined as*

$$|\text{Acc}(\mathcal{G}_\pi(\tilde{m}))| = \begin{cases} r + 1 & \text{if } \mathcal{G}_\pi(\tilde{m}) \text{ has a zero collision,} \\ r & \text{otherwise.} \end{cases}$$

Example A.2. Consider the structure graph from Figure A.1. Here, we have two true collisions, namely $((1, 0), (1, 2); (1, 1))$ and $((1, 0), (1, 2); (1, 5))$, and the associated system of equations is

$$Z_1[0] \oplus Z_1[2] = m_1[1] \oplus m_1[3]$$

$$Z_1[0] \oplus Z_1[2] = m_1[5] \oplus m_2[1]$$

Clearly, the two equations are dependent. So the graph has just one accident, and that accident is $((1,0), (1,2); (1,1))$, since it occurs before $((1,0), (1,2); (1,5))$. We encourage the readers to see [24,11] for further exposition on true collisions and accidents.

EXISTING RESULTS ON STRUCTURE GRAPHS: We now recall some known and useful combinatorial results on structure graphs. The proof of these results are already available in [24,11].

Lemma A.1 ([11]). *For any structure graph G , if there is a vertex α with in-degree d then $\text{Acc}(G) \geq d - 1$. Moreover, if the graph has a zero collision then $\text{Acc}(G) \geq d$.*

Lemma A.2 ([24,11]). *The number of structures graphs associated to \tilde{m} with a accidents is at most $\binom{\sigma}{2}^a$. In particular, there exists exactly one structure graph with 0 accidents.*

Lemma A.3 ([24,11]). *For any structure graph G with a accidents, we have*

$$\Pr_{\mathbf{P}}(\mathcal{G}_{\mathbf{P}}(\tilde{m}) = G) \leq \frac{1}{(2^n - q - \sigma)^a},$$

where \mathbf{P} denotes the partial function, introduced in Phase II of the ideal world sampling (see section 4.1), that samples each new point in a without replacement manner from a set of size $(2^n - q)$.

Proof. The proof of this lemma is identical to the proof of [11, Lemma 2], with a small change in the probability distribution. For all $\alpha \in \mathcal{Q}^+$, the values for $Z_{\text{in}(\alpha)}$ are now chosen in a without replacement manner from a set of size $(2^n - q)$ (instead of 2^n in case of [11, Lemma 2]).

Corollary A.1 ([24,11]). *For $a \in \mathbb{N}$ and $q + \sigma < 2^{n-1}$, we have*

$$\Pr_{\mathbf{P}}(\text{Acc}(\mathcal{G}_{\mathbf{P}}(\tilde{m})) \geq a) \leq \frac{\sigma^{2a}}{2^{an}},$$

AN EXTENSION OF LEMMA A.3: In addition to the system of linear equations $\mathcal{S}(\mathcal{G}_{\mathbf{P}}(\tilde{m}))$, sometimes we also consider additional equations over the intermediate output variables tailor-made to our analysis. Suppose the system of these additional equations is denoted simply by $\mathcal{S}'(\tilde{Z})$, and $\mathcal{S}'(\tilde{Z}) \cup \mathcal{S}(\mathcal{G}_{\mathbf{P}}(\tilde{m}))$ denotes the system consisting of equations from both $\mathcal{S}(\mathcal{G}_{\mathbf{P}}(\tilde{m}))$ and $\mathcal{S}'(\tilde{Z})$. Let r denote the rank of the combined system of equations, $\mathcal{S}'(\tilde{Z}) \cup \mathcal{S}(\mathcal{G}_{\mathbf{P}}(\tilde{m}))$.

Suppose $|\tilde{Z}| = t \leq \sigma$ and let 1_0 be an indicator variable that results in 1 if $\mathcal{G}_{\mathbf{P}}(\tilde{m})$ has a zero collision, and 0 otherwise. Then, we can have at most $(2^n - q)_{t+1_0}$ valid intermediate output tuples, since we must have $Z_{\text{in}(\alpha)} \neq Z_{\text{in}(\beta)}$

whenever $\text{in}(\alpha) \neq \text{in}(\beta)$, whence the probability to realize one such valid intermediate output tuple is $1/(2^n - q)_{t+1_0}$. Since the rank of $\mathcal{S}'(\tilde{Z}) \cup \mathcal{S}(\mathcal{G}_\pi(\tilde{m}))$ is r , by simple linear algebraic argument we know that by choosing the value of any $t - r$ variables, we may get a unique solution for the other r variables, in such a way that it also realizes $\mathcal{G}_\pi(\tilde{m})$ and satisfies $\mathcal{S}'(\tilde{Z})$. Thus, the probability to realize an intermediate output tuple that results in $\mathcal{G}_\pi(\tilde{m})$ and also satisfies $\mathcal{S}'(\tilde{Z})$ is bounded by at most $(2^n - q)_{t-r}/(2^n - q)_{t+1_0}$. We summarize this discussion in the following result which extends Lemma A.3. We remark that a similar result has already been proved in [25].

Lemma A.4. *For any structure graph G and additional system of equations $\mathcal{S}'(\tilde{Z})$, we have*

$$\Pr_{\mathbb{P}} \left(\mathcal{G}_{\mathbb{P}}(\tilde{m}) = G \wedge \mathcal{S}'(\tilde{Z}) \text{ is satisfied} \right) \leq \frac{1}{(2^n - q - \sigma + r)^{r+1_0}},$$

where r denotes the rank of $\mathcal{S}'(\tilde{Z}) \cup \mathcal{S}(\mathcal{G}_{\mathbb{P}}(\tilde{m}))$ and 1_0 is an indicator variable that results in 1 if $\mathcal{G}_{\mathbb{P}}(\tilde{m})$ has a zero collision, and 0 otherwise.

Proof. The result follows from the preceding discussion by using the fact that $t < \sigma$.

B Bounding $\text{BadX} | \neg(\text{BadT} \vee \text{unman} \vee \text{BadW})$

Let $\text{TUW} := \neg(\text{BadT} \vee \text{unman} \vee \text{BadW})$, and $\text{Fi} = \neg(\text{TUW} \vee \text{BadX1} \vee \dots \vee \text{BadXi})$. We have

$$\begin{aligned} \Pr(\text{BadX} | \text{TUW}) &\leq \Pr(\text{BadX1} | \text{TUW}) + \Pr(\text{BadX2} | \text{F1}) + \Pr(\text{BadX3} | \text{F2}) + \Pr(\text{BadX4} | \text{F3}) \\ &\quad + \Pr(\text{BadX5} | \text{F4}) + \Pr(\text{BadX6} | \text{F5}) + \Pr(\text{BadX7} | \text{F6}) \end{aligned} \tag{28}$$

We bound the individual terms on the right hand side as follows:

Bounding $\Pr(\text{BadX1} | \text{TUW})$: Fix some $i, j, k \in [q]$. Since $\neg \text{unman}$ holds, we must have $\text{Acc}(\mathcal{G}_{\mathbb{P}}(\mathbf{M}_i, \mathbf{M}_j, \mathbf{M}_k)) \leq 2$. Accordingly, we have the following three cases:

- (A) $\text{Acc}(\mathcal{G}_{\mathbb{P}}(\mathbf{M}_i, \mathbf{M}_j, \mathbf{M}_k)) = 2$: This can be bounded by using exactly the same argument as used in Case A for $\text{BadW3} | \text{E2}$. So, we have

$$\Pr(\text{BadX1} \wedge \text{Case A} | \text{TUW}) \leq \frac{q^3 \ell^5}{2(2^n - q - \sigma + 3)^3}.$$

- (B) $\text{Acc}(\mathcal{G}_{\mathbb{P}}(\mathbf{M}_i, \mathbf{M}_j, \mathbf{M}_k)) = 1$: Suppose (α_1, β_1) and (α_2, β_2) are collision source leading to the accident. Assuming $a < \ell_i - 1$ and $0 < b < \ell_k - 1$, we have the following system of equations

$$\mathbf{Z}_{\alpha_1}[\beta_1] \oplus \mathbf{Z}_{\alpha_2}[\beta_2] = \mathbf{M}_{\alpha_1}[\beta_1 + 1] \oplus \mathbf{M}_{\alpha_2}[\beta_2 + 1]$$

$$\begin{aligned} Z_i[a-1] \oplus \mu_{\delta_{M_j}} \odot L \oplus Z_j[\ell_j-1] &= \overline{M}_j[\ell_j] \oplus M_i[a] \\ Z_k[b-1] &= M_i[a+1] \oplus M_k[b] \oplus T_j \end{aligned}$$

Since the three equations involve an odd number of Z variables and $\neg\text{BadW1}$ holds, we can straightaway conclude that the system has rank 3. Note that it is true irrespective of our choice of a , and k as long as $b > 0$. So, we have

$$\Pr(\text{BadX1} \wedge \text{Case B} \wedge b > 0 | \text{TUW}) \leq \frac{q^3 \ell^4}{2(2^n - q - \sigma + 3)^3}.$$

If $b = 0$. The last equation simply boils down to the condition $T_j = M_i[a+1]$. This gives a unique choice of j , once we fix i, a . Further, there are q choices for k , 3 choices for (α_1, α_2) , and ℓ^2 choices for (β_1, β_2) . So, we get

$$\Pr(\text{BadX1} \wedge \text{Case B} \wedge b = 0 | \text{TUW}) \leq \frac{q^2 \ell^3}{2(2^n - q - \sigma + 2)^2}.$$

- (C) $\text{Acc}(\mathcal{G}_P(M_i, M_j, M_k)) = 0$: Assuming $a < \ell_i - 1$ and $0 < b < \ell_k - 1$, we have the following system of equations

$$\begin{aligned} Z_j[a-1] \oplus \mu_{\delta_{M_j}} \odot L \oplus Z_j[\ell_j-1] &= \overline{M}_j[\ell_j] \oplus M_i[a] \\ Z_k[b-1] &= M_i[a+1] \oplus M_k[b] \oplus T_j \end{aligned}$$

Using the same argument as applied in case B above, we get

$$\Pr(\text{BadX1} \wedge \text{Case C} | \text{TUW}) \leq \frac{q^3 \ell^2}{6(2^n - q - \sigma + 2)^2}.$$

For $b = 0$, using the previous argument, we get

$$\Pr(\text{BadX1} \wedge \text{Case B} \wedge b = 0 | \text{TUW}) \leq \frac{\sigma}{2^n - q - \sigma + 1}.$$

By combining all three cases, we have

$$\Pr(\text{BadX1} | \text{TUW}) \leq \frac{2\sigma}{2^n} + \frac{q^3 \ell^5}{(2^n - q - \sigma + 3)^3} + \frac{q^3 \ell^2}{6(2^n - q - \sigma + 2)^2} + \frac{q^2 \ell^3}{2(2^n - q - \sigma + 2)^2}. \quad (29)$$

Bounding $\Pr(\text{BadX2} | \text{F1})$: First note that once we fix i and k , $a = \ell_i - \ell_k$, and $T_j = M_k[1] \oplus M_i[a+1]$ which in combination with $\neg\text{BadW2}$ gives a unique choice for j . So we have at most $q^2/2$ choices for (i, j, k, a) . Now, considering $W_i[a] = W_j[\ell_j]$, we have the equation

$$Z_j[a-1] \oplus \mu_{\delta_{M_j}} \odot L \oplus Z_j[\ell_j-1] = \overline{M}_j[\ell_j] \oplus M_i[a]$$

Using Lemma A.4, we get

$$\Pr(\text{BadX2} | \text{F1}) \leq \frac{q^2}{2(2^n - q - \sigma + 1)}. \quad (30)$$

Bounding $\Pr(\text{BadX3}|\text{F2})$: Recalling the counting argument of previous case, we know that j is fixed once we choose (i, k, a) . Now, we can have three cases:

- (A) $\text{Acc}(\mathcal{G}_P(M_i, M_j, M_k, M_l)) \geq 2$: We consider the system, consisting of any two accident equations and $W_i[a] = W_j[\ell_j]$. Using similar argument as used in Case A for $\text{BadW3}|\text{E2}$, and the counting argument, we get

$$\Pr(\text{BadX3} \wedge \text{Case A}|\text{F2}) \leq \frac{3q^3 \ell^5}{(2^n - q - \sigma + 3)^3}.$$

- (B) $\text{Acc}(\mathcal{G}_P(M_i, M_j, M_k, M_l)) = 1$: Suppose (α_1, β_1) and (α_2, β_2) are collision source leading to the accident. We have the following system of equations

$$\begin{aligned} Z_{\alpha_1}[\beta_1] \oplus Z_{\alpha_2}[\beta_2] &= M_{\alpha_1}[\beta_1 + 1] \oplus M_{\alpha_2}[\beta_2 + 1] \\ Z_i[a - 1] \oplus \mu_{\delta_{M_j}} \odot L \oplus Z_j[\ell_j - 1] &= \overline{M}_j[\ell_j] \oplus M_i[a] \\ Z_k[b - 1] \oplus \mu_{\delta_{M_l}} \odot L \oplus Z_l[\ell_l - 1] &= \overline{M}_l[\ell_l] \oplus M_k[b] \end{aligned}$$

We can have two sub-cases:

- (B.1) Suppose the third equation is a consequence of the second equation. Using a similar line of arguments as used in case B.1 of $\text{BadW5}|\text{E4}$, we can conclude that a and b have some fixed choices. In particular, a has at most 18 choices. Then, using Lemma A.4, we have

$$\Pr(\text{BadX3} \wedge \text{Case B.1}|\text{F2}) \leq \frac{3q^3 \ell^2}{(2^n - q - \sigma + 2)^2}.$$

- (B.2) The last two equations are independent. Then, we can simply consider these two equations and ignore the accident equation. Using Lemma A.4, we have

$$\Pr(\text{BadX3} \wedge \text{Case B.2}|\text{F2}) \leq \frac{q^3 \ell^2}{6(2^n - q - \sigma + 2)^2}.$$

- (C) $\text{Acc}(\mathcal{G}_P(M_i, M_j, M_k)) = 0$: We have the following system of equations

$$\begin{aligned} Z_i[a - 1] \oplus \mu_{\delta_{M_j}} \odot L \oplus Z_j[\ell_j - 1] &= \overline{M}_j[\ell_j] \oplus M_i[a] \\ Z_k[b - 1] \oplus \mu_{\delta_{M_l}} \odot L \oplus Z_l[\ell_l - 1] &= \overline{M}_l[\ell_l] \oplus M_k[b] \end{aligned}$$

Let r denote the rank of the above system. If $r = 1$, then we must have $\mu_{\delta_{M_j}} = \mu_{\delta_{M_l}}$ and $\overline{M}_j[\ell_j] \oplus M_i[a] \oplus \overline{M}_l[\ell_l] \oplus M_k[b] = 0^n$. Now, we can have two sub-cases:

- (a) $M_i[1, \dots, a-1] = M_k[1, \dots, b-1]$ and $M_j[1, \dots, \ell_j-1] = M_l[1, \dots, \ell_l-1]$.
(b) $M_i[1, \dots, a-1] = M_l[1, \dots, \ell_l-1]$ and $M_j[1, \dots, \ell_j-1] = M_k[1, \dots, b-1]$.

In both cases, we consider the first equation for probability calculation. The two cases are similar. So we only handle the second case. The key idea here is to note that the i -th message shares $a - 1$ block prefix with l -th message and similarly k -th message shares $b - 1$ block prefix with j -th message.

Let $\mathcal{S} = \{M_i[1, \dots, j] : (i, j) \in [q] \times [\ell_i]\} \cup \{\perp\}$ be the set of all prefixes over all q messages. For any $x \in \mathcal{S}$, we define $\text{pred}(x) = x'$, where $x = x' \| y$ for some $y \in \mathcal{B}$. Consider a rooted tree \mathcal{T} with vertices from \mathcal{S} where \perp acts as the root of the tree and there's a directed edge $(\text{pred}(x), x)$ for each $x \in \mathcal{S} \setminus \{\perp\}$. Clearly, it is a directed tree rooted at \perp and all leaf nodes are the $1 \leq i \leq q$ messages M_i . Except the leaves, all nodes have out-degree (denoted d_x for the node x) at least one. A node is called *fork node* if it has out-degree at least two. Let Fork denote the set of all fork nodes. Now, $\sum_{x \in \mathcal{S}} d_x = \sigma - 1$. Then, we make the following claim.

Claim. $\sum_{x \in \text{Fork}} (d_x - 1) = q - 1$, and $|\text{Fork}| \leq q - 1$.

Proof. All edges at the node x introduce exactly new $d_x - 1$ messages. So the first part is done. The second part also follows from the same argument. A new message is introduced only when we have a forking. So the number of messages should be at least the number of forking points.

Now let us come back to our problem. We have to choose two prefixes u and v of length $a - 1$ and $b - 1$ blocks, respectively. Now, clearly u and v must be fork nodes in \mathcal{T} , as each of them have out-degree at least 2. We know that the number of fork nodes can be at most $q - 1$ and for every such choices the probability that first equation holds is at most $1/(2^n - q - \sigma + 1)$. So, total probability for having this types of equation is at most

$$\sum_{u, v \in \text{Fork}} \frac{d_u d_v}{2^n - q - \sigma + 1} \leq \frac{1}{2^n - q - \sigma + 1} \left[\sum_{u \in \text{Fork}} (d_u - 1) \right]^2 \leq \frac{q^2}{2^n - q - \sigma + 1}.$$

So, combining the bound for the two sub-cases, we get

$$\Pr(\text{BadX3} \wedge \text{Case C} \wedge r = 1 | \text{F2}) \leq \frac{2q^2}{2^n - q - \sigma + 1}.$$

If $r = 2$, we get

$$\Pr(\text{BadX3} \wedge \text{Case C} \wedge r = 2 | \text{F2}) \leq \frac{q^3 \ell^2}{6(2^n - q - \sigma + 2)^2}.$$

Combining all three cases, we get

$$\Pr(\text{BadX3} | \text{F2}) \leq \frac{3q^3 \ell^5}{(2^n - q - \sigma + 3)^3} + \frac{20q^3 \ell^2}{6(2^n - q - \sigma + 2)^2} + \frac{2q^2}{2^n - q - \sigma + 1}. \quad (31)$$

Bounding $\Pr(\text{BadX4} | \text{F3})$: $(i, a), (j, b) \in \widetilde{\text{FCI}}$ if and only if there exists $k, l \in [q]$, such that $W_i[a] = W_k[\ell_k]$ and $W_j[b] = W_l[\ell_l]$. We first note that fixing (i, a) ,

(j, b) , and anyone of k and l fixed the remaining index, since $\mathsf{T}_k \oplus \mathsf{T}_l \oplus \mathsf{M}_i[a] \oplus \mathsf{M}_j[b]$. So we can have at most $q^3\ell^2/6$ choices for (i, j, k, l, a, b) . As in the case of $\mathsf{BadX3|F2}$, we can have three cases:

- (A) $\text{Acc}(\mathcal{G}_{\mathsf{P}}(\mathsf{M}_i, \mathsf{M}_j, \mathsf{M}_k, \mathsf{M}_l)) \geq 2$: Using similar argument as used in Case A for $\mathsf{BadX3|F2}$, and the counting argument, we get

$$\Pr(\mathsf{BadX4} \wedge \text{Case A} | \mathsf{F3}) \leq \frac{5q^3\ell^6}{(2^n - q - \sigma + 3)^3}.$$

- (B) $\text{Acc}(\mathcal{G}_{\mathsf{P}}(\mathsf{M}_i, \mathsf{M}_j, \mathsf{M}_k, \mathsf{M}_l)) = 1$: Suppose (α_1, β_1) and (α_2, β_2) are collision source leading to the accident. We have the following system of equations

$$\begin{aligned} \mathsf{Z}_{\alpha_1}[\beta_1] \oplus \mathsf{Z}_{\alpha_2}[\beta_2] &= \mathsf{M}_{\alpha_1}[\beta_1 + 1] \oplus \mathsf{M}_{\alpha_2}[\beta_2 + 1] \\ \mathsf{Z}_i[a - 1] \oplus \mu_{\delta_{\mathsf{M}_k}} \odot \mathsf{L} \oplus \mathsf{Z}_k[\ell_k - 1] &= \overline{\mathsf{M}}_k[\ell_k] \oplus \mathsf{M}_i[a] \\ \mathsf{Z}_j[b - 1] \oplus \mu_{\delta_{\mathsf{M}_l}} \odot \mathsf{L} \oplus \mathsf{Z}_l[\ell_l - 1] &= \overline{\mathsf{M}}_l[\ell_l] \oplus \mathsf{M}_j[b] \end{aligned}$$

We can have two sub-cases:

- (B.1) Suppose the third equation is a consequence of the second equation. Using prefix and suffix backtracing arguments, we can conclude that a, b, β_1 and β_2 can be chosen in at most $18\ell^2$ ways, once we fix α_1 and α_2 . Then, using Lemma A.4, we have

$$\Pr(\mathsf{BadX4} \wedge \text{Case B.1} | \mathsf{F3}) \leq \frac{3q^3\ell^2}{(2^n - q - \sigma + 2)^2}.$$

- (B.2) The last two equations are independent. Then, we can simply consider these two equations and ignore the accident equation. Using Lemma A.4, we have

$$\Pr(\mathsf{BadX4} \wedge \text{Case B.2} | \mathsf{F3}) \leq \frac{q^3\ell^2}{6(2^n - q - \sigma + 2)^2}.$$

- (C) $\text{Acc}(\mathcal{G}_{\mathsf{P}}(\mathsf{M}_i, \mathsf{M}_j, \mathsf{M}_k)) = 0$: We have the following system of equations

$$\begin{aligned} \mathsf{Z}_i[a - 1] \oplus \mu_{\delta_{\mathsf{M}_j}} \odot \mathsf{L} \oplus \mathsf{Z}_j[\ell_j - 1] &= \overline{\mathsf{M}}_j[\ell_j] \oplus \mathsf{M}_i[a] \\ \mathsf{Z}_k[b - 1] \oplus \mu_{\delta_{\mathsf{M}_l}} \odot \mathsf{L} \oplus \mathsf{Z}_l[\ell_l - 1] &= \overline{\mathsf{M}}_l[\ell_l] \oplus \mathsf{M}_k[b] \end{aligned}$$

Using similar arguments as in case C of $\mathsf{BadX3|F2}$, we get

$$\Pr(\mathsf{BadX4} \wedge \text{Case C} \wedge r = 1 | \mathsf{F3}) \leq \frac{q^3\ell^2}{6(2^n - q - \sigma + 2)^2} + \frac{2q^2}{2^n - q - \sigma + 1}.$$

Combining all three cases, we get

$$\Pr(\mathsf{BadX4} | \mathsf{F3}) \leq \frac{5q^3\ell^6}{(2^n - q - \sigma + 3)^3} + \frac{20q^3\ell^2}{6(2^n - q - \sigma + 2)^2} + \frac{2q^2}{2^n - q - \sigma + 1}. \quad (32)$$

Bounding $\Pr(\text{BadX5}|\text{F4})$: Similar counting argument as used in previous cases apply here as well, i.e., index j is fixed once we choose (i, k, a) . First, we handle the corner case, when $b = 0$. In this case we get the system of equations

$$\begin{aligned} Z_i[a-1] \oplus \mu_{\delta_{M_j}} \odot L \oplus Z_j[\ell_j-1] &= \bar{M}_j[\ell_j] \oplus M_i[a] \\ Z_k[p+1] &= M_i[a+2+p] \oplus \star \end{aligned}$$

where $\star = 0^n$ if $a+2+p \neq \ell_i$, and $\mu_{\delta_{M_i}} \odot L$ otherwise. In both the cases, the two equations are independent. So we get

$$\Pr(\text{BadX5} \wedge b=0|\text{F4}) \leq \frac{q^2 \ell}{2(2^n - q - \sigma + 2)^2}.$$

Assuming $b > 0$, we can have three cases:

- (A) $\text{Acc}(\mathcal{G}_P(M_i, M_j, M_k, M_l)) \geq 2$: Using similar argument as used in Case A for $\text{BadX3}|\text{F2}$, and the counting argument, we get

$$\Pr(\text{BadX5} \wedge \text{Case A}|\text{F4}) \leq \frac{3q^3 \ell^5}{(2^n - q - \sigma + 3)^3}.$$

- (B) $\text{Acc}(\mathcal{G}_P(M_i, M_j, M_k, M_l)) = 1$: Suppose (α_1, β_1) and (α_2, β_2) are collision source leading to the accident. Assuming $a+2+p < \ell_i$ and $b < \ell_l$, we have the following system of equations

$$\begin{aligned} Z_{\alpha_1}[\beta_1] \oplus Z_{\alpha_2}[\beta_2] &= M_{\alpha_1}[\beta_1+1] \oplus M_{\alpha_2}[\beta_2+1] \\ Z_i[a-1] \oplus \mu_{\delta_{M_j}} \odot L \oplus Z_j[\ell_j-1] &= \bar{M}_j[\ell_j] \oplus M_i[a] \\ Z_k[p+1] \oplus Z_l[b-1] &= M_l[b] \oplus M_i[a+2+p] \end{aligned}$$

Now, we can have two cases:

- B.1 Third equation is a consequence of the first equation. Then, using the previously used prefix and suffix backtracing arguments, we can conclude that β_1 and β_2 have fixed choices. In particular, we have at most 3 choices for (β_1, β_2) . So, using Lemma A.4, we get

$$\Pr(\text{BadX5} \wedge \text{Case B.1}|\text{F4}) \leq \frac{3q^3 \ell^2}{(2^n - q - \sigma + 2)^2}.$$

- B.2 Third equation is independent of the first equation. Then, we simply consider the second and third equation, which are obviously independent. So, using Lemma A.4, we get

$$\Pr(\text{BadX5} \wedge \text{Case B.2}|\text{F4}) \leq \frac{q^3 \ell^2}{6(2^n - q - \sigma + 2)^2}.$$

Now, assume $a = \ell_i - p - 2$ and consider the two equations:

$$Z_{\alpha_1}[\beta_1] \oplus Z_{\alpha_2}[\beta_2] = M_{\alpha_1}[\beta_1+1] \oplus M_{\alpha_2}[\beta_2+1]$$

$$Z_i[\ell_i - p - 3] \oplus \mu_{\delta_{M_j}} \odot L \oplus Z_j[\ell_j - 1] = \overline{M}_j[\ell_j] \oplus M_i[\ell_i - p - 2]$$

The two equations are obviously independent due to the presence of L . So we get

$$\Pr(\text{BadX5} \wedge \text{Case B} \wedge a = \ell_i - p - 2 | \mathbf{F4}) \leq \frac{q^3 \ell^2}{(2^n - q - \sigma + 2)^2}.$$

The case where $a < \ell_i = p - 2$ and $b = \ell_l$ is similarly bounded.

(C) $\text{Acc}(\mathcal{G}_P(M_i, M_j, M_k)) = 0$: Assuming $a + 2 + p < \ell_i$ and $b < \ell_l$, we have the following system of equations

$$\begin{aligned} Z_i[a - 1] \oplus \mu_{\delta_{M_j}} \odot L \oplus Z_j[\ell_j - 1] &= \overline{M}_j[\ell_j] \oplus M_i[a] \\ Z_k[p + 1] \oplus Z_l[b - 1] &= M_l[b] \oplus M_i[a + 2 + p] \end{aligned}$$

The two equations are clearly independent due to the presence of L , whence we have

$$\Pr(\text{BadX5} \wedge \text{Case C} \wedge a < \ell_i - p - 2 \wedge b < \ell_l | \mathbf{F4}) \leq \frac{q^3 \ell^2}{6(2^n - q - \sigma + 2)^2}.$$

Suppose $a = \ell_i - p - 2$. Then we can simply consider the first equation, whence we get

$$\Pr(\text{BadX5} \wedge \text{Case C} \wedge a = \ell_i - p - 2 | \mathbf{F4}) \leq \frac{q^2}{2(2^n - q - \sigma + 1)}.$$

Finally assume $a < \ell_i - p - 2$ and $b = \ell_l$. Then, we have the following system of equations

$$\begin{aligned} Z_i[a - 1] \oplus \mu_{\delta_{M_j}} \odot L \oplus Z_j[\ell_j - 1] &= \overline{M}_j[\ell_j] \oplus M_i[a] \\ Z_k[p + 1] \oplus \mu_{\delta_{M_l}} \odot L \oplus Z_l[\ell_l - 1] &= \overline{M}_l[\ell_l] \oplus M_i[a + 2 + p] \end{aligned}$$

Using similar argument as in case C of BadX3|F2 , we get

$$\Pr(\text{BadX5} \wedge \text{Case C} \wedge a < \ell_i - p - 2 \wedge b = \ell_l | \mathbf{F4}) \leq \frac{2q^2}{2^n - q - \sigma + 1} + \frac{q^3 \ell^2}{6(2^n - q - \sigma + 2)^2}.$$

Combining all three cases, we get

$$\Pr(\text{BadX5} | \mathbf{F4}) \leq \frac{3q^3 \ell^5}{(2^n - q - \sigma + 3)^3} + \frac{27q^3 \ell^2}{6(2^n - q - \sigma + 2)^2} + \frac{3q^2}{(2^n - q - \sigma + 1)}. \quad (33)$$

Bounding $\Pr(\text{BadX6} | \mathbf{F5})$: $(i, a) \in \widetilde{\text{FCI}}$ and $(j, b) \in \widetilde{\text{ICI}}$ if and only if there exists $i', k, l \in [q]$ and $c \in [\ell_k - 1]$, such that $W_i[a] = W_{i'}[\ell_{i'}]$, $W_j[b - c] = W_l[\ell_l]$ and $X_j[b - c + 1] = M_k[1]$. We first note that fixing (i, a) , $j, k, b - c$, and i' fixes b, c , and l . So we can have at most $q^4 \ell^2 / 12$ choices for $(i, j, k, i', l, a, b, c)$. As in the case of BadX3|F2 , we can have three cases:

- (A) $\text{Acc}(\mathcal{G}_P(M_i, M_j, M_k, M_l, M_{i'})) \geq 2$: In this case we consider the two accident equations along with $W_i[a] = W_{i'}[\ell_{i'}]$ and $X_i[a+1] = X_j[b+1]$. We claim that all four equations are independent due to odd number of Z variables (last equation is univariate in $Z_k[c]$). Then, we get

$$\Pr(\text{BadX6} \wedge \text{Case A} | \text{F5}) \leq \frac{4q^4\ell^6}{(2^n - q - \sigma + 4)^4}.$$

- (B) $\text{Acc}(\mathcal{G}_P(M_i, M_j, M_k, M_l, M_{i'})) = 1$: Suppose (α_1, β_1) and (α_2, β_2) are collision source leading to the accident. We have the following system of equations

$$\begin{aligned} Z_{\alpha_1}[\beta_1] \oplus Z_{\alpha_2}[\beta_2] &= M_{\alpha_1}[\beta_1 + 1] \oplus M_{\alpha_2}[\beta_2 + 1] \\ Z_i[a-1] \oplus \mu_{\delta_{M_{i'}}} \odot L \oplus Z_{i'}[\ell_{i'} - 1] &= \overline{M}_{i'}[\ell_{i'}] \oplus M_i[a] \\ Z_j[b-c-1] \oplus \mu_{\delta_{M_l}} \odot L \oplus Z_l[\ell_l - 1] &= \overline{M}_l[\ell_l] \oplus M_j[b-c] \\ Z_k[c] &= M_j[b+1] \oplus M_i[a+1] \oplus T_{i'} \end{aligned}$$

Again concentrating on whether the second and third equations are independent or not, and using similar argumentation as before, we get

$$\Pr(\text{BadX6} \wedge \text{Case B} | \text{F5}) \leq \frac{5q^4\ell^3}{(2^n - q - \sigma + 2)^3}.$$

- (C) $\text{Acc}(\mathcal{G}_P(M_i, M_j, M_k)) = 0$: We have the following system of equations

$$\begin{aligned} Z_i[a-1] \oplus \mu_{\delta_{M_{i'}}} \odot L \oplus Z_{i'}[\ell_{i'} - 1] &= \overline{M}_{i'}[\ell_{i'}] \oplus M_i[a] \\ Z_j[b-c-1] \oplus \mu_{\delta_{M_l}} \odot L \oplus Z_l[\ell_l - 1] &= \overline{M}_l[\ell_l] \oplus M_j[b-c] \\ Z_k[c] &= M_j[b+1] \oplus M_i[a+1] \oplus T_{i'} \end{aligned}$$

Using similar arguments as in previous cases, we get

$$\Pr(\text{BadX6} \wedge \text{Case C} | \text{F5}) \leq \frac{q^4\ell^2}{4(2^n - q - \sigma + 3)^3} + \frac{2q^3\ell}{(2^n - q - \sigma + 2)^2}.$$

Combining all three cases, we get

$$\Pr(\text{BadX6} | \text{F5}) \leq \frac{4q^4\ell^6}{(2^n - q - \sigma + 4)^4} + \frac{6q^4\ell^3}{(2^n - q - \sigma + 3)^3} + \frac{2q^3\ell}{(2^n - q - \sigma + 2)^2}. \quad (34)$$

Bounding $\Pr(\text{BadX7} | \text{F6})$: $(i, a), (j, b) \in \widetilde{\text{ICl}}$ if and only if there exists $i', j', k, l \in \overline{[q]}$, $c \in \overline{[\ell_k]}$ and $d \in \overline{[\ell_l]}$, such that $W_i[a] = W_{i'}[\ell_{i'}]$, $W_j[b-c] = W_{j'}[\ell'_{j'}]$, $X_i[a-c+1] = M_k[1]$ and $X_j[b-d+1] = M_l[1]$. We first note that fixing $i, j, k, l, a-c$, and $b-d$, fixes a, b, c, d, i' and j' . So we can have at most $q^4\ell^2/12$ choices for $(i, j, k, l, i', j', a, b, c, d)$. As in the case of $\text{BadX3} | \neg \text{F2}$, we can have three cases:

- (A) $\text{Acc}(\mathcal{G}_P(\mathbf{M}_i, \mathbf{M}_j, \mathbf{M}_k, \mathbf{M}_l, \mathbf{M}_{i'}, \mathbf{M}_{j'})) \geq 2$: In this case we consider the two accident equations along with $W_i[a] = W_{i'}[\ell_{i'}]$, $W_j[b] = W_{j'}[\ell_{j'}]$. As in the previous cases, we conclude that all four equations are independent. Then, we get

$$\Pr(\text{BadX7} \wedge \text{Case A} | \text{F6}) \leq \frac{4q^4 \ell^6}{(2^n - q - \sigma + 4)^4}.$$

- (B) $\text{Acc}(\mathcal{G}_P(\mathbf{M}_i, \mathbf{M}_j, \mathbf{M}_k, \mathbf{M}_l, \mathbf{M}_{i'})) = 1$: Suppose (α_1, β_1) and (α_2, β_2) are collision source leading to the accident. We have the following system of equations

$$\begin{aligned} Z_{\alpha_1}[\beta_1] \oplus Z_{\alpha_2}[\beta_2] &= \mathbf{M}_{\alpha_1}[\beta_1 + 1] \oplus \mathbf{M}_{\alpha_2}[\beta_2 + 1] \\ Z_i[a - c - 1] \oplus \mu_{\delta_{\mathbf{M}_{i'}}} \odot \mathbf{L} \oplus Z_i[\ell_{i'} - 1] &= \overline{\mathbf{M}}_{i'}[\ell_{i'}] \oplus \mathbf{M}_i[a - c] \\ Z_j[b - d - 1] \oplus \mu_{\delta_{\mathbf{M}_{j'}}} \odot \mathbf{L} \oplus Z_j[\ell_{j'} - 1] &= \overline{\mathbf{M}}_{j'}[\ell_{j'}] \oplus \mathbf{M}_j[b - d] \\ Z_k[c] \oplus Z_l[d] &= \mathbf{M}_j[b + 1] \oplus \mathbf{M}_i[a + 1] \oplus \mathbf{T}_{i'} \oplus \mathbf{T}_{j'} \end{aligned}$$

Again concentrating on whether the second and third equations are independent or not, and using similar argumentation as before, we get

$$\Pr(\text{BadX7} \wedge \text{Case B} | \text{F6}) \leq \frac{6q^4 \ell^2}{(2^n - q - \sigma + 2)^3}.$$

- (C) $\text{Acc}(\mathcal{G}_P(\mathbf{M}_i, \mathbf{M}_j, \mathbf{M}_k)) = 0$: We have the following system of equations

$$\begin{aligned} Z_i[a - c - 1] \oplus \mu_{\delta_{\mathbf{M}_{i'}}} \odot \mathbf{L} \oplus Z_{i'}[\ell_{i'} - 1] &= \overline{\mathbf{M}}_{i'}[\ell_{i'}] \oplus \mathbf{M}_i[a - c] \\ Z_j[b - d - 1] \oplus \mu_{\delta_{\mathbf{M}_{j'}}} \odot \mathbf{L} \oplus Z_{j'}[\ell_{j'} - 1] &= \overline{\mathbf{M}}_{j'}[\ell_{j'}] \oplus \mathbf{M}_j[b - d] \\ Z_k[c] \oplus Z_l[d] &= \mathbf{M}_j[b + 1] \oplus \mathbf{M}_i[a + 1] \oplus \mathbf{T}_{i'} \oplus \mathbf{T}_{j'} \end{aligned}$$

Using similar arguments as in previous cases, we get

$$\Pr(\text{BadX7} \wedge \text{Case C} | \text{F6}) \leq \frac{q^4 \ell^2}{4(2^n - q - \sigma + 3)^3} + \frac{2q^3}{(2^n - q - \sigma + 2)^2}.$$

Combining all three cases, we get

$$\Pr(\text{BadX7} | \text{F6}) \leq \frac{4q^4 \ell^6}{(2^n - q - \sigma + 4)^4} + \frac{7q^4 \ell^2}{(2^n - q - \sigma + 3)^3} + \frac{2q^3}{(2^n - q - \sigma + 2)^2}. \quad (35)$$

Finally, accumulating all the bounds from Eq. (29)-(35), and assuming $\ell < q$, we get Eq. (26), i.e.,

$$\begin{aligned} \Pr(\text{BadX} | \text{TUW}) &\leq \frac{2\sigma}{2^n} + \frac{10q^2}{2^n - q - \sigma + 1} + \frac{15q^3 \ell^2 + q^2 \ell^3}{(2^n - q - \sigma + 2)^2} \\ &\quad + \frac{12q^3 \ell^6 + 6q^4 \ell^3}{(2^n - q - \sigma + 3)^3} + \frac{8q^4 \ell^6}{(2^n - q - \sigma + 4)^4} \end{aligned}$$