Bayesian Spillover Graphs for Dynamic Networks: Supplementary Material

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A MOVING AVERAGE REPRESENTATION OF VAR(1)

We can rewrite a VAR(1) model with a moving average representation [\[Tsay, 2013\]](#page-4-0) using the mean-adjusted model, which is useful for computing variances of forecast errors.

We define the **mean-adjusted model** $\tilde{\mathbf{z}}_t = \mathbf{z}_t - \mu$, where $\mu = (I_d - \phi_1)^{-1} \phi_0$.

Then,

$$
\tilde{\mathbf{z}}_t = \mathbf{a}_t + \phi_1 \tilde{\mathbf{z}}_{t-1} \n= \mathbf{a}_t + \phi_1 (\mathbf{a}_{t-1} + \phi_1 \tilde{\mathbf{z}}_{t-2}) \n= \mathbf{a}_t + \phi_1 \mathbf{a}_{t-1} + \phi_1^2 (\mathbf{a}_{t-2} + \phi_1 \tilde{\mathbf{z}}_{t-3}) \n= \mathbf{a}_t + \phi_1 \mathbf{a}_{t-1} + \phi_1^2 \mathbf{a}_{t-2} + \phi_1^3 \mathbf{a}_{t-3} + \dots
$$

Hence,

$$
\mathbf{z}_t = \mu + \tilde{\mathbf{z}}_t
$$

= $\mu + \mathbf{a}_t + \phi_1 \mathbf{a}_{t-1} + \phi_1^2 \mathbf{a}_{t-2} + \phi_1^3 \mathbf{a}_{t-3} + \dots$
= $\mu + \mathbf{a}_t + \psi_1 \mathbf{a}_{t-1} + \psi_2 \mathbf{a}_{t-2} + \dots$
= $\mu + \sum_{i=0}^{\infty} \psi_i \mathbf{a}_{t-i}$

where $\psi_i = \phi_1^i$ for $i \geq 0$.

B PROOF OF THEOREM 1

Theorem 1. *If* ϕ_1 *is a DAG, then (1) no autocorrelation exists, (2)* ϕ_1 *can be specified by a strictly triangular matrix, (3) all eigenvalues of* ϕ_1 *are* 0 *and hence* z_t *is stationary.*

Proof: By definition of DAG, no cycles can exist in the adjacency matrix, in this case, ϕ_1 . Hence, the diagonal entries which indicate dependency of z_{it} on $z_{i,t+1}$ is necessarily 0, and thereby proving point (1).

Note that by definition, there exists a topological ordering on the vertices if and only if a graph has no directed cycles. Because ϕ_1 is a DAG, we can relabel the d vertices (time series components) as $v_1, v_2, ..., v_d$. If $v_i v_{i'}$ is a directed edge into i from i' (indicating Granger-causality), then $i > i'$. Hence, all entries above the main diagonal are also 0 because these are entries for which $i < i'$. Combined with point (1) where main diagonal entries are also 0, this satisfies the definition of a strictly lower-triangular matrix (2).

We've shown that the adjacency matrix of a DAG is strictly lower-triangular via permutation, and note that the order of individual time series components does not matter, although in this case the d vertices are ordered from source to sink nodes. The eigenvalues of any lower-triangular matrix is just its diagonal components [\[Axler, 1997\]](#page-4-1), meaning that all eigenvalues for ϕ_1 is just 0. Since these are strictly less than 1 in magnitude, we can conclude that z_t is stationary (3).

C EVALUATING ACCURACY FOR SOURCE & SINK NODE IDENTIFICATION

First, define Discounted Cumulative Gain (DCG) at position d , for d nodes arranged in a particular order:

$$
\text{DCG}_{\text{d}} = \sum_{i=1}^{d} \frac{rel_i}{\log_2(i+1)}
$$

where rel_i is the graded precision score of node at position i, e.g. $\{1, 0.5, 0\}$ for {source, intermediary, sink} nodes respectively. Greater penalty is given for source or sink nodes ranked in lower positions. NDCG [\[Valizadegan et al., 2009\]](#page-4-2) then equals DCG divided by Ideal Discounted Cumulative Gain (IDCG):

$$
\text{NDCG}_d = \frac{DCG_d}{IDCG_d}, \ \text{IDCG}_d = \sum_{i=1}^{|rel_d|} \frac{rel_i}{\log_2(i+1)}
$$

and $|rel_d|$ represents the optimal order of nodes, which is given by the ground truth labels of each node.

D BSG FOR IDENTIFYING SINK AND SOURCE NODES

D.1 ABLATION EXPERIMENT - ERROR COVARIANCE Σ_a

— indicates retrieved NGC graph is degenerate, e.g., only edges are self-directed.

† indicates network measure cannot distinguish between nodes, e.g., all in/out degrees are equal.

D.2 MULTISPECIES LOTKA-VOLTERRA - NONLINEAR DYNAMIC SYSTEMS

E EVALUATING KINCADE FIRE SPILLOVERS

Multi-species LV	$d = 20, T = 50$		$d = 20, T = 200$		$d = 20, T = 1000$	
Method	NDCG@20 Source (Predator)	NDCG@20 Sink (Prey)	Source (Predator)	Sink (Prey)	Source (Predator)	Sink (Prey)
BSG, $h=1$ BSG. $h=5$ BSG, $h = 10$	0.995 ± 0.004 0.995 ± 0.002 0.989 ± 0.01	0.865 ± 0.045 0.905 ± 0.046 0.946 ± 0.015	0.973 ± 0.013 0.945 ± 0.021 0.892 ± 0.058	0.939 ± 0.039 0.931 ± 0.047 0.907 ± 0.056	0.982 ± 0.015 0.967 ± 0.024 0.932 ± 0.031	0.811 ± 0.069 0.755 ± 0.035 0.711 ± 0.074
VAR-Between VAR-Closeness VAR-Degree VAR-Eigen	0.71 ± 0.058 0.781 ± 0.093 0.768 ± 0.091 0.812 ± 0.087	0.84 ± 0.058 0.768 ± 0.093 0.748 ± 0.071 0.738 ± 0.087	$0.721 + 0.145$ 0.78 ± 0.09 0.679 ± 0.084 0.881 ± 0.037	0.828 ± 0.145 0.769 ± 0.09 0.737 ± 0.077 0.669 ± 0.037		
DBN-Between DBN-Closeness DBN-Degree DBN-Eigen	$0.796 + 0.125$ 0.796 ± 0.075 0.801 ± 0.072 0.753 ± 0.086	$0.753 + 0.125$ 0.754 ± 0.075 0.756 ± 0.101 0.797 ± 0.086	$0.808 + 0.091$ 0.806 ± 0.074 0.825 ± 0.093 0.8 ± 0.111	$0.742 + 0.091$ 0.743 ± 0.074 0.724 ± 0.112 0.75 ± 0.111	$0.892 + 0.107$ 0.854 ± 0.086 0.891 ± 0.061 0.797 ± 0.067	0.657 ± 0.107 0.696 ± 0.086 0.704 ± 0.072 0.748 ± 0.073
GVAR-Between GVAR-Closeness GVAR-Degree GVAR-Eigen	0.736 ± 0.077 0.744 ± 0.093 0.791 ± 0.129	0.814 ± 0.077 0.806 ± 0.093 0.758 ± 0.129	0.816 ± 0.111 0.83 ± 0.114 0.746 ± 0.098	0.733 ± 0.111 0.72 ± 0.114 0.803 ± 0.098	0.741 ± 0.063 0.996 ± 0.01 0.816 ± 0.077	0.809 ± 0.063 0.554 ± 0.01 0.734 ± 0.077

Table 2: Average NDCG (Accuracy) for Identifying Sink & Source Nodes with Nonlinear Systems, 5 Rep.

— indicates retrieved NGC graph is degenerate, e.g., only edges are self-directed.

† indicates network measure cannot distinguish between nodes, e.g., all in/out degrees are equal.

Figure 1: Example Multi-species Lotka-Volterra Population with $d = 20$ and $T = 1000$. Warm colors refer to the 10 predator species and cool colors refer to the 10 prey species. Each predator hunts 2 prey and each prey is hunted by 2 predators.

Figure 2: Hourly PM 2.5 Concentration (FOD) by County During Kincade Fire - Oct. 22 to Nov. 7, 2019.

Kincade Wildfire - 12 Hour Spillover Effects

Figure 4: From left to right: Lower 95% HPDI Bound, Posterior Mean, and Upper 95% HPDI Bound. BSG for Kincade Fire, $h=12$ hours ahead. Note the strong variability in spillovers (edge weights) originating from Sonoma County and tighter intervals for indirect spillovers between San Francisco, Contra Costa, and Alameda counties.

References

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