# Supplementary Material

### Appendix A.1

We prove Theorem 2.1.

### Theorem 2.1:

Let  $\mathscr{F}$  be a family of functions from domain P to [0,1], for each  $f \in \mathscr{F}$  taken the probability uniform distribution  $\mu$  over the P, it holds:

$$var = \sup_{f \in \mathscr{F}} E_{\mu}[f^2].$$
<sup>(1)</sup>

Proof: Due to the normalization factor of  $\frac{1}{n(n-1)}$ , the value of BC is almost zero, especially on large-scale networks<sup>1</sup>. It is reasonable to be considered:

$$var = E_{\mu}[f^2] - \{E(f)\}^2 \le E_{\mu}[f^2].$$

# Appendix A.2

We are now ready to prove Theorem 2.2. Our most significant technical contributions are in this paper.

#### Theorem 2.2:

For  $k, m \ge 1$  and the function  $f \in \mathscr{F}$ , where  $\mathscr{F}$  be a family of functions from P to [0,1]. Let  $\lambda \in \{-1,+1\}^{k \times m}$  be an  $k \times m$ matrix of Rademacher random variables, so that  $\lambda \in \{-1,+1\}$  independently and with equal probability  $\frac{1}{2}$ . Let S be a sample size of *m* drawn i.i.d. from *P*, taken a distribution  $\mu$ . For each  $\delta \in (0,1)$ , define:

$$V(f) \doteq \alpha + \frac{\ln^3_{\overline{\delta}}}{m} + \sqrt{\left(\frac{\ln^3_{\overline{\delta}}}{m}\right)^2 + \frac{2\alpha \ln^3_{\overline{\delta}}}{m}}$$

$$\tilde{R}(\mathscr{F}, S) \doteq \tilde{R}_m^k(\mathscr{F}, S, \sigma) + \frac{2\ln^3_{\overline{\delta}}}{km} + \sqrt{\left(\frac{2\ln^3_{\overline{\delta}}}{km}\right)^2 + \frac{4(\tilde{R}_m^k(\mathscr{F}, S, \lambda) + \alpha)\ln^3_{\overline{\delta}}}{km}}$$

$$R(\mathscr{F}, m) \doteq \tilde{R}(\mathscr{F}, S) + \frac{\ln^3_{\overline{\delta}}}{m} + \sqrt{\left(\frac{\ln^3_{\overline{\delta}}}{m}\right)^2 + \frac{2\tilde{R}(\mathscr{F}, S)\ln^3_{\overline{\delta}}}{m}}$$

$$\varepsilon \doteq 2R(\mathscr{F}, m) + \frac{\ln^3_{\overline{\delta}}}{3m} + \sqrt{\left(\frac{\ln^3_{\overline{\delta}}}{3m}\right)^2 + \frac{2R(\mathscr{F}, m)\ln^3_{\overline{\delta}}}{m}},$$
(2)

With the probability at least  $1 - \delta$  over the choice of *S* and  $\lambda$ .

We first prove the origin of each formula in four steps.

### Step 1:

Theorem A.2.1 (Thm.7.5.8<sup>2</sup>) :

With probability  $\geq 1 - \eta$  over *S*, it holds:

$$\sup_{f\in\mathscr{F}}E(f^2)\leq \alpha+\frac{\ln\frac{1}{\eta}}{m}+\sqrt{(\frac{\ln\frac{1}{\delta}}{m})^2+\frac{2\alpha\ln\frac{1}{\eta}}{m}}.$$

*Proof: From Theorem 2.1,*  $V(f) = \sup_{f \in \mathscr{F}} E(f^2)$ , thus replacing  $\frac{1}{\lambda}$  with  $\frac{3}{\delta}$ , we can obtain the result.

# Step 2:

Before proving that  $\tilde{R}(\mathscr{F}, S) \doteq \tilde{R}_m^k(\mathscr{F}, S, \lambda) + \frac{2ln\frac{3}{\delta}}{nm} + \sqrt{\left(\frac{2ln\frac{3}{\delta}}{nm}\right)^2 + \frac{4(\tilde{R}_m^n(\mathscr{F}, S, \lambda) + \alpha)ln\frac{3}{\delta}}{nm}}$ , we need to define two functions that we need to use: the self-boundary function and  $g_{j,i}(\sigma)$ .

**Definition 1:**  $(\phi, \gamma)$ -self-bounding function

Let  $X = (X_1, ..., X_n)$  be a vector of random variables  $X_i$ , each taking values in a measurable set  $\chi$  and let g:  $\chi^n$  map R be a non-negative measurable function. The denote  $g_i$  a function from  $\chi^{n-1}$  map R.

A function *g* is a  $(\varphi, \gamma)$ -self-bounding function, for each  $X \in \chi^n$ :

$$0 \le g(X) - g_i(X^{(i)}) \le 1$$

$$\sum_{i=1}^n \{g(X) - g_i(X^i)\} \le \phi g(X) + \gamma,$$
(3)

where  $X^{(i)} = (X_1, ..., X_{i-1}, X_{i+1}, ..., X_N) \in \chi^{n-1}$ . **Theorem A.2.2:** 

Let  $\sigma \in \{-1,+1\}^{n \times m}$  be a  $n \times m$  matrix, define the function of  $g(\sigma)$ :

$$g(\boldsymbol{\sigma}) \doteq nm\tilde{R}^n_m(\mathscr{F}, m, \boldsymbol{\sigma}), \tag{4}$$

 $g(\sigma)$  is a self-bounding function of  $(1, 2nm\alpha)$ . Where  $\alpha = \sup_{f \in \mathscr{F}} \frac{1}{m} \sum_{i=1}^{m} (f(s_i))^2$ . **Definition 2:** the function of  $g_{j,i}(\sigma)$ , for  $j \in [1, n]$  and  $i \in [1, m]$  is defined as:

$$g_{j,i}(\boldsymbol{\sigma}) \doteq \inf_{\substack{\boldsymbol{\sigma}'_{j,i} \in \{-1,+1\}}} \left\{ \sum_{q=1,q\neq j}^{n} \left[ \sup_{f \in \mathscr{F}} \sum_{q=1}^{m} f(s_z) \boldsymbol{\sigma}_{q,z} \right] + \sup_{f \in \mathscr{F}} \left\{ \sum_{z=1, z\neq i} (\boldsymbol{\sigma}_{j,z} f(s_z)) + (\boldsymbol{\sigma}'_{j,i} f(s_i)) \right\} \right\},$$

we denote by  $g(\sigma)$  the function that replaces the element  $\sigma_{j,i}$  at the position (i, j) of  $\sigma$  with  $\sigma'$ , and we take the smallest value over  $\sigma'$ .

**Proof of Theorem A.2.2:** 

Now proceed to prove Theorem A.2.2, i.e., prove that:

$$0 \le g(\sigma) - g_{j,i}(\sigma) \le 1$$
$$\sum_{j=1}^{n} \sum_{i=1}^{m} (g(\sigma) - g_{j,i}(\sigma))^2 \le \phi g(\sigma) + \gamma, (\phi, \gamma) \ge 0.$$

First, we can rewrite  $g(\sigma)$  as the following:

$$g_{j,i}(\boldsymbol{\sigma}) = \min\left[\sum_{q=1,q\neq j}^{n} \left[\sup_{f\in\mathscr{F}}\sum_{z=1}^{m} \boldsymbol{\sigma}_{q,z}f(s_{z})\right] + \sup_{f\in\mathscr{F}}\left\{\sum_{z=1,z\neq i}^{m} (\boldsymbol{\sigma}_{j,z}f(s_{z}) - f(s_{i}))\right\},\\ \sum_{q=1,q\neq j}^{n} \left[\sup_{f\in\mathscr{F}}\sum_{z=1}^{m} \boldsymbol{\sigma}_{q,z}f(s_{z})\right] + \sup_{f\in\mathscr{F}}\left\{\sum_{z=1,z\neq i}^{m} (\boldsymbol{\sigma}_{j,z}f(s_{z}) + f(s_{i}))\right\}\right].$$

It is easy to see that at least one element of this min equation is equal to  $g(\sigma)$ , so the minimum is either  $g(\sigma)$  or smaller than  $g(\sigma)$ . Next, we start the proof of Theorem A.2.2 that:

$$g_{j,i}(\sigma) \doteq \inf_{\sigma'_{j,i}} \sum_{q=1,q\neq j}^{n} \left[ \sup_{f \in \mathscr{F}} \sum_{z=1}^{m} \sigma_{q,z} f(s_z) \right] + \sup_{f \in \mathscr{F}} \left\{ \sum_{z=1,z\neq i}^{m} (\sigma_{j,z} f(s_z) + \sigma'_{j,i} f(s_i))) \right\}$$
$$= \sum_{q=1,q\neq j}^{n} \left[ \sup_{f \in \mathscr{F}} \sum_{z=1}^{m} \sigma_{q,z} f(s_z) \right] + \inf_{\sigma'_{j,i}} \left\{ \sup_{f \in \mathscr{F}} \left\{ \sum_{z=1,z\neq i}^{m} (\sigma_{j,i} f(s_z)) + \sigma'_{j,i} f(s_i) \right\} \right\}$$
$$\geq \sum_{q=1,q\neq j}^{n} \left[ \sup_{f \in \mathscr{F}} \sum_{z=1}^{m} \sigma_{q,z} f(s_z) \right] + \sup_{f \in \mathscr{F}} \left\{ \inf_{\sigma'_{j,i} \in \{-1,+1\}} \left\{ \sum_{z=1,z\neq i}^{m} (\sigma_{j,i} f(s_z)) + \sigma'_{j,i} f(s_i) \right\} \right\}$$
$$= \sum_{q=1,q\neq j}^{n} \left[ \sup_{f \in \mathscr{F}} \sum_{z=1}^{m} \sigma_{q,z} f(s_z) \right] + \sup_{f \in \mathscr{F}} \left\{ \sum_{z=1,z\neq i}^{m} (\sigma_{j,i} f(s_z)) + \inf_{\sigma'_{j,i} \in \{-1,+1\}} \left\{ \sigma'_{j,i} f(s_i) \right\} \right\}.$$

2/<mark>6</mark>

For a given  $\sigma$ , let  $f_j^*$  be one of the functions of  $\mathscr{F}$  attaining the supremum of  $\sup_{f \in \mathscr{F}_{z=1}} \sum_{z=1}^m \sigma_{j,i} f(s_z)$ . Thus, We can keep writing this up here:

$$\begin{split} g_{j,i}(\sigma) &\geq \sum_{q=1,q\neq j}^{n} \left[ \sup_{f\in\mathscr{F}} \sum_{z=1}^{m} \sigma_{q,z} f(s_{z}) \right] + \sup_{f\in\mathscr{F}} \left\{ \sum_{z=1,z\neq i}^{m} (\sigma_{j,i}f(s_{z})) + \inf_{\sigma'_{j,i}\in\{-1,+1\}} \{\sigma'_{j,i}f(s_{i})\} \right\} \\ &\geq \sum_{q=1,q\neq j}^{n} \left[ \sup_{f\in\mathscr{F}} \sum_{z=1}^{m} \sigma_{q,z}f(s_{z}) \right] + \sum_{z=1,z\neq i}^{m} (\sigma_{j,i}f_{j}^{\star}(s_{z})) + \inf_{\sigma'_{j,i}\in\{-1,+1\}} \{\sigma'_{j,i}f_{j}^{\star}(s_{i})\} \\ &= \sum_{q=1,q\neq j}^{n} \left[ \sup_{f\in\mathscr{F}} \sum_{z=1}^{m} \sigma_{q,z}f(s_{z}) \right] + \sum_{z=1,z\neq i}^{m} (\sigma_{j,i}f_{j}^{\star}(s_{z})) + \sigma_{j,i}f_{j}^{\star}(s_{i}) - \sigma_{j,i}f_{j}^{\star}(s_{i}) \\ &+ \inf_{\sigma'_{j,i}\in\{-1,+1\}} \{\sigma'_{j,i}f_{j}^{\star}(s_{i})\} \\ &= g(\sigma) - \sigma_{j,i}f_{j}^{\star}(s_{i}) + \inf_{\sigma'_{j,i}\in\{-1,+1\}} \{\sigma'_{j,i}f_{j}^{\star}(s_{i})\}, \end{split}$$

where  $f_j^{\star}(s_i) \in [0, 1]$ , we can obtain that:  $g_{j,i}(\sigma) \ge g(\sigma) - \sigma_{j,i} f_j^{\star}(s_i) - |f_j^{\star}(s_i)| \ge g(\sigma) - 1$ . Now, the proof of  $\phi = 1, \gamma = 2nm$  as follows:

$$\begin{split} \sum_{j=1}^{n} \sum_{i=1}^{m} \left( g(\sigma) - g_{j,i}(\sigma) \right)^2 &\leq \sum_{j=1}^{n} \sum_{i=1}^{m} \left( \sigma_{j,i} f_j^{\star}(s_i) + |f_j^{\star}(s_i)| \right)^2 \\ &= \sum_{j=1}^{n} \sum_{i=1}^{m} \left( [\sigma_{j,i} f_j^{\star}(s_i)]^2 + |f_j^{\star}(s_i)|^2 + 2\sigma_{j,i} f_j^{\star}(s_i)|f_j^{\star}(s_i)| \right) \\ &= \sum_{j=1}^{n} \sum_{i=1}^{m} \left( 2f_j^{\star}(s_i)^2 + 2\sigma_{j,i} f_j^{\star}(s_i)|f_j^{\star}(s_i)| \right) \\ &\leq \sum_{j=1}^{n} \sum_{i=1}^{m} \sigma_{j,i} f_j^{\star}(s_i) + 2\sum_{j=1}^{n} \sum_{i=1}^{m} f_j^{\star}(s_i)^2 \\ &= g(\sigma) + 2\sum_{j=1}^{n} \sum_{i=1}^{m} f_j^{\star}(s_i)^2 \\ &\leq g(\sigma) + 2n \sup_{f \in \mathscr{F}} \sum_{i=1}^{m} (f(s_i))^2 \\ &= g(\sigma) + 2nm\alpha \,, \end{split}$$

obtaining the statement.

#### Theorem A.2.3<sup>3</sup>

*let*  $\lambda \in \{-1,+1\}^{n \times m}$  *be an*  $n \times m$  *matrix of matrix of Rademcher random variables,*  $\lambda_{j,i} \in \{-1,+1\}$  *with probability*  $\frac{1}{2}$  *taken each one and independent. Then, for all*  $0 \le \tau \le \tilde{R}(\mathscr{F},S)$ :

$$Pr(\tilde{R}(\mathscr{F},S) \ge \tilde{R}_{m}^{n}(\mathscr{F},S,\lambda) + \tau) \le exp(-\frac{nm\tau^{2}}{4(\check{R}(\mathscr{F},S) + \alpha)}).$$
(5)

Theorem A.2.3 plays a role in solving the second formula, so we prove it as follows: The function of f, defines:

$$\hat{\mathscr{F}} \doteq \hat{f}(x) \doteq \frac{f(x)}{2} : f \in \mathscr{F}, \forall x \in \chi.$$

With theorem A.2.2, we know that  $g(\sigma) = nm\tilde{R}_m^n(\hat{F}, S, \lambda)$  is a 1,2 $nm\sigma$  self-bounding function. It is easily to get the statement that  $E_{\lambda}[nm\tilde{R}_m^n(\hat{\mathcal{F}}, S, \lambda)] = nm\tilde{R}(\hat{\mathcal{F}}, S)$ . We apply (Theorem.7<sup>3</sup>), obtaining:

$$Pr(nm\tilde{R}(\hat{\mathscr{F}},S) \ge nm\tilde{R}_{m}^{n}(\hat{\mathscr{F}},S,\lambda) + t) \le exp(-\frac{t^{2}}{2(nm\tilde{R}(\hat{\mathscr{F}},S) + 2nm\alpha)}),$$
(6)

we use the fact that  $\tilde{R}(\hat{\mathscr{F}}, S) = \frac{\tilde{R}(\hat{\mathscr{F}}, S)}{2}$ ,  $\tilde{R}_m^n(\hat{\mathscr{F}}, S, \lambda) = \frac{\tilde{R}_m^n(\hat{\mathscr{F}}, S, \lambda)}{2}$  and  $\alpha_{\hat{\mathscr{F}}} = \frac{\alpha}{4}$ . It follows that:

$$Pr(\frac{nm}{2}\tilde{R}(\mathscr{F},S) \ge \frac{nm}{2}\tilde{R}_{m}^{n}(\mathscr{F},S,\lambda) + t \le exp(-\frac{t^{2}}{nm(\tilde{R}(\mathscr{F},S)) + \alpha})),$$
(7)

replacing t by  $\frac{\tau_{nm}}{2}$  achieve the Theorem A.2.3.

The proof of the key equation begins below:

$$\tilde{R}(\mathscr{F},S) = \tilde{R}_m^n(\mathscr{F},S,\lambda) + \frac{2ln\frac{1}{\delta}}{nm} + \sqrt{(\frac{2ln\frac{1}{\delta}}{nm})^2 + \frac{4(\tilde{R}_m^n(\mathscr{F},S,\lambda) + \alpha)ln\frac{1}{\delta}}{nm}}$$

**Proof: From Theorem A.2.3**, we have the fact that, with probability  $\geq 1 - \delta$ ,

$$\tilde{R}(\mathscr{F},S) \leq \tilde{R}_m^n(\mathscr{F},S,\lambda) + \sqrt{\frac{4(\tilde{R}(\mathscr{F},S)+\alpha)\ln\frac{1}{\delta}}{nm}}$$

The bound of  $\tilde{R}(\mathcal{F}, S)$  can be obtained by the function of b(x), which can find the fixed point.

$$b(x) \doteq \tilde{R}_m^n(\mathscr{F}, S, \lambda) + \sqrt{\frac{4(x+\alpha)ln\frac{1}{\delta}}{nm}}$$

**Lemma 1.** *let*  $d, j, v \ge 0$ *. The fixed point of* 

$$b(x) \doteq d + \sqrt{j + vx},$$

is at

$$x \doteq d + \frac{v}{2} + \sqrt{\frac{v^2}{4} + j + vj}$$

Therefore, we apply lemma 1 to obtain the statement.

#### Step 3:

We have rigorously proved the two formulas before, and then our main technical proof Theorem 2.2 is based on the concentration inequality for Rademachaer Averages, for the supremum deviation  $S(\mathscr{F}, S) = \sup_{\substack{f \in \mathscr{F} \\ f \in \mathscr{F}}} |\rho_s(f) - \rho_\mu(f)|$ . To facilitate

the discussion, we write  $S(\mathscr{F}, S)$  as  $Z^4$ .

We now start to describe. Then, define the *Rademacher complexity*  $(RC)R(\mathscr{F},m)$  of a set of functions  $\mathscr{F}$  as the expection of the ERA over *S*.  $R(\mathscr{F},m) \doteq E_S[\tilde{R}(\mathscr{F},S)]$ . The following central results correlate  $\tilde{R}(\mathscr{F},m)$  with the expected supremum deviation.

Lemma 2. Symmetrization lemma<sup>5</sup>

$$E_{\mathcal{S}}[Z] \le 2R(\mathscr{F},m). \tag{8}$$

The following shows the deviation of the variance-dependent constraint above its expected value. **Theorem A.2.4**<sup>6</sup> Let  $Z = \sup_{f \in \mathscr{F}} |\rho_s(f) - \rho_\mu(f)|$ . Then, with probability at least  $1 - \lambda$  over S, it holds

$$Z \leq E[Z] + \frac{ln\frac{1}{\lambda}}{3m} + \sqrt{\left(\frac{ln\frac{1}{\lambda}}{3m}\right)^2 + \frac{2(E[Z]+1)ln\frac{1}{\lambda}}{m}}.$$

We apply lemma 2. to Theorem A.2.3 can obtain

$$\varepsilon \doteq 2R(\mathscr{F},m) + \frac{ln\frac{1}{\lambda}}{3m} + \sqrt{(\frac{ln\frac{1}{\lambda}}{3m})^2 + \frac{2R(\mathscr{F},m)ln\frac{1}{\lambda}}{m}},$$

the next result bounds  $R(\mathscr{F}, m)$  above its estimated  $\tilde{R}(\mathscr{F}, S)$ .

Step 4:

Theorem A.2.5.<sup>7</sup>

With probability  $\geq 1 - \lambda$  over *S*, it holds

$$R(\mathscr{F},m) \leq \tilde{R}(\mathscr{F},S) + \frac{ln\frac{1}{\lambda}}{m} + \sqrt{(\frac{ln\frac{1}{\lambda}}{m})^2 + \frac{2\tilde{R}(\mathscr{F},S)ln\frac{1}{\lambda}}{m}}$$

#### **Conclusion:**

Now we prove Theorem 2.2 in its entirety, the most important part in our paper. Proof. In the 4 formulas for our most vital results, replace  $\frac{3}{\delta}$  with  $\frac{1}{\lambda}$  to obtain Theorem 2.2.

### Appendix A.3

We now prove Theorem 3.1, which provides probabilistic quality assurance for the CBCA algorithm.

### Theorem 3.1:

With probability at least  $1 - \delta$  for the CBCA algorithm, the output  $(\tilde{B}, \varepsilon)$ , such that  $|b(v) - \tilde{b}(v) \le \varepsilon|$ . Before proving, we need to understand the following facts will improve the efficiency of proof. Facts:

(1). At the end of each iteration, the wimpy variance V(f), for any  $f \in \mathscr{F}$ , can be computed by replacing  $\delta$  with probability  $\frac{\delta}{2^{i+1}}$  by the Eq.(2) in Theorem 2.2.

(2). At the end of each iteration, Monte Carlo empirical Rademacher, empirical Rademacher values, and Rademacher values, for any  $fin\mathscr{F}$ , can be computed by replacing  $\delta$  with probability  $\frac{\delta}{2^{i+1}}$  by the Eq.(2) in Theorem 2.2.

(3). At the end of each iteration,  $\varepsilon$ , for any  $f \in \mathscr{F}$ , can be computed by replacing  $\delta$  with probability  $\frac{\delta}{2^{i+1}}$  by the Eq.(2) in Theorem 2.2.

(4). After sampling m samples, the maximum error  $S(\mathscr{F}, S)$  is at most  $\varepsilon$  (i.e.,  $m_i \ge m_1$ ), and the probability is  $1 - \frac{\delta}{2}$ . (this can refer to the Matteo<sup>8</sup>)

(5). For each  $i \ge 1, S_i = \{m_1, ..., m_{s_i}\}$  is the set of  $S_i$  independent uniform samples from P.

The probability in Theorem 3.1 is taken on all realizations of sequence  $S_i$ , that is, on all realizations of sequence  $m_j$ . Proof: Let events *E*1 and *E*2 be defined as:

 $E1 = \{ \exists i \ge 1 \, s.t. \, |b(v) - \tilde{b}(v)| > \varepsilon \}$  $E2 = \{ \exists j, i > 1 \, s.t. \, |b(v) - \tilde{b}(v)| > \varepsilon \}.$ 

It can be known that the output of the algorithm satisfies the  $\varepsilon$  – *approximation* condition when both events *E*1, *E*2 are wrong. Thus, we only need to prove  $Pr(E1 \cup E2) \leq \delta$ .

For statement the sake of fact (4), we can obtain that  $Pr(E1) \le \frac{\delta}{2}$ . In consideration of fact (1)(2)(3)(5), we can be calculated:

$$\begin{aligned} \Pr(E2) &= \Pr(\exists j, i \quad s.t. | b(v) - \tilde{b}_{s_{i_j}} | > \varepsilon) \leq \sum_i \Pr(|b(v) - \tilde{b}_{s_{i_j}} | > \varepsilon) \\ &\leq \sum_i \frac{\delta}{2^{i+1}} \leq \frac{\delta}{2}. \end{aligned}$$

We can obtain that

 $Pr(E1 \cup E2) \leq Pr(E1) + Pr(E2) \leq \delta.$ 

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